

Reproducible Statistical Inference

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+ Faculty of Computing & Data Sciences
Boston University*

Based on work with Jeff Miller and Jiawei Li

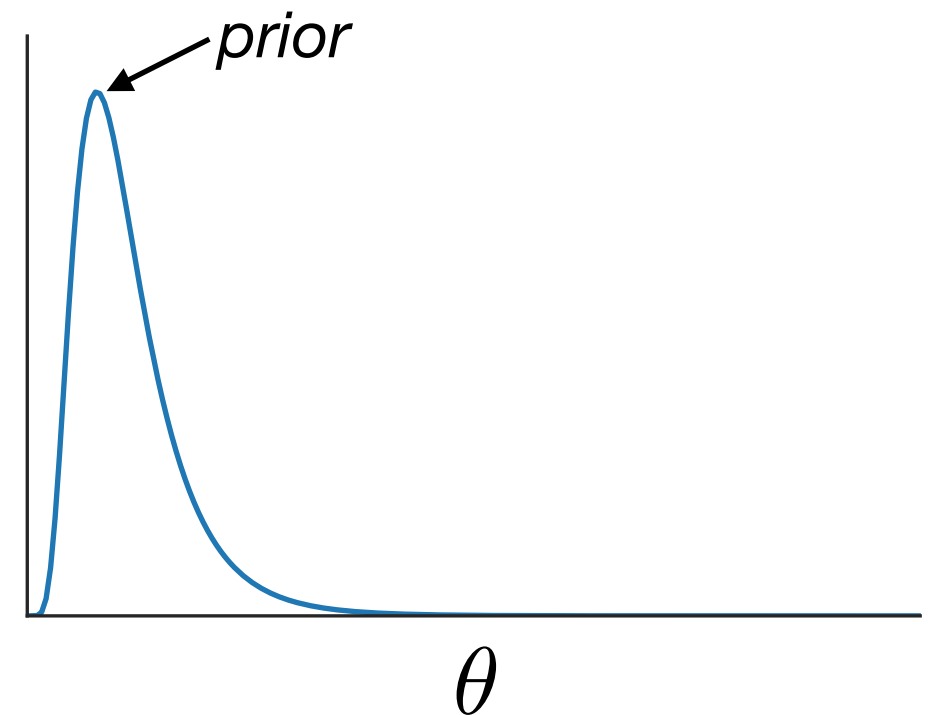
Bayesian inference

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- **Goal:** learn about unobserved phenomenon (parameter) of interest θ [e.g., “skill” of a baseball player]

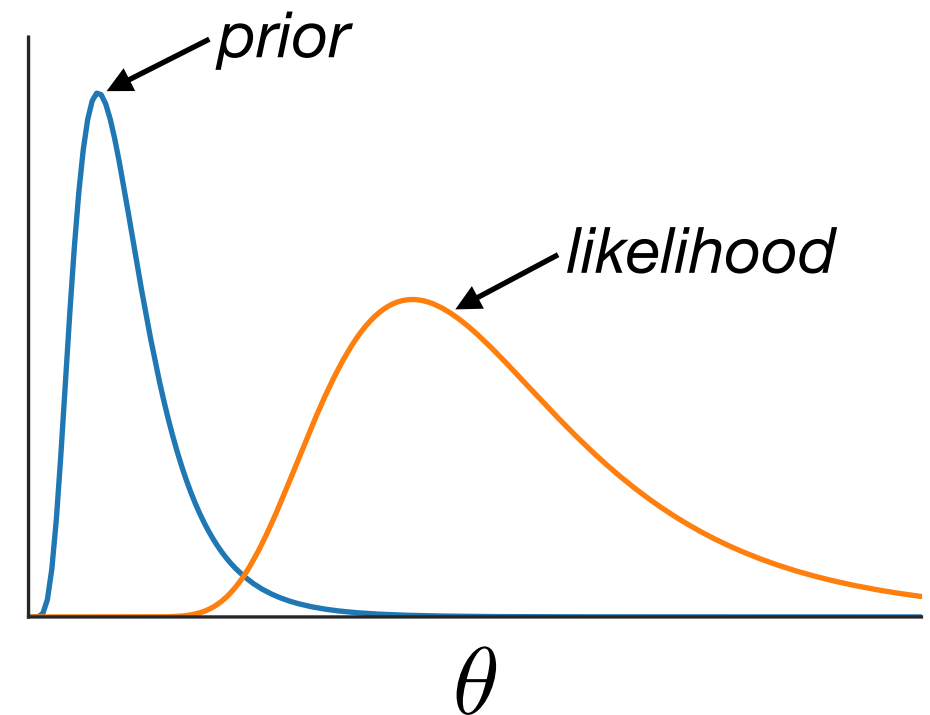
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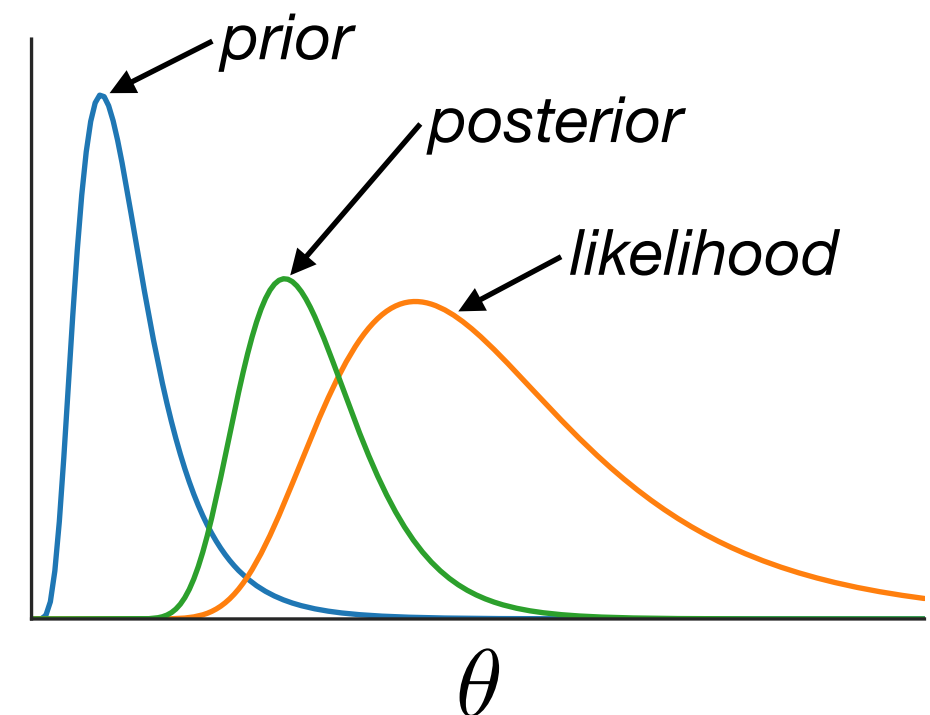
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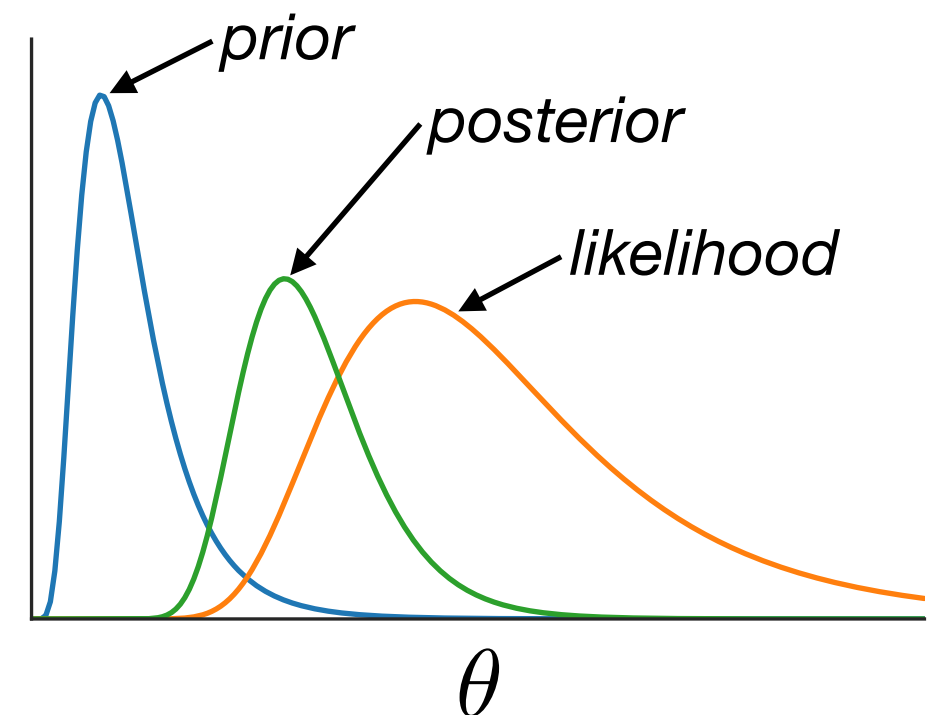
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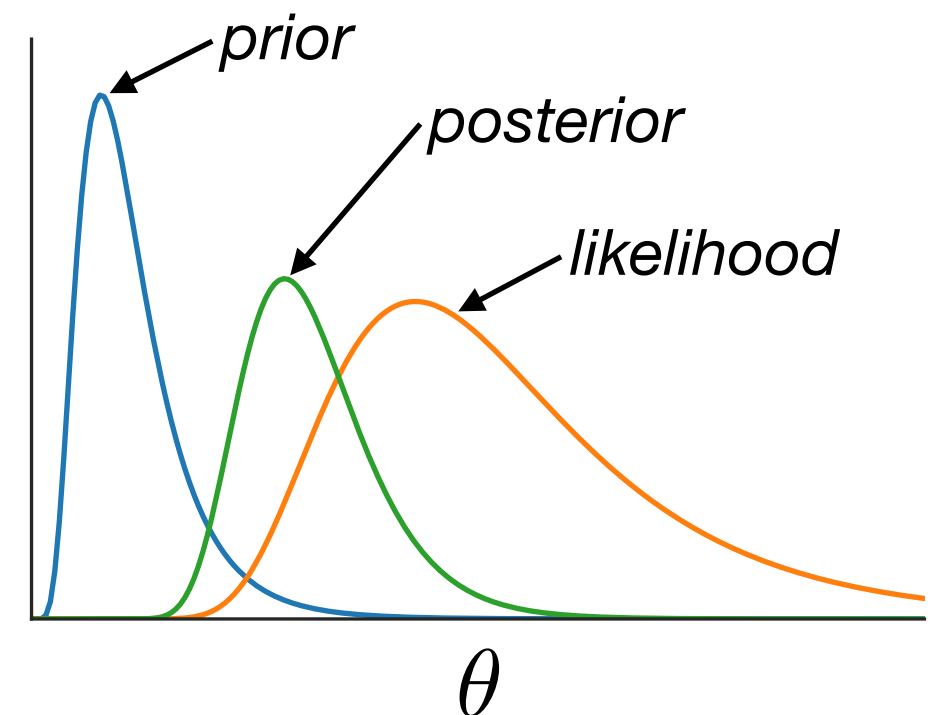


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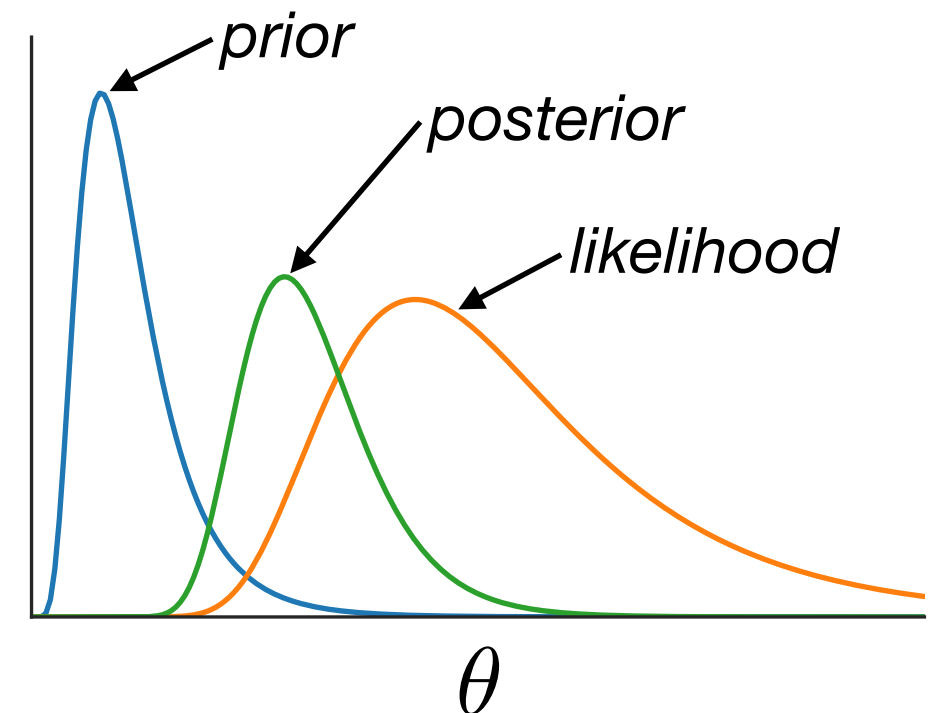


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- **Assumption:** measurement model correct: *observed* x has distribution $p(x | \theta_{\text{true}})$



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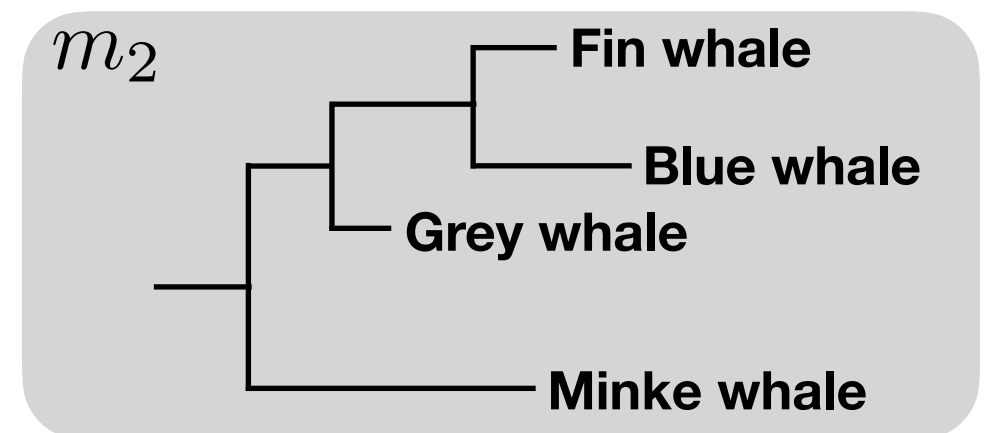
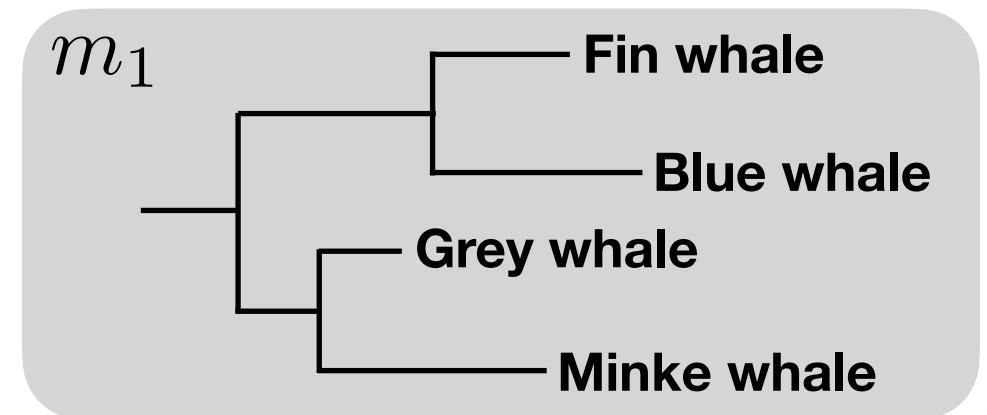
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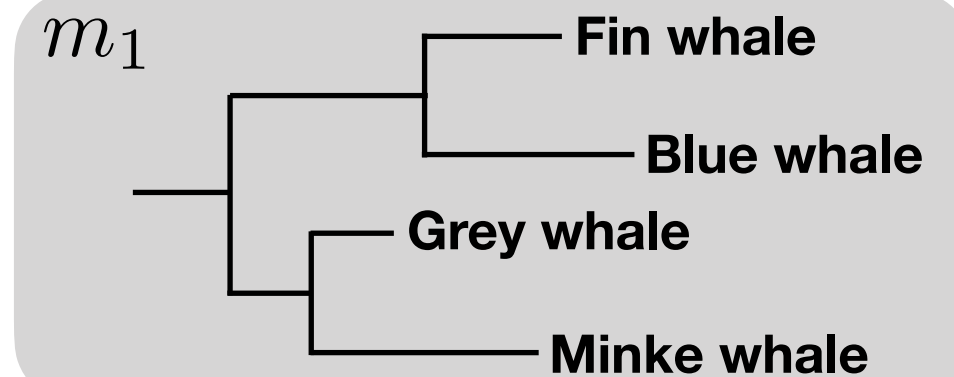
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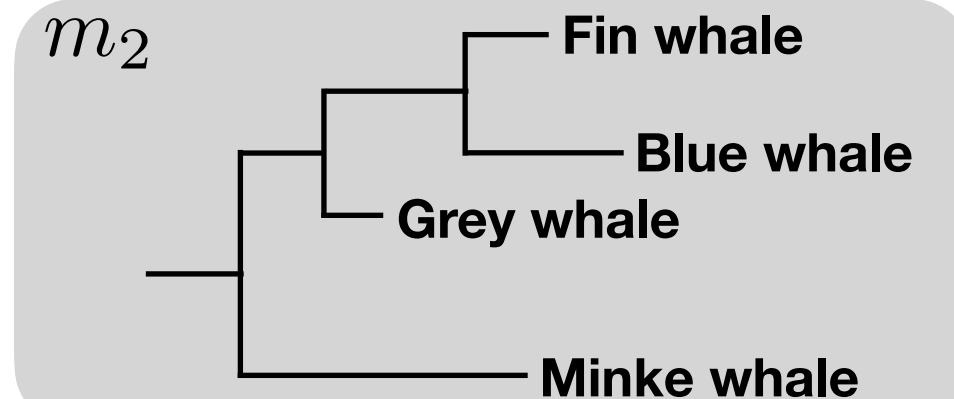
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$$\pi(m_i | x) = \frac{p(x | m_i)\pi_0(m_i)}{\sum_j p(x | m_j)\pi_0(m_j)}$$

$$\pi(m_1 | x) = 0.8$$



$$\pi(m_2 | x) = 0.1$$



⋮

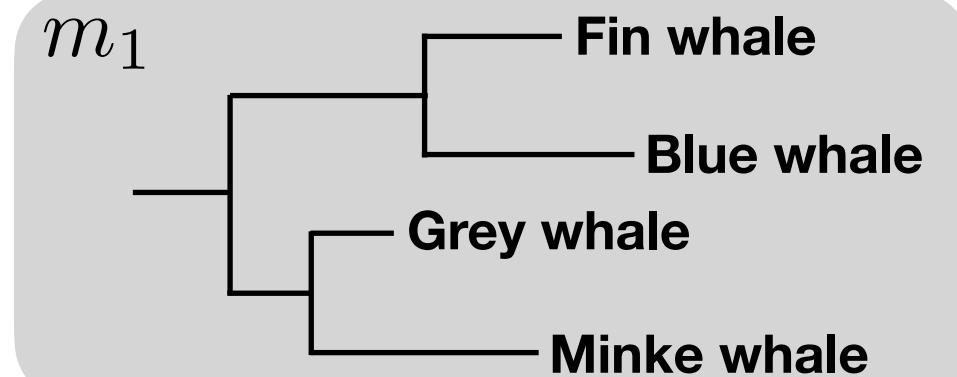
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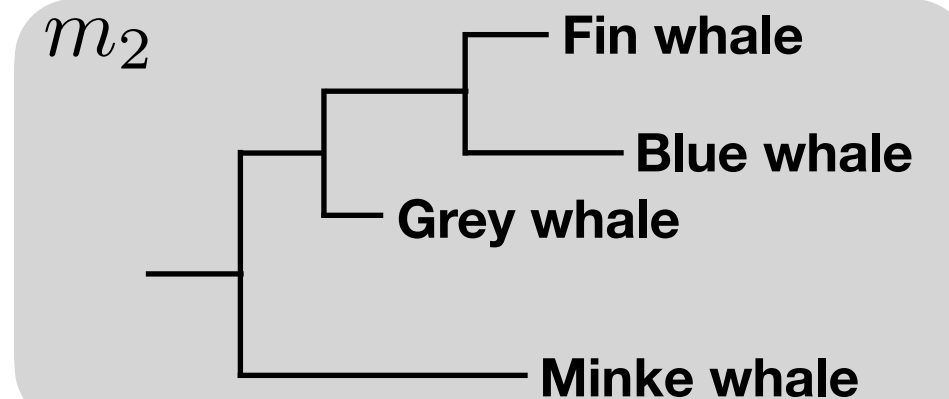
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- **Problem:** (Bayesian) model selection doesn't always work as we might hope...

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- **Problem:** infer phylogeny of 13 whale species from mitochondrial coding DNA

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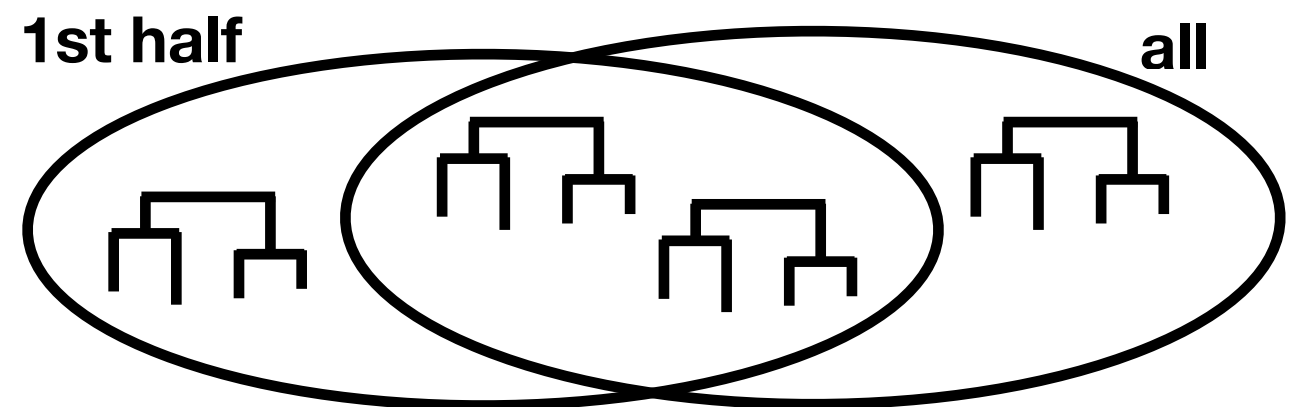
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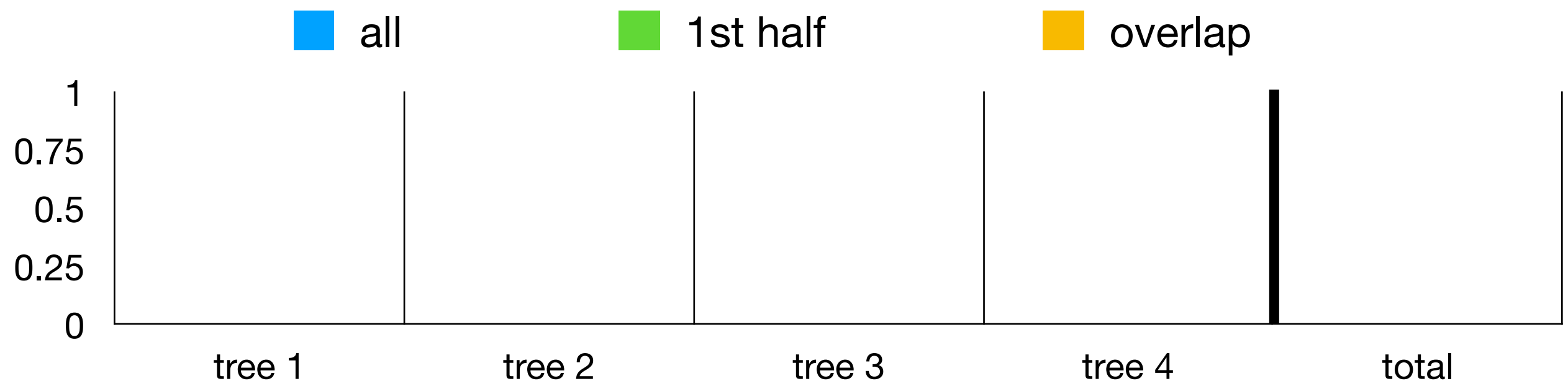
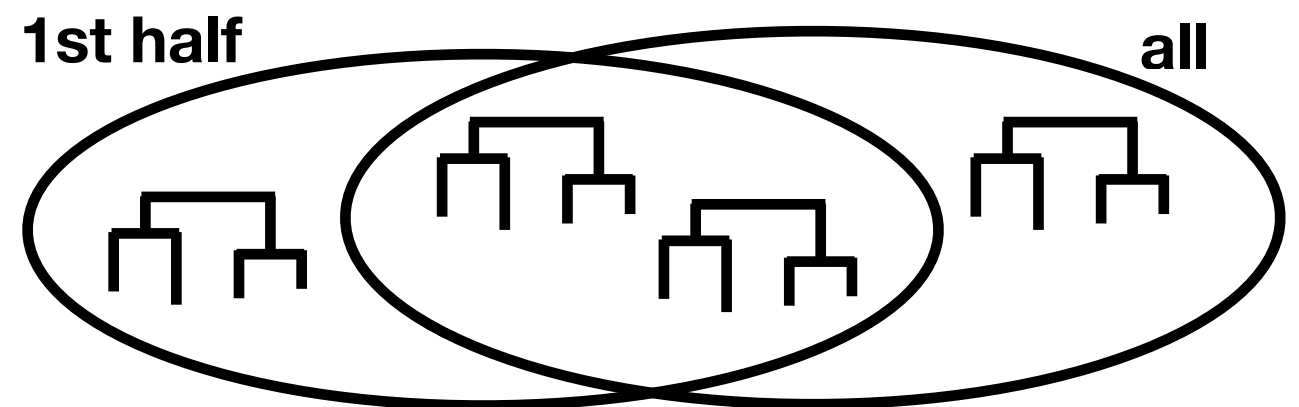
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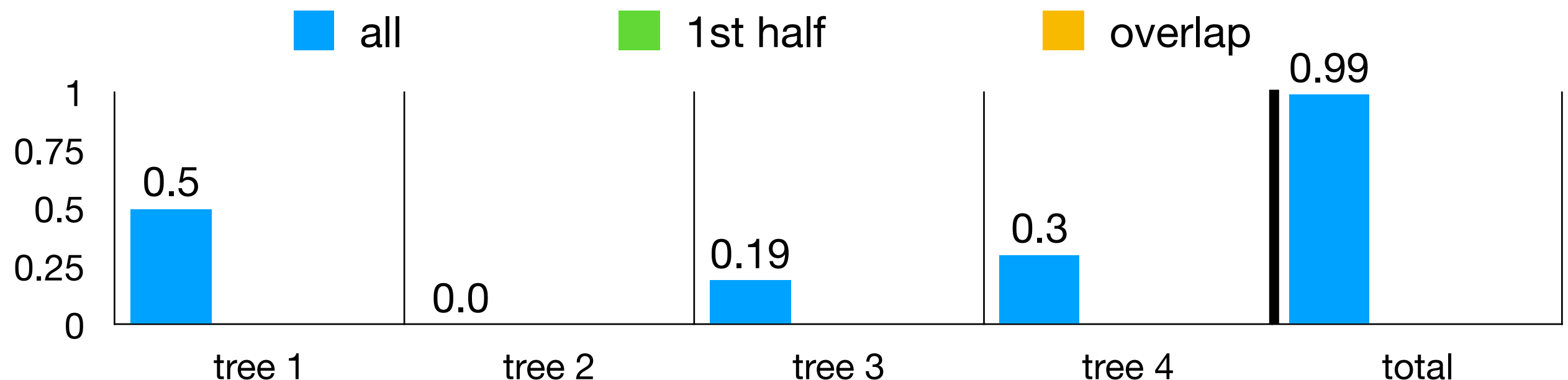
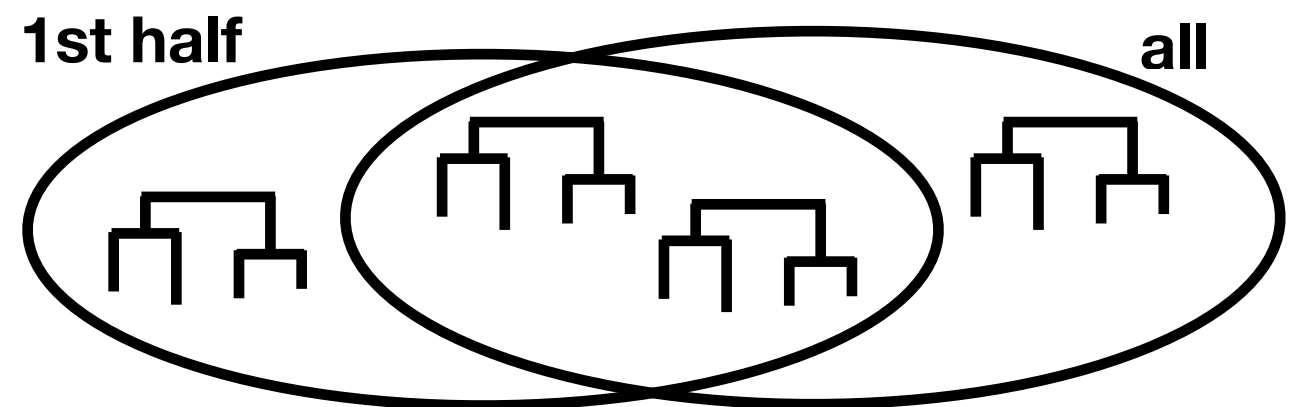
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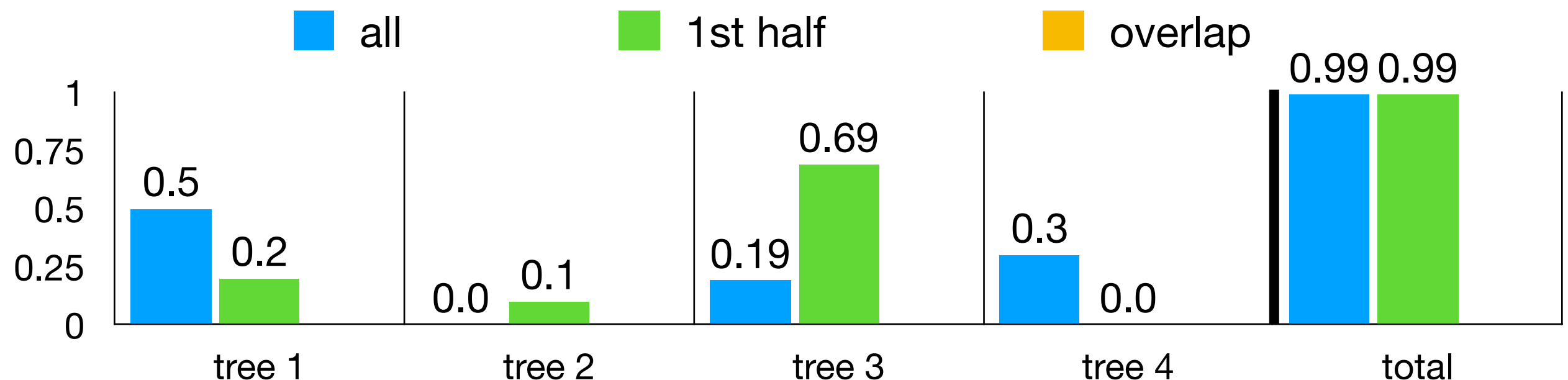
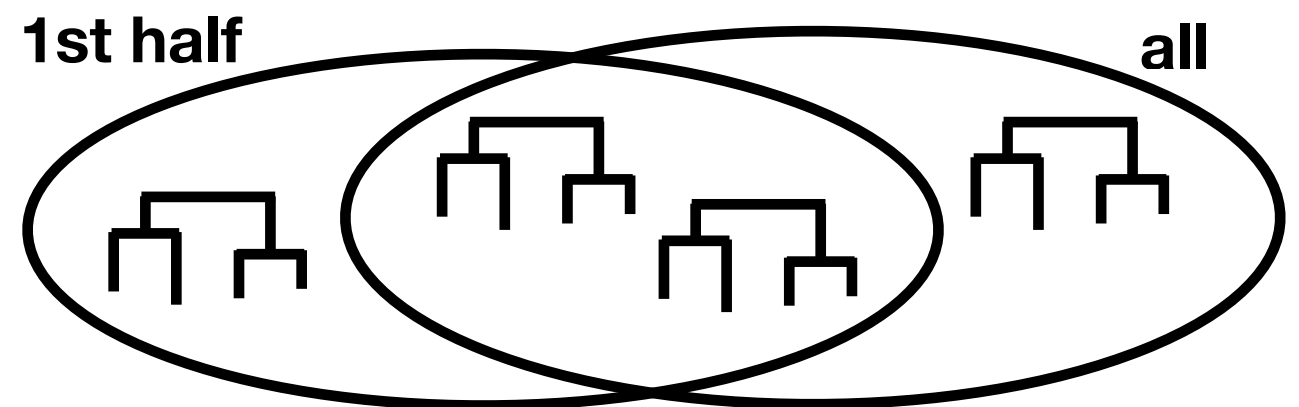
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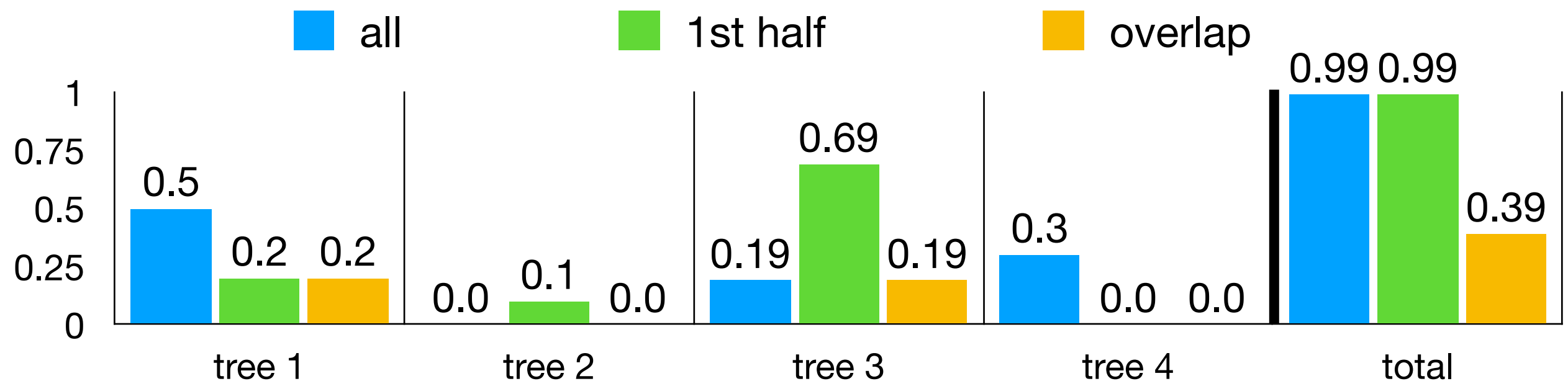
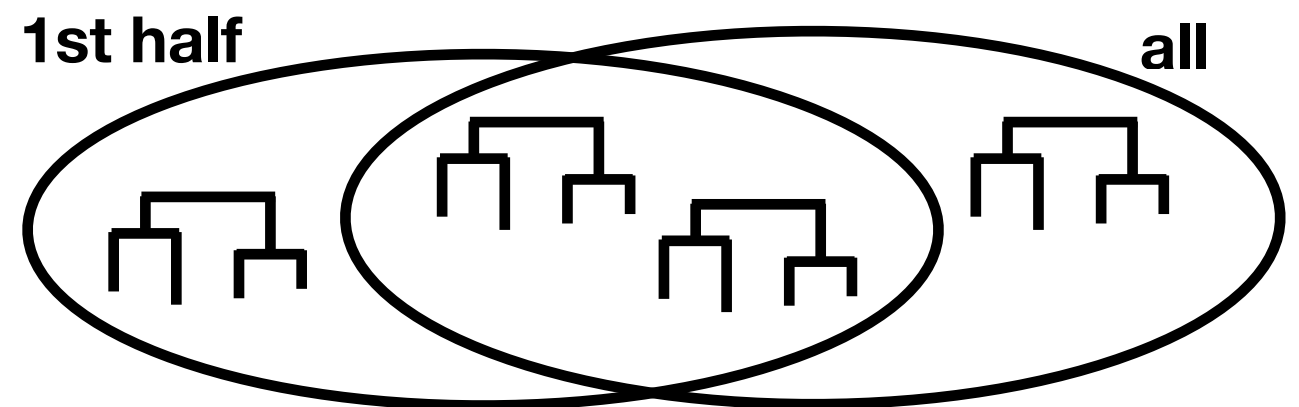
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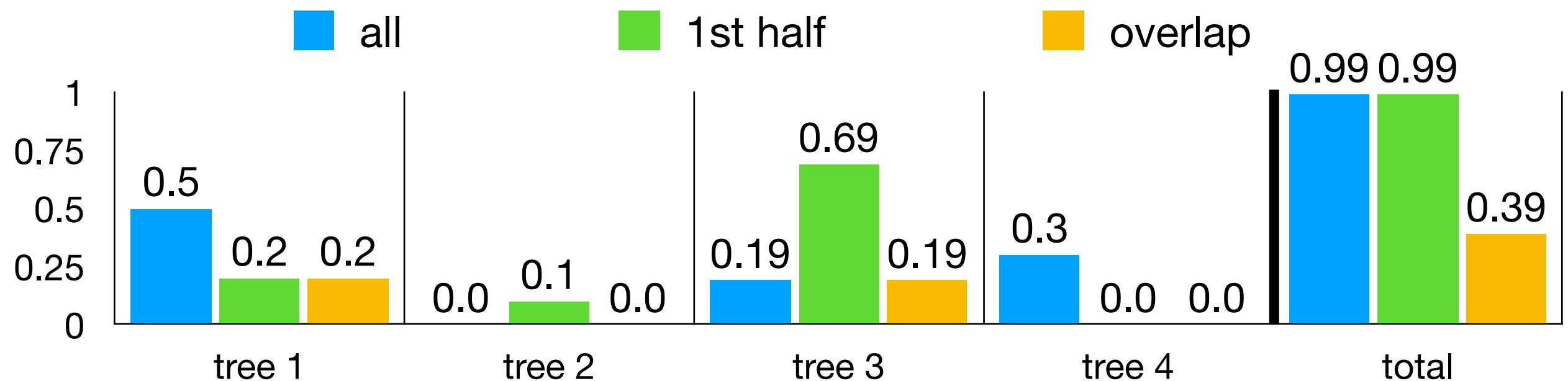
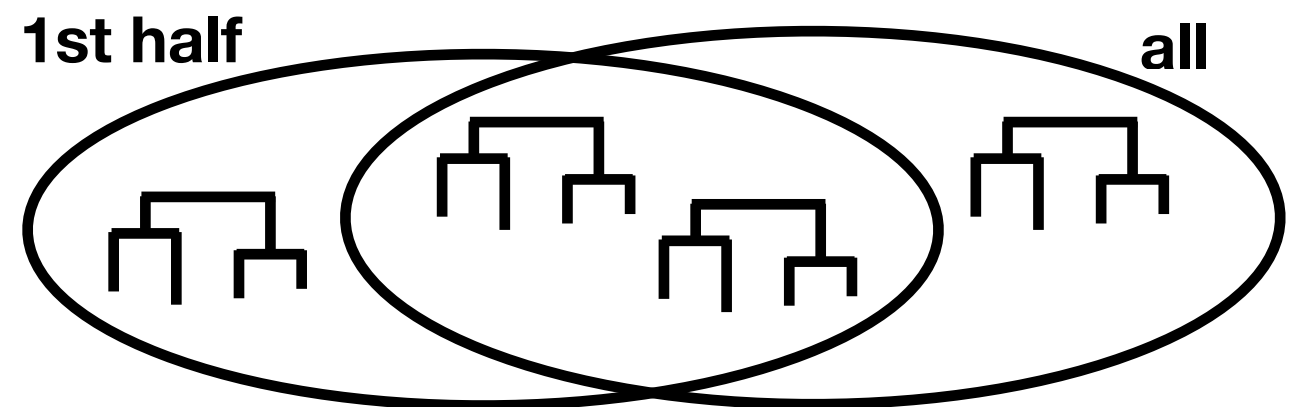
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- 0% overlap = contradiction

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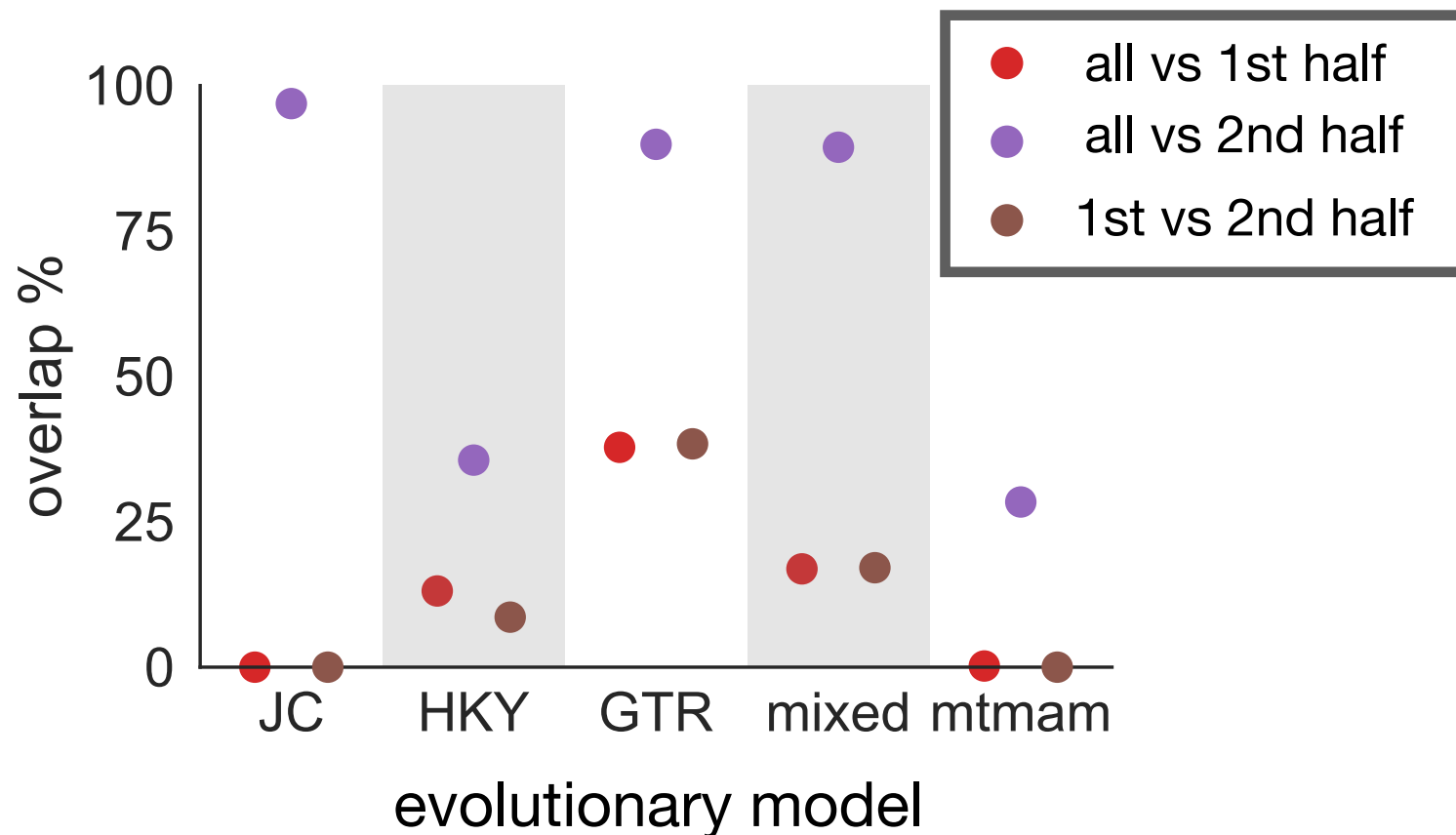
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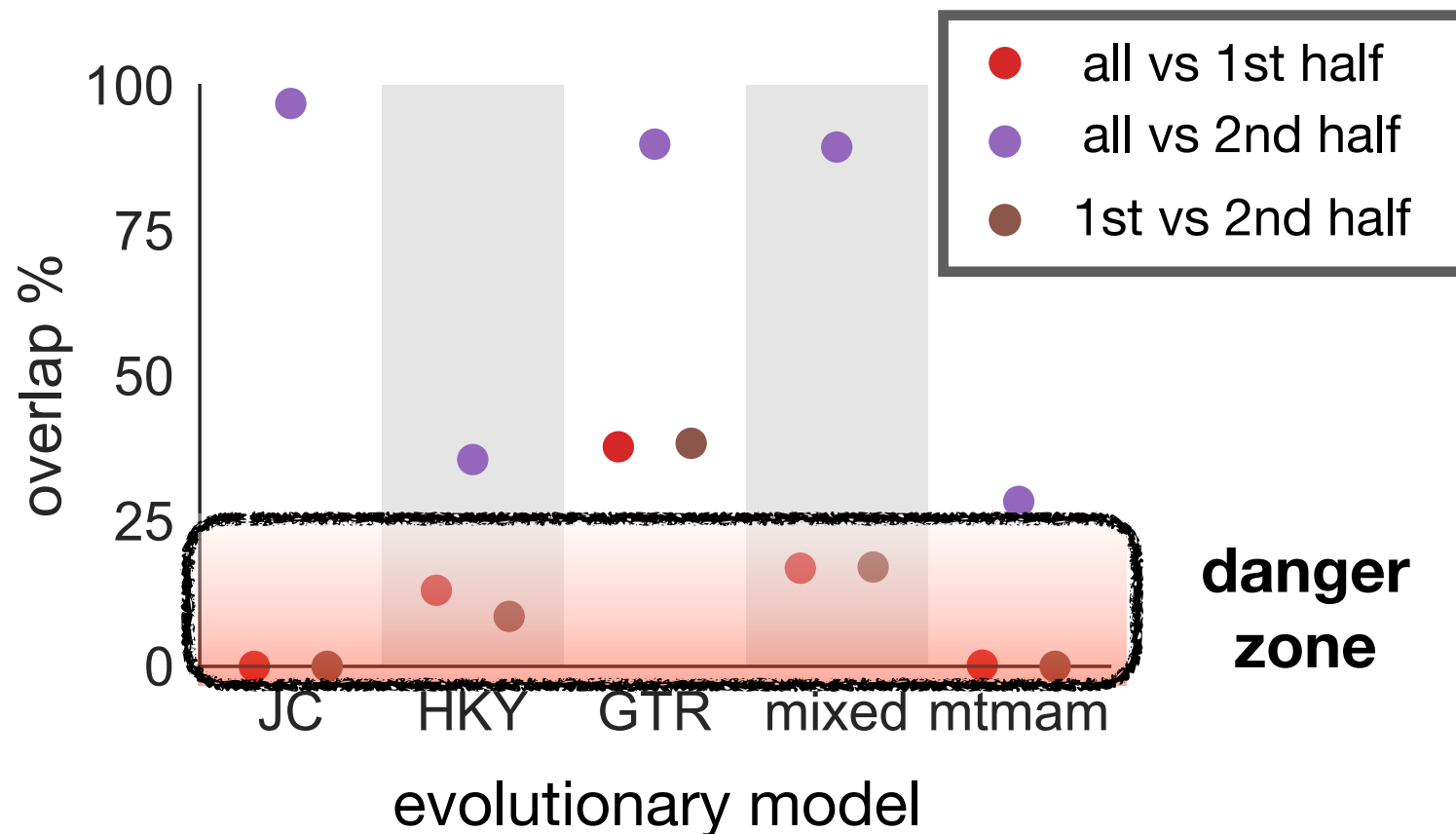


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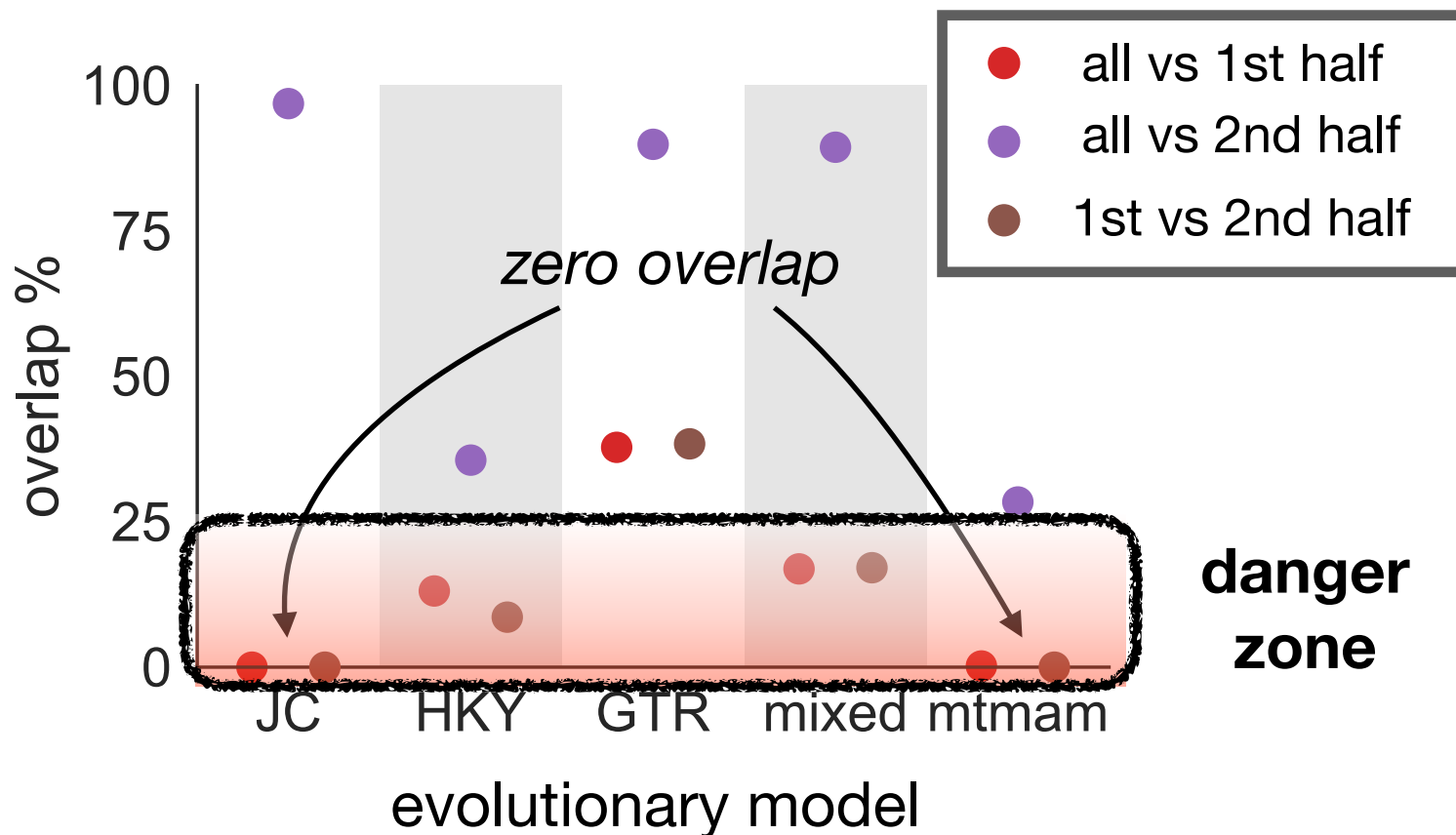


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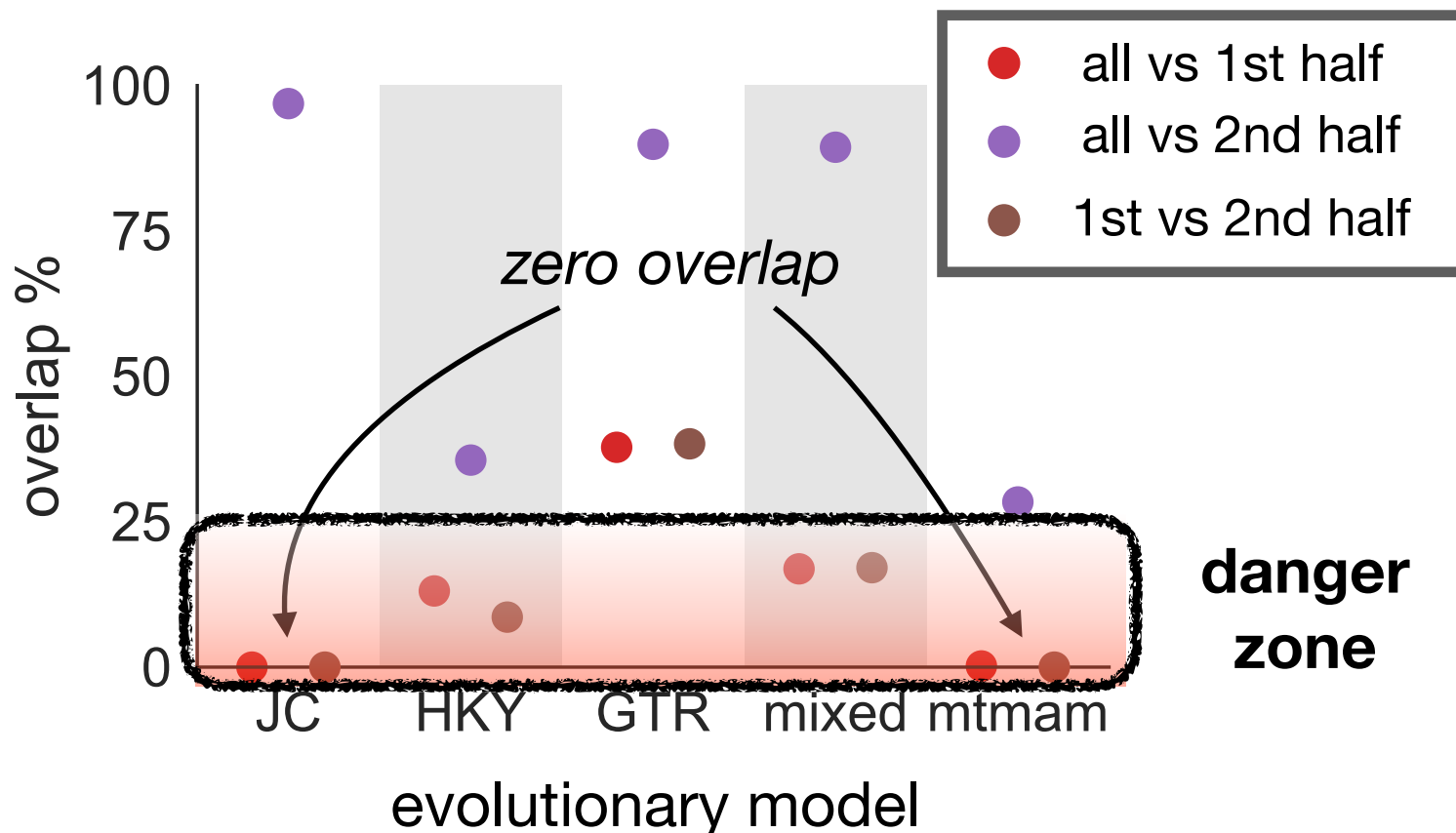


Bayesian phylogenetic inference may not be reproducible

Problem: infer phylogeny of 13 whale species

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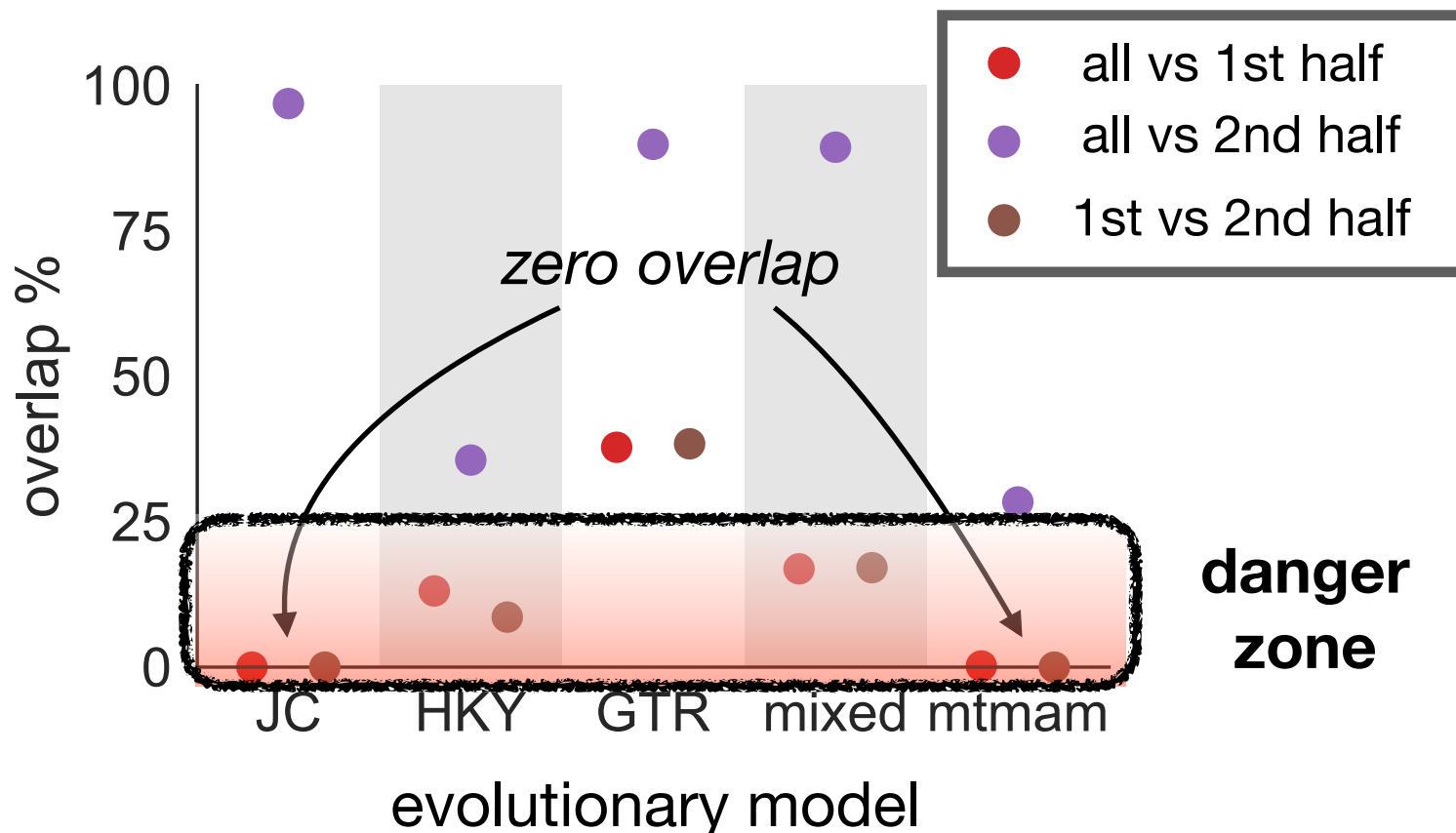


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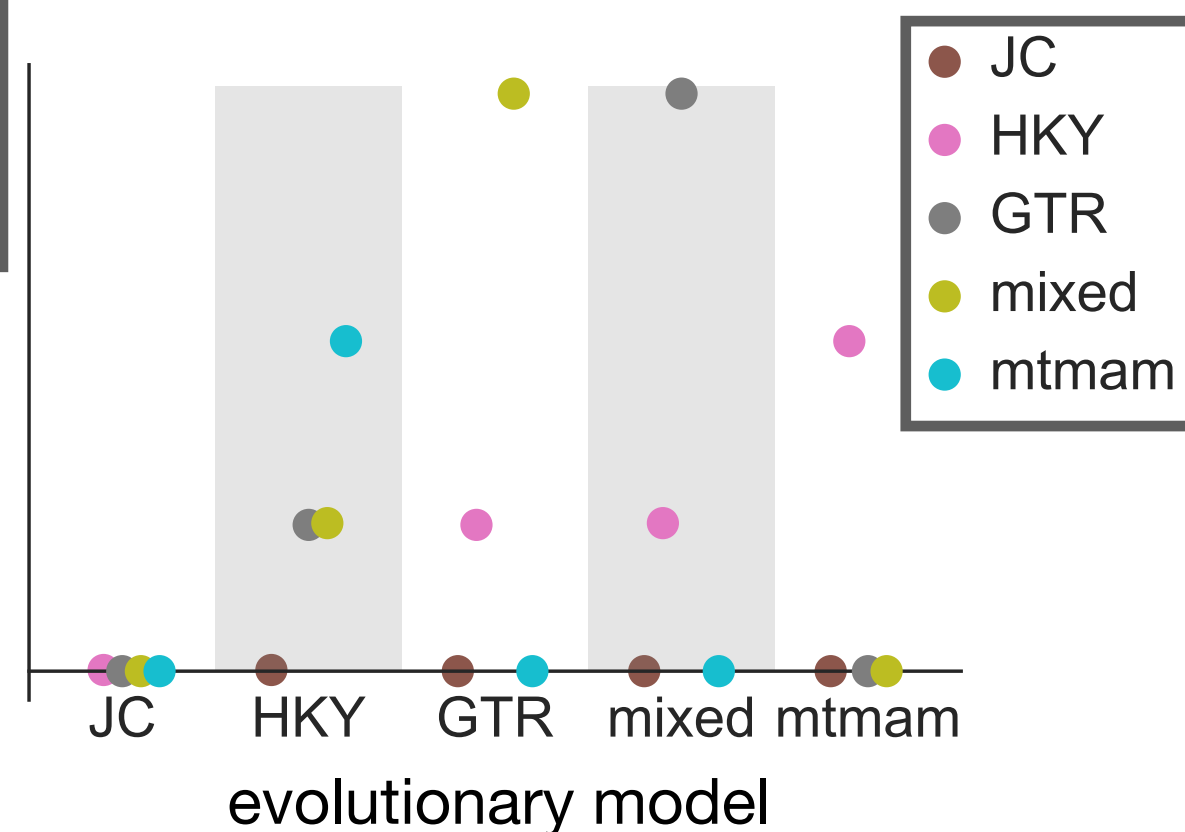
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- Same problem comparing evolutionary models with data fixed

Cross-data consistency



Cross-model consistency



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- **This talk:** when and how contradictions can arise in
 1. model selection
 2. prediction with high-dimensional models
 3. unsupervised learning
- **Takeaways:**
 - A. Non-reproducibility can be **subtle** (and is **problem-dependent**)
 - B. Not specific to Bayes
 - C. Need default, low-cost **protective methods** that remain **statistically efficient**

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all posterior mass on a single, arbitrary model

This is not an abstract concern...

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...We show that ***model misspecification explains the variation in previous estimates of divergence time***

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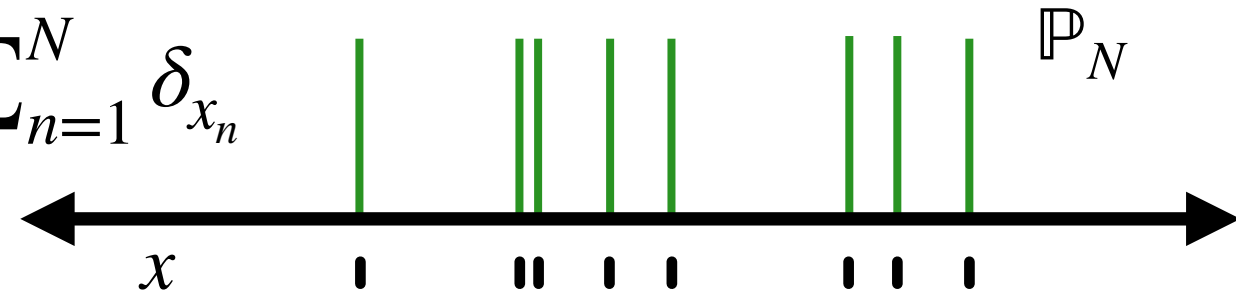
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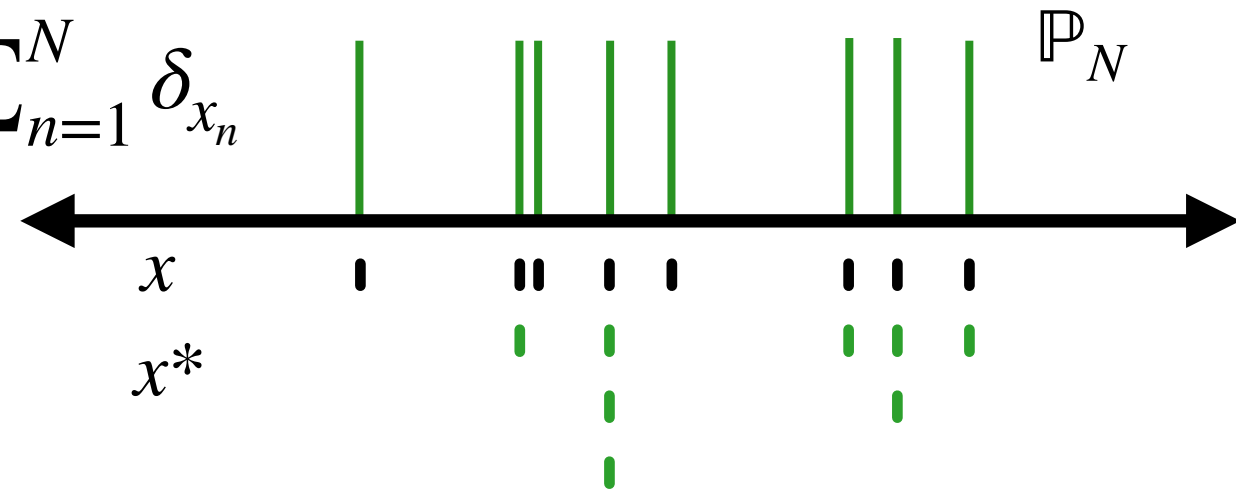
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- Bootstrap dataset $x^* = (x_1^*, \dots, x_M^*)$,
where x_m^* i.i.d. $\sim \mathbb{P}_N$



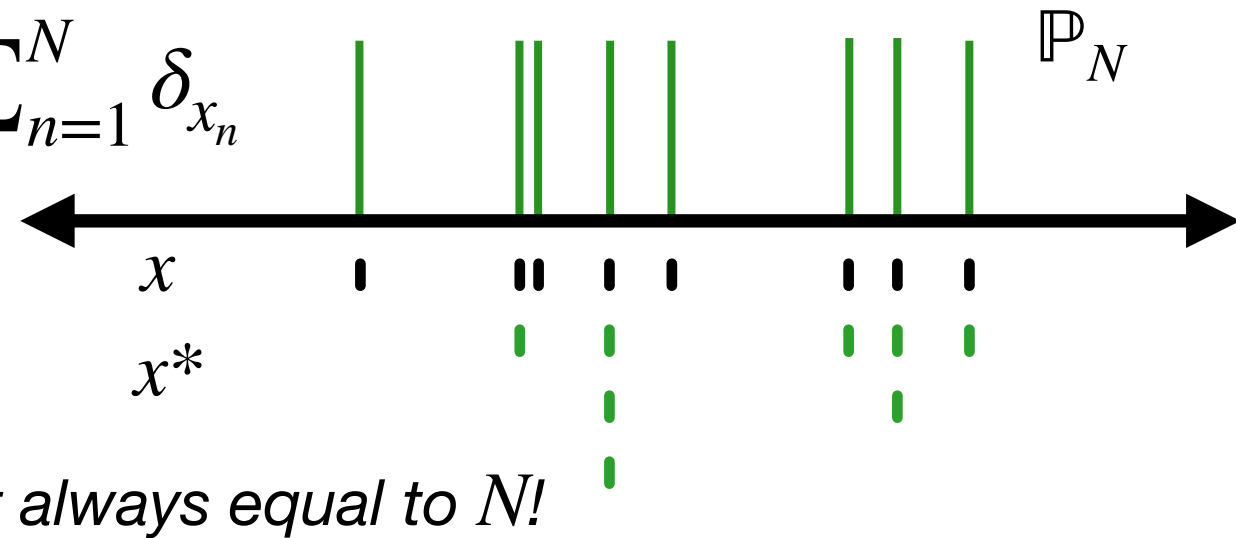
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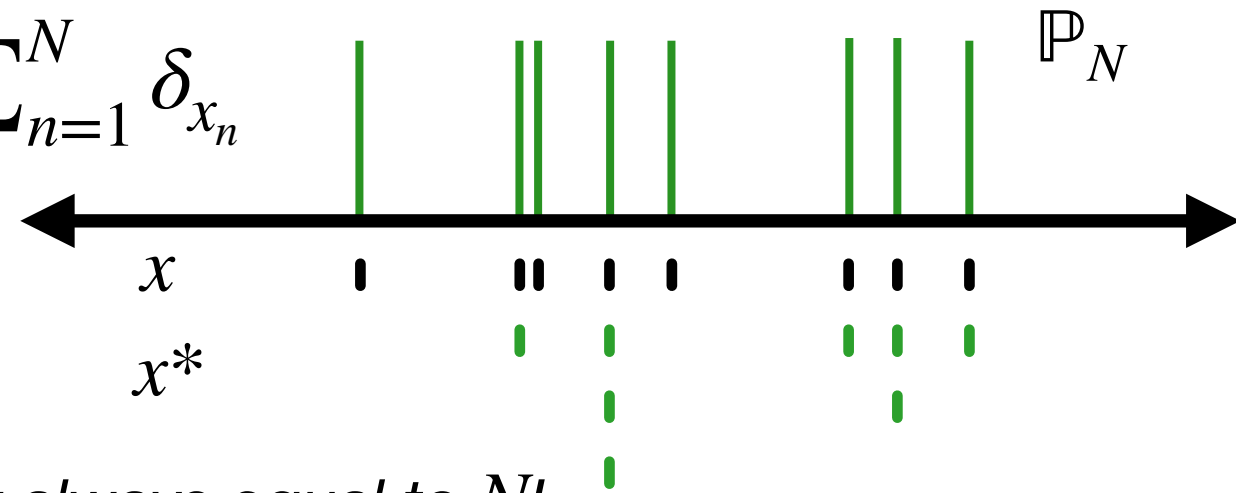
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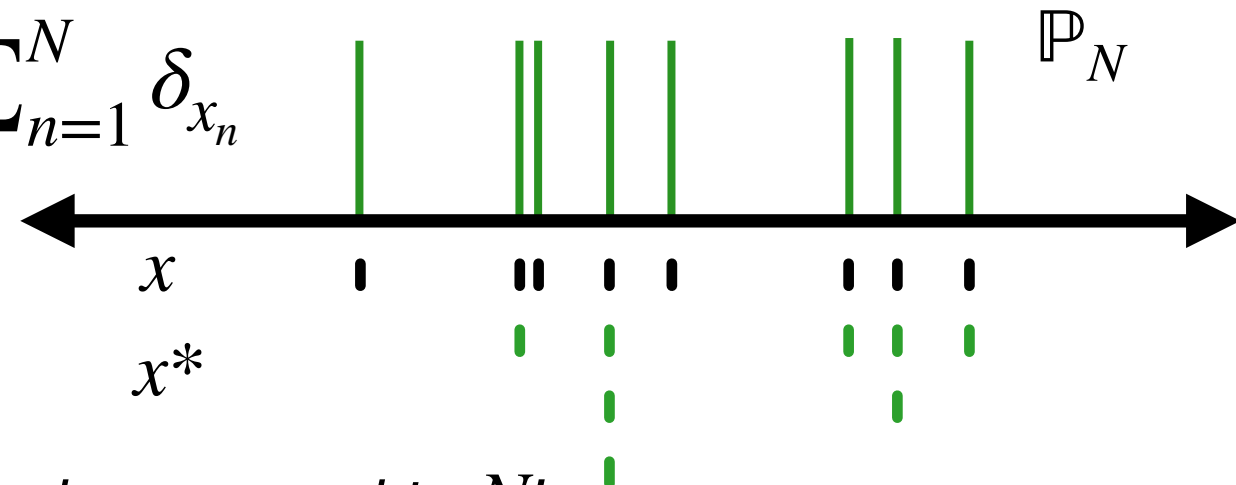
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- In practice, sample B bootstrap datasets: $\pi^*(\theta | x) \approx \sum_{b=1}^B \pi(\theta | x_{(b)}^*)$

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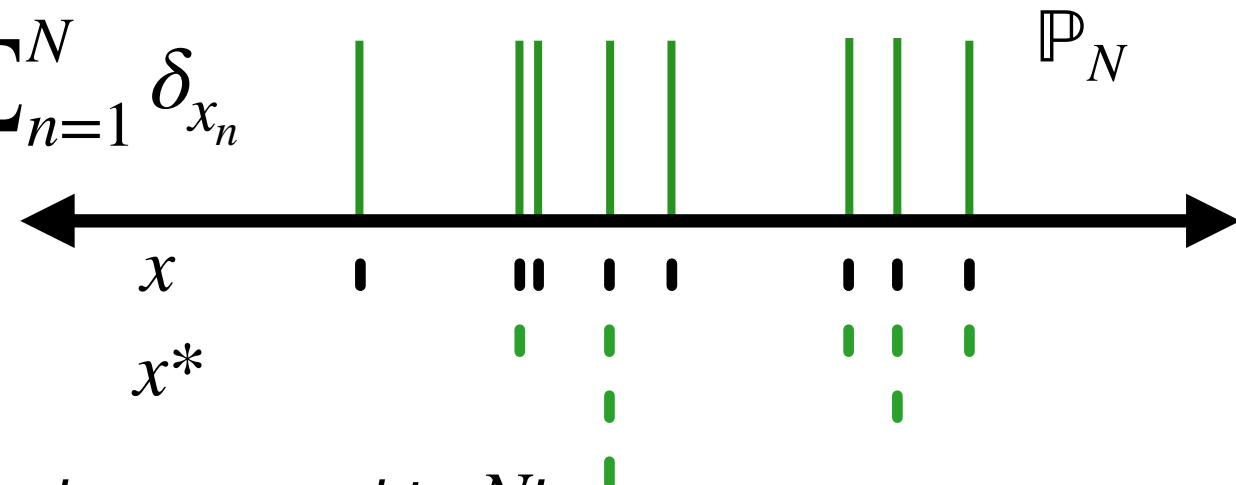
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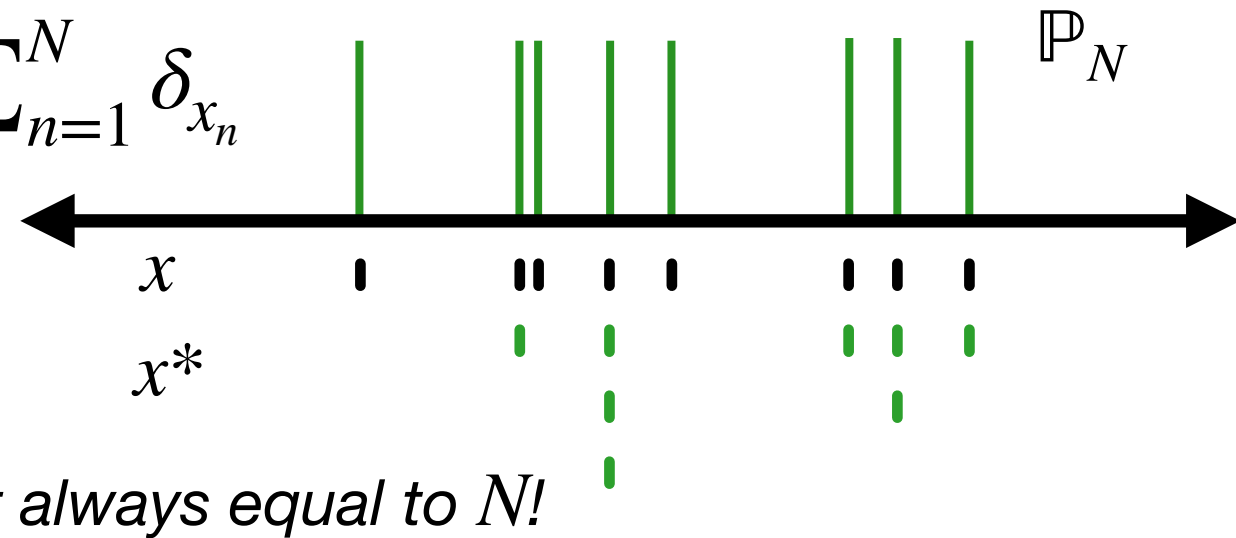
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▸ Benefits: **easy to use**, can **parallelize** across B

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- Assume data x and two models m_1 and m_2
- Assume they explain the data-generating distribution equally well:
 $\mathbb{E}\{\log p(x | m_1)\} = \mathbb{E}\{\log p(x | m_2)\}$
- We'd **hope** models have equal poster probability as $N \rightarrow \infty$:
 $\pi(m_1 | Y) = \pi(m_2 | Y) = 1/2$

- However...

Theorem [Yang & Zhu 2019, H & Miller 2023]:

As $N \rightarrow \infty$, $\pi(m_1 | x) \xrightarrow{d} \text{Bernoulli}(0.5)$

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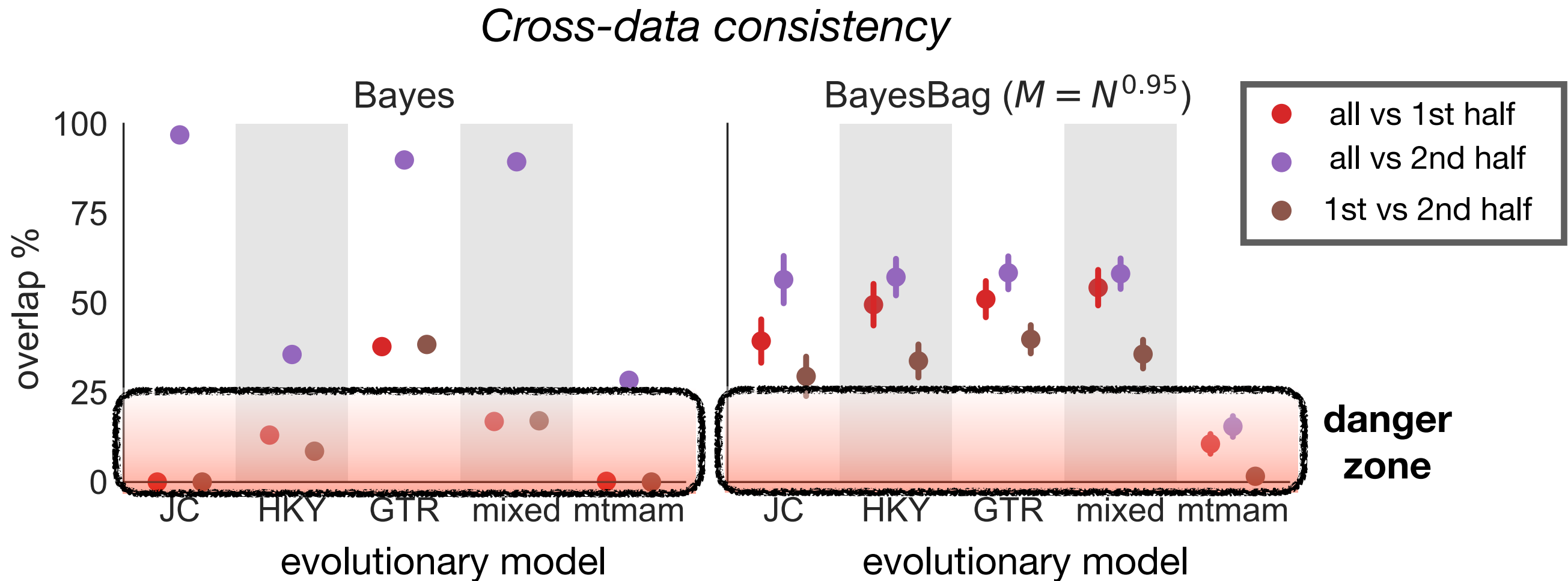
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Recommendation:

- ▶ $M = N^{0.95}$ default
- ▶ $M = N^{0.75}$ if significant misspecification and/or many models

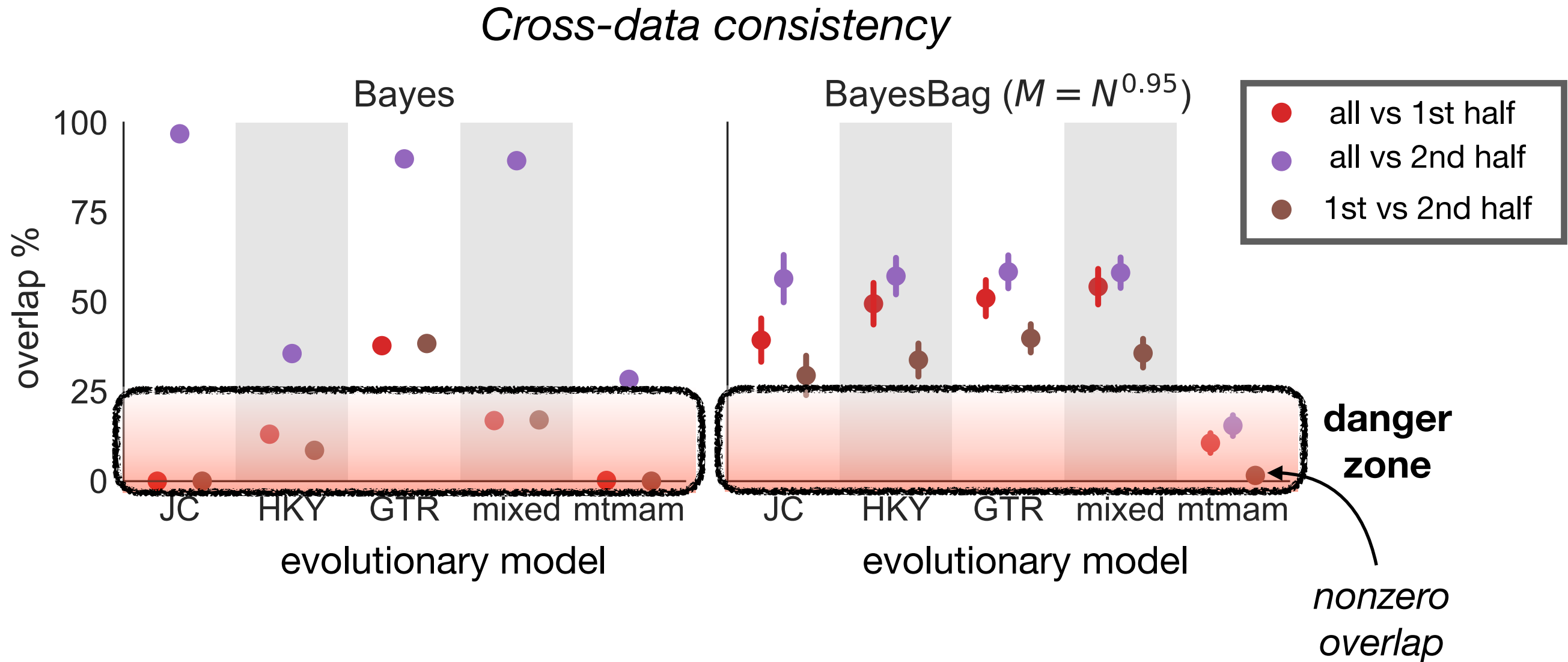
Bagged posterior improves stability

Goal: infer phylogeny of 13 whale species



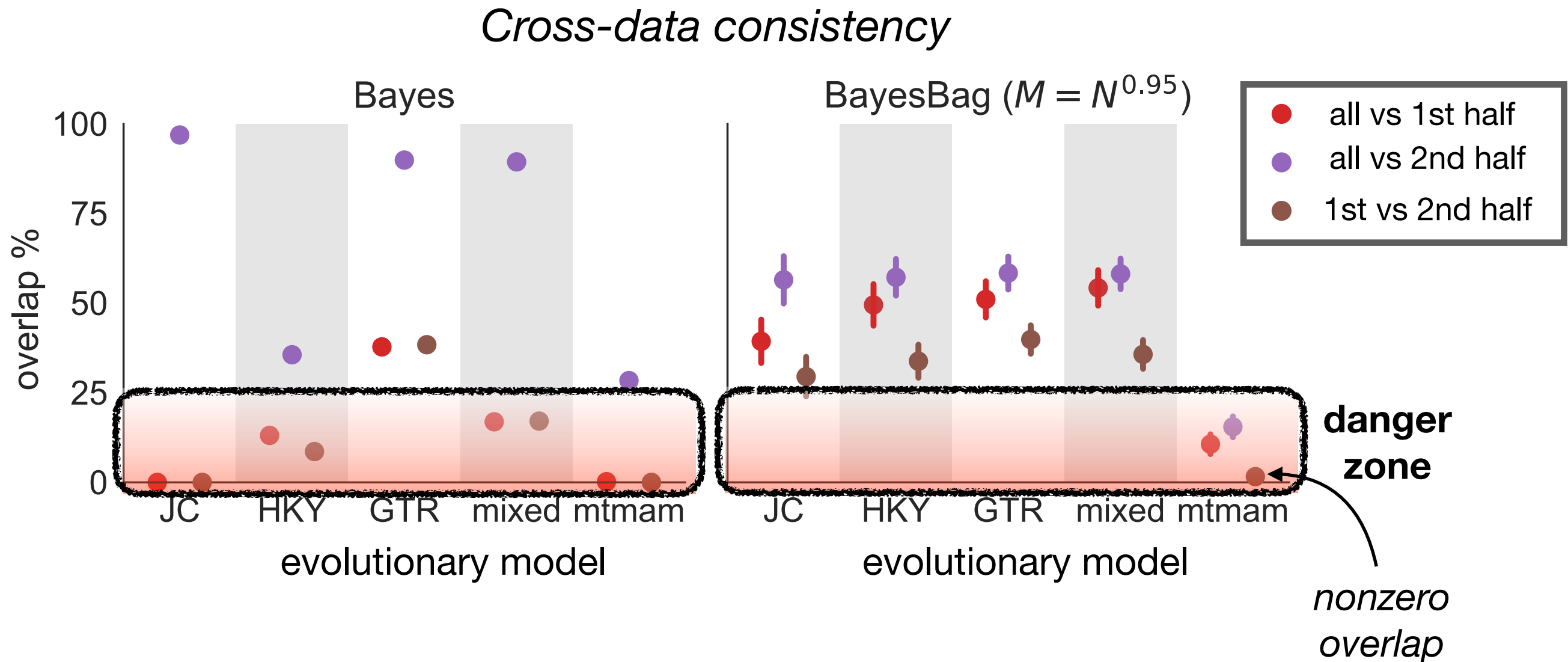
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- Significant overlap between non-MTMAM BayesBag and mixed Bayes
- BayesBag dramatically improves cross-model consistency too

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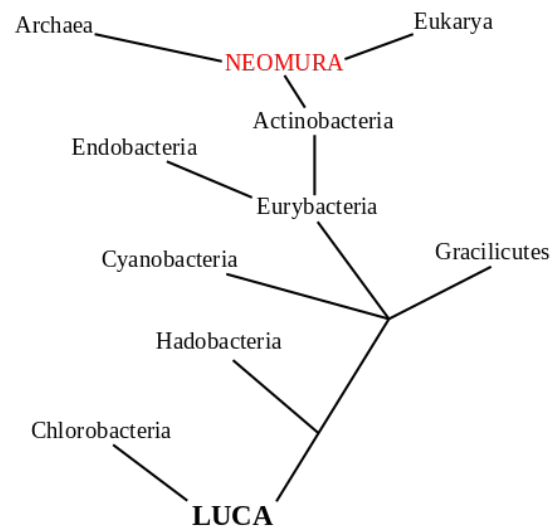
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Part 3: Unsupervised Learning of *Bona Fide* Latent Structure

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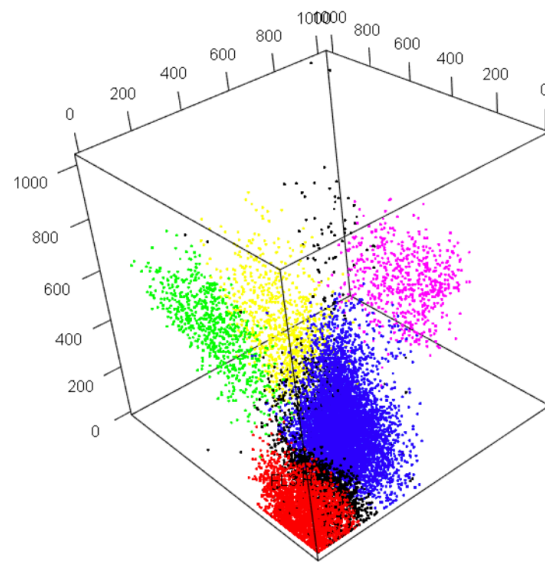
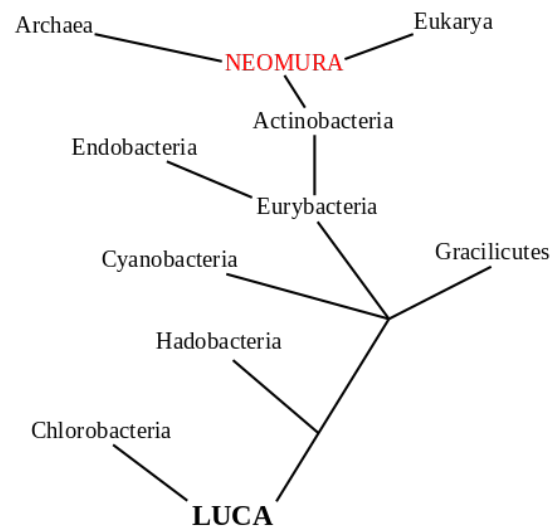
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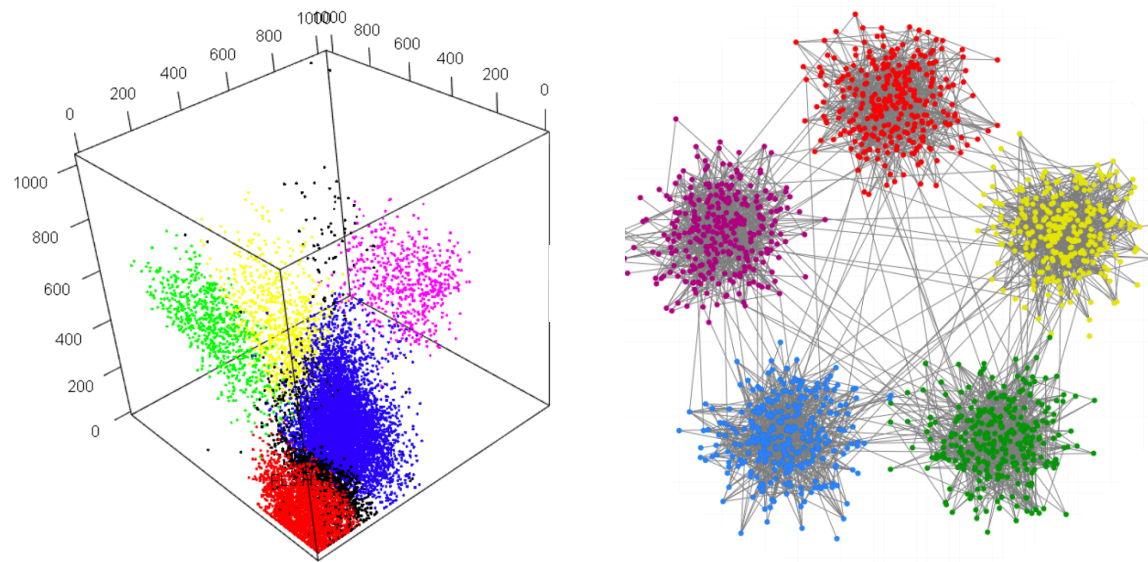
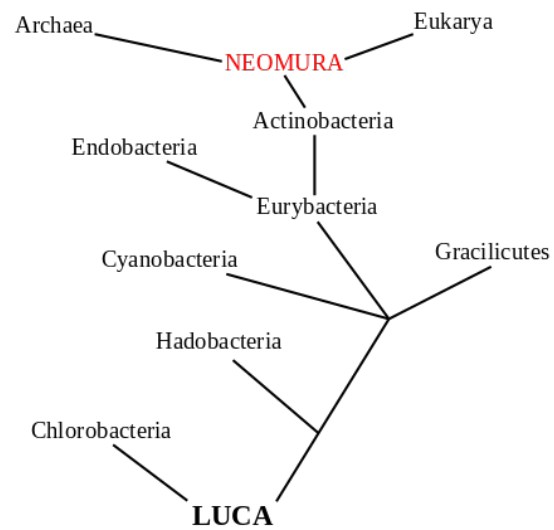
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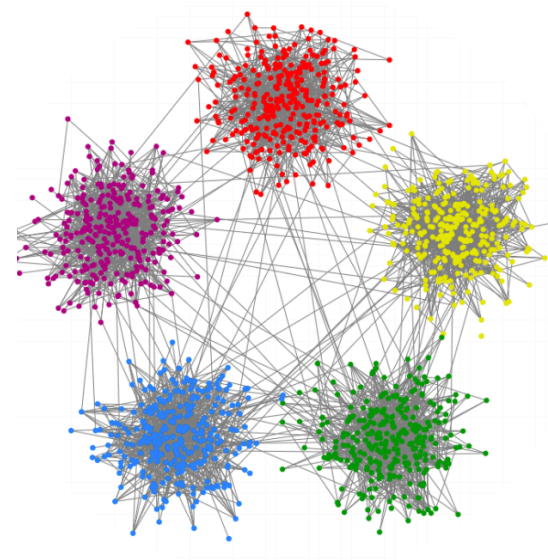
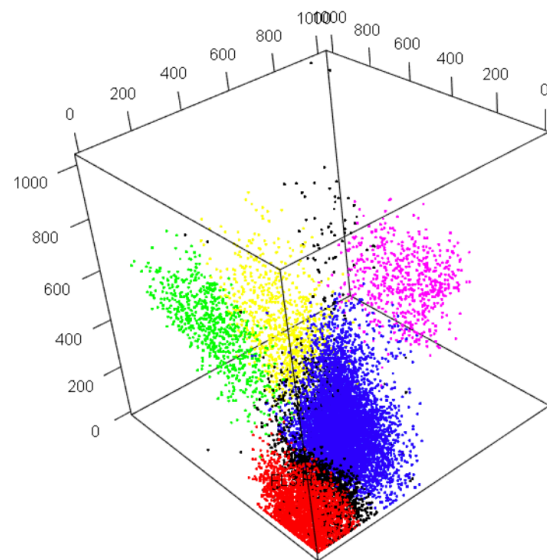
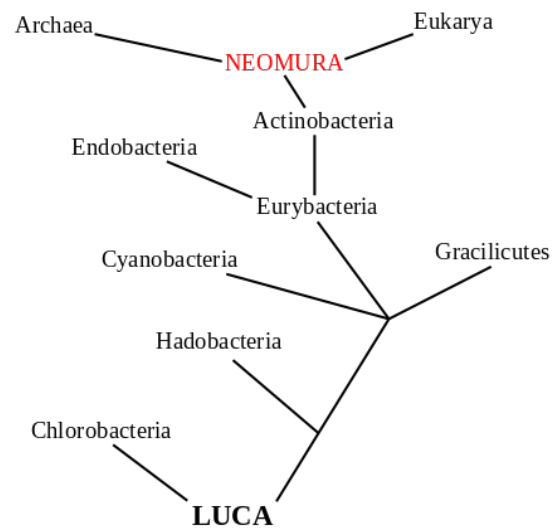
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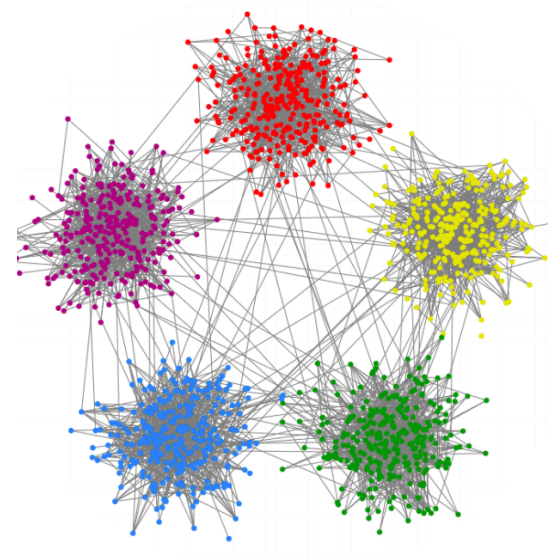
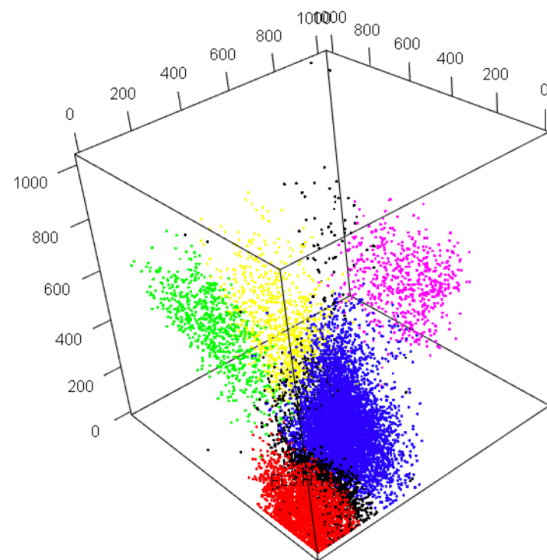
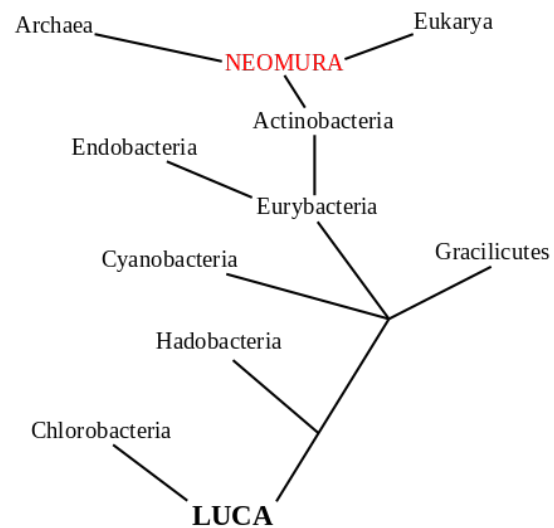
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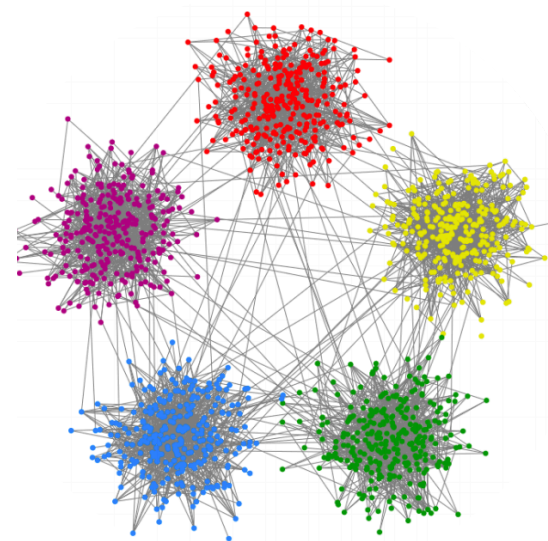
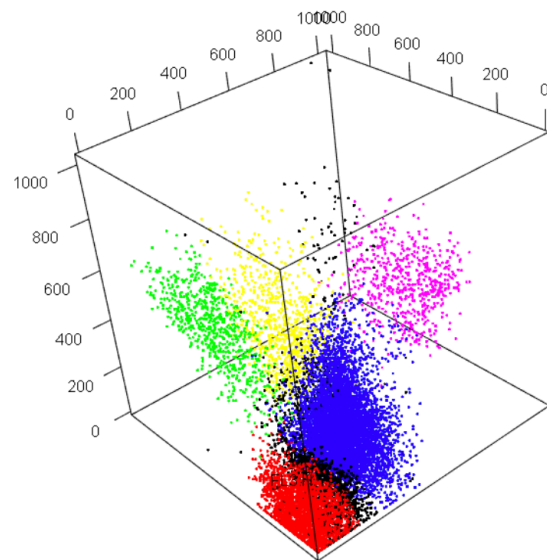
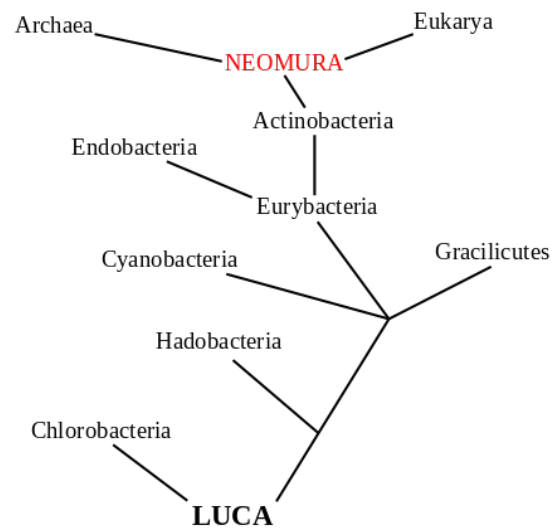
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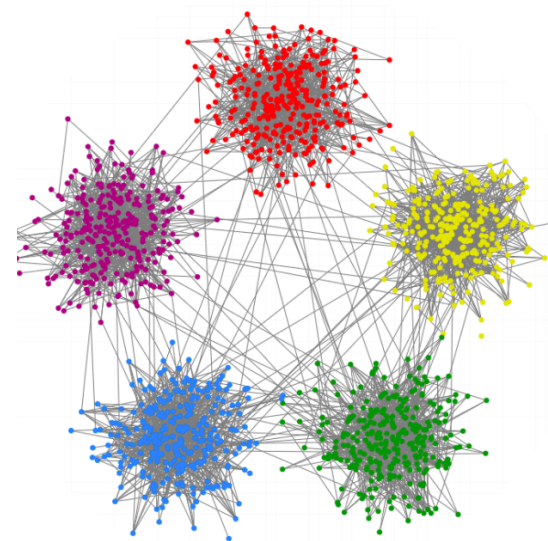
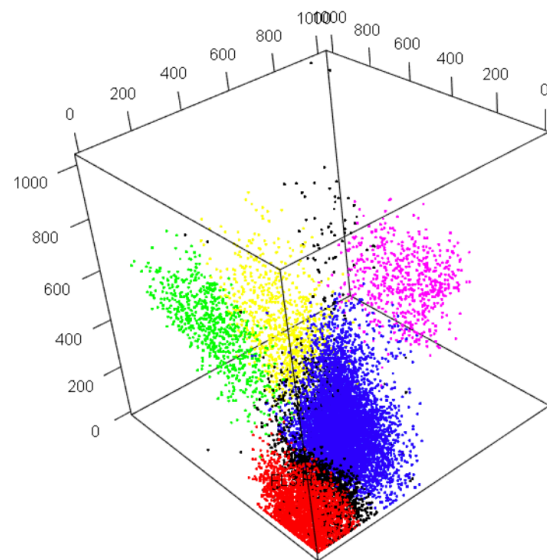
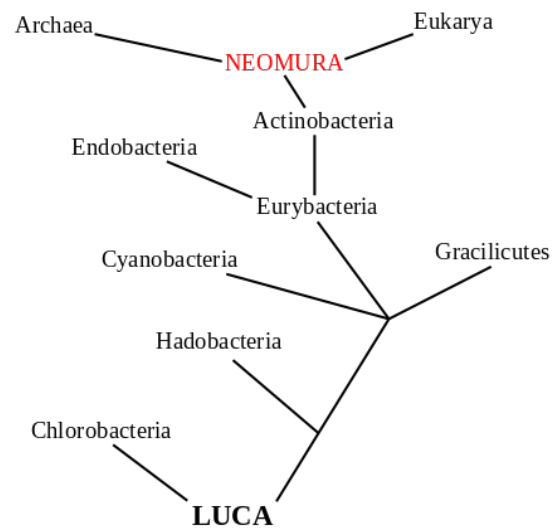
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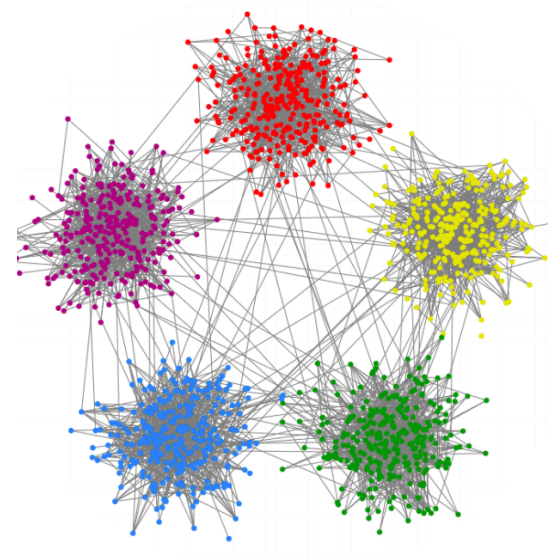
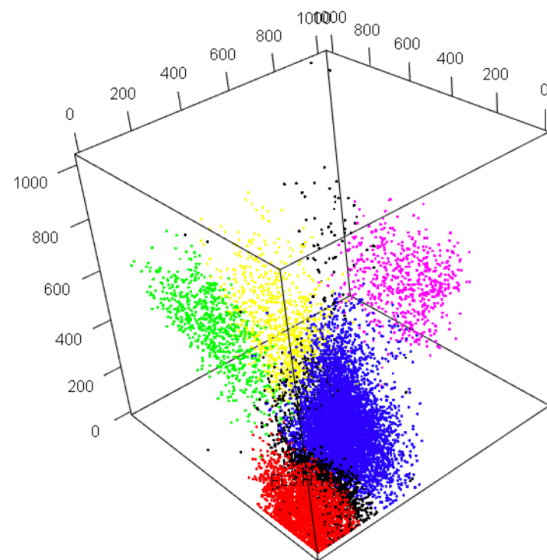
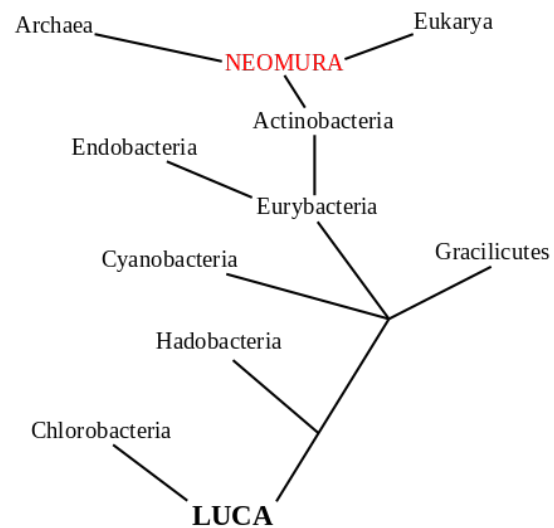
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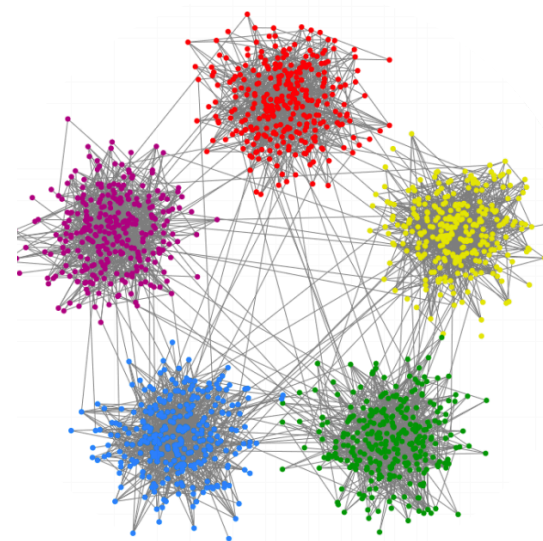
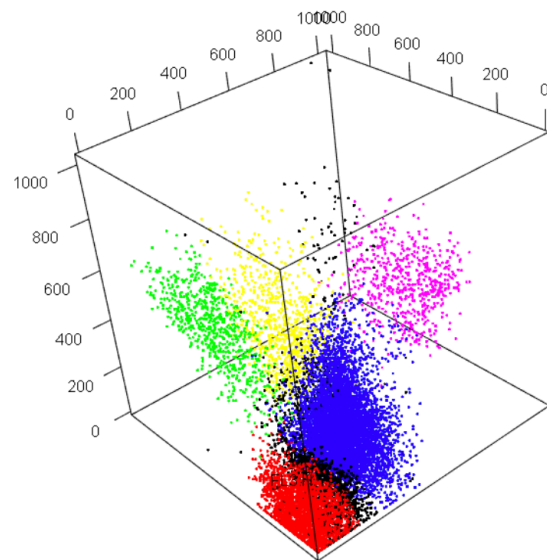
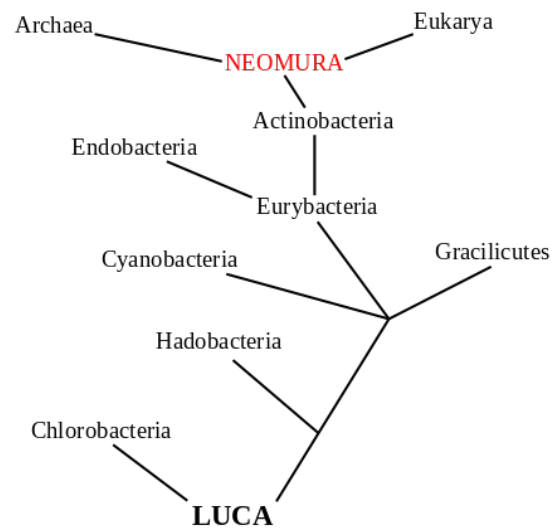
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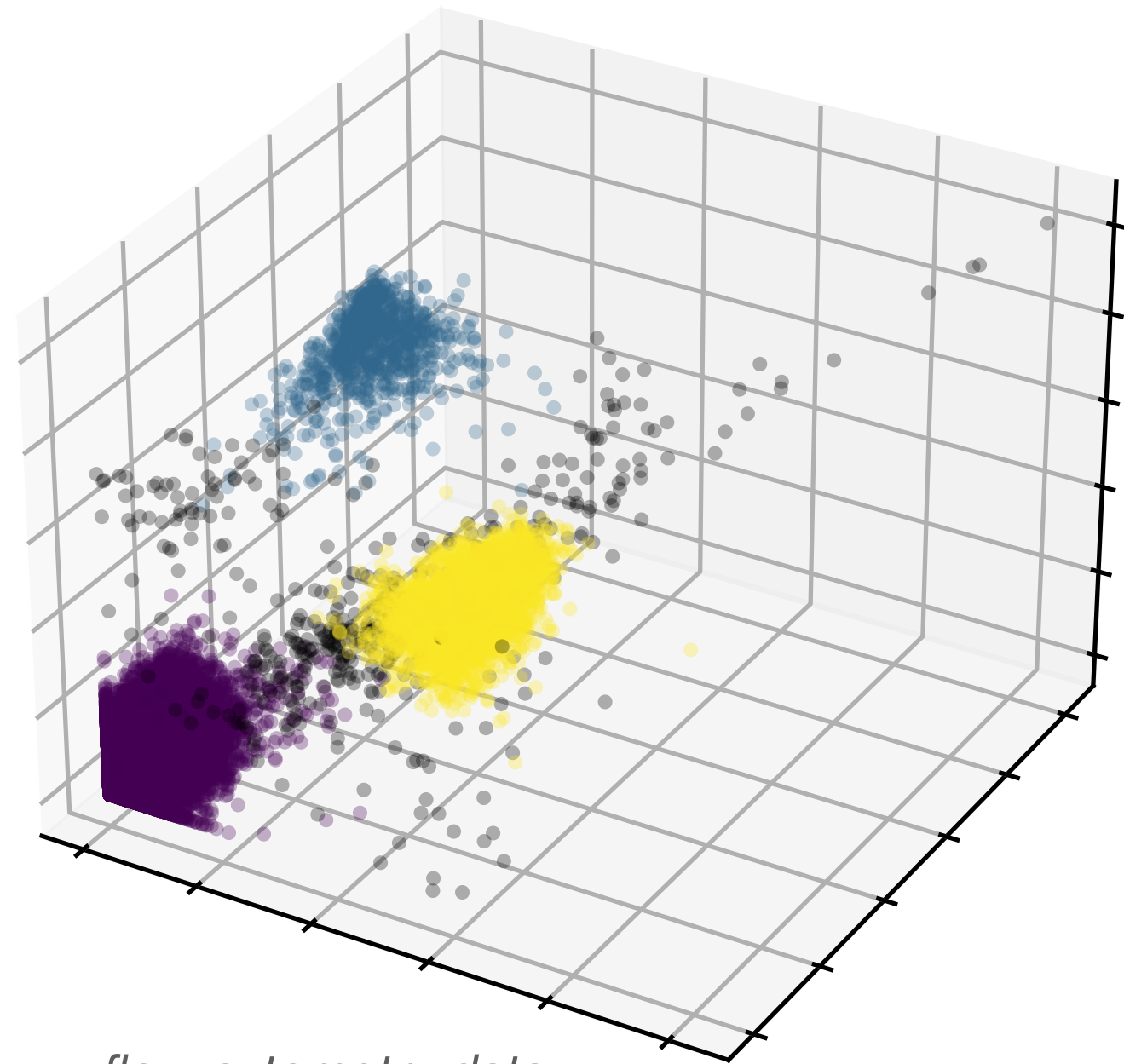
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- **Key ideas:** use known causal structure and domain knowledge

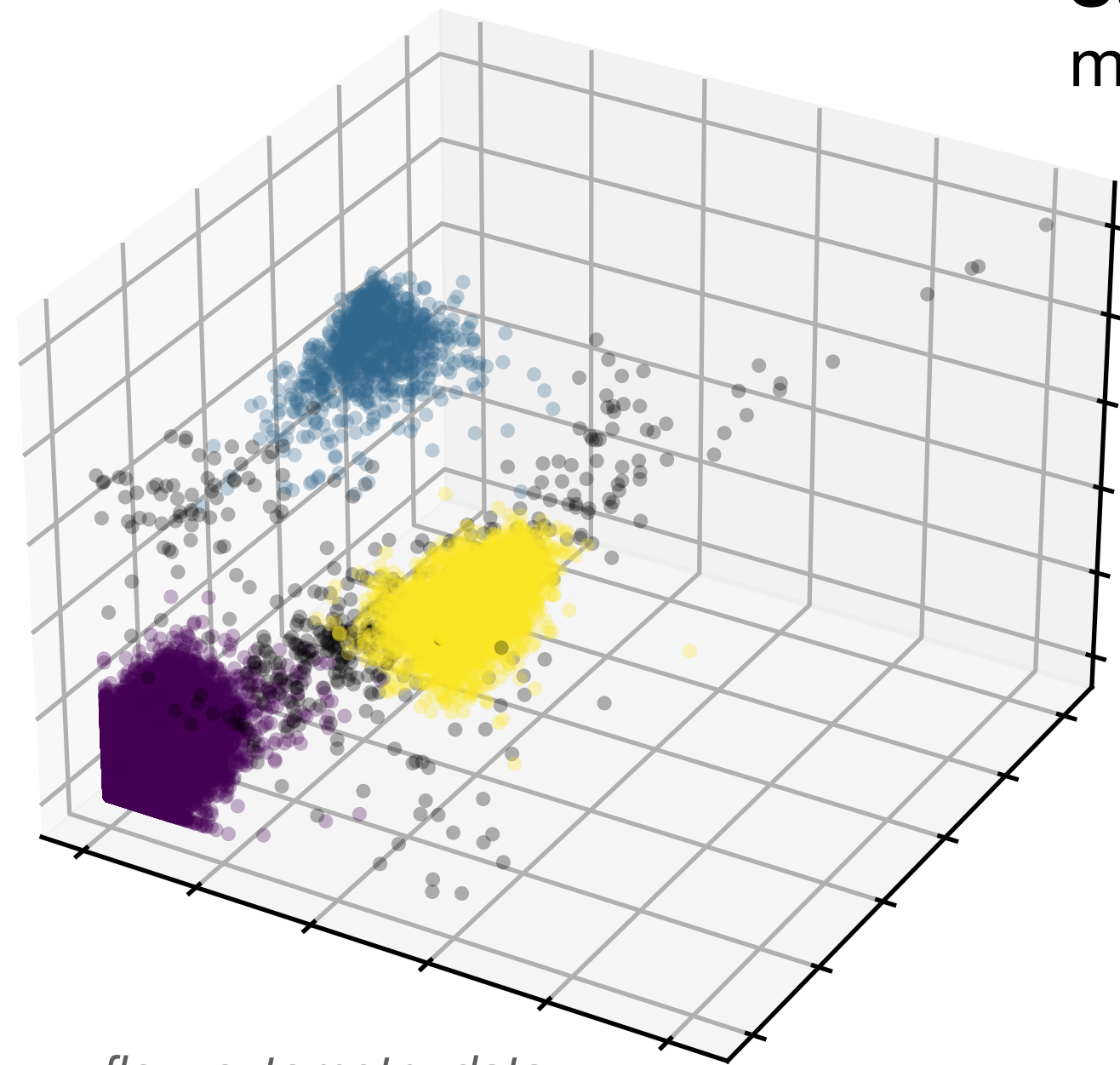
Clustering cells by type



flow cytometry data

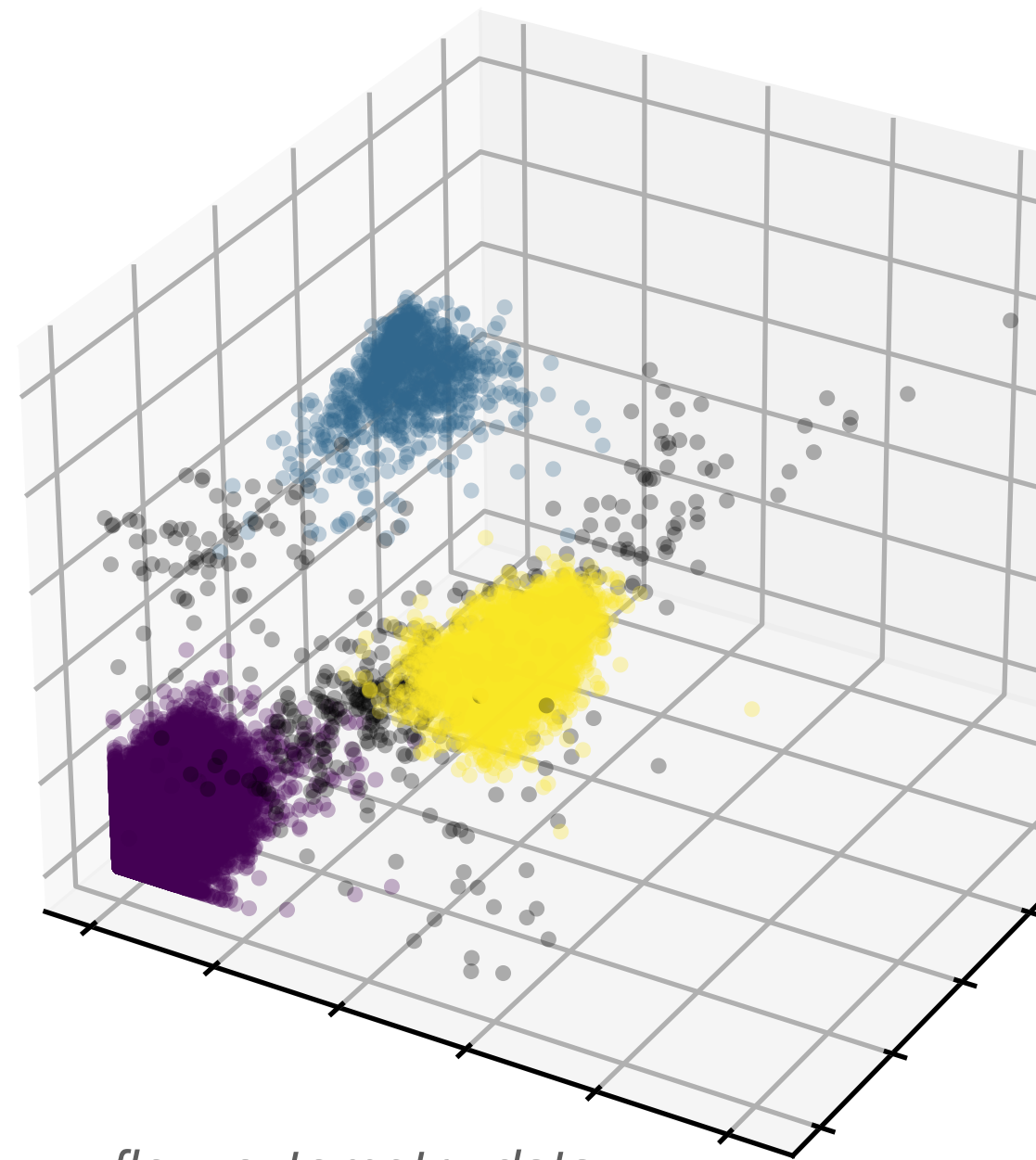
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Standard approach: Gaussian mixture model with prior on # of components k



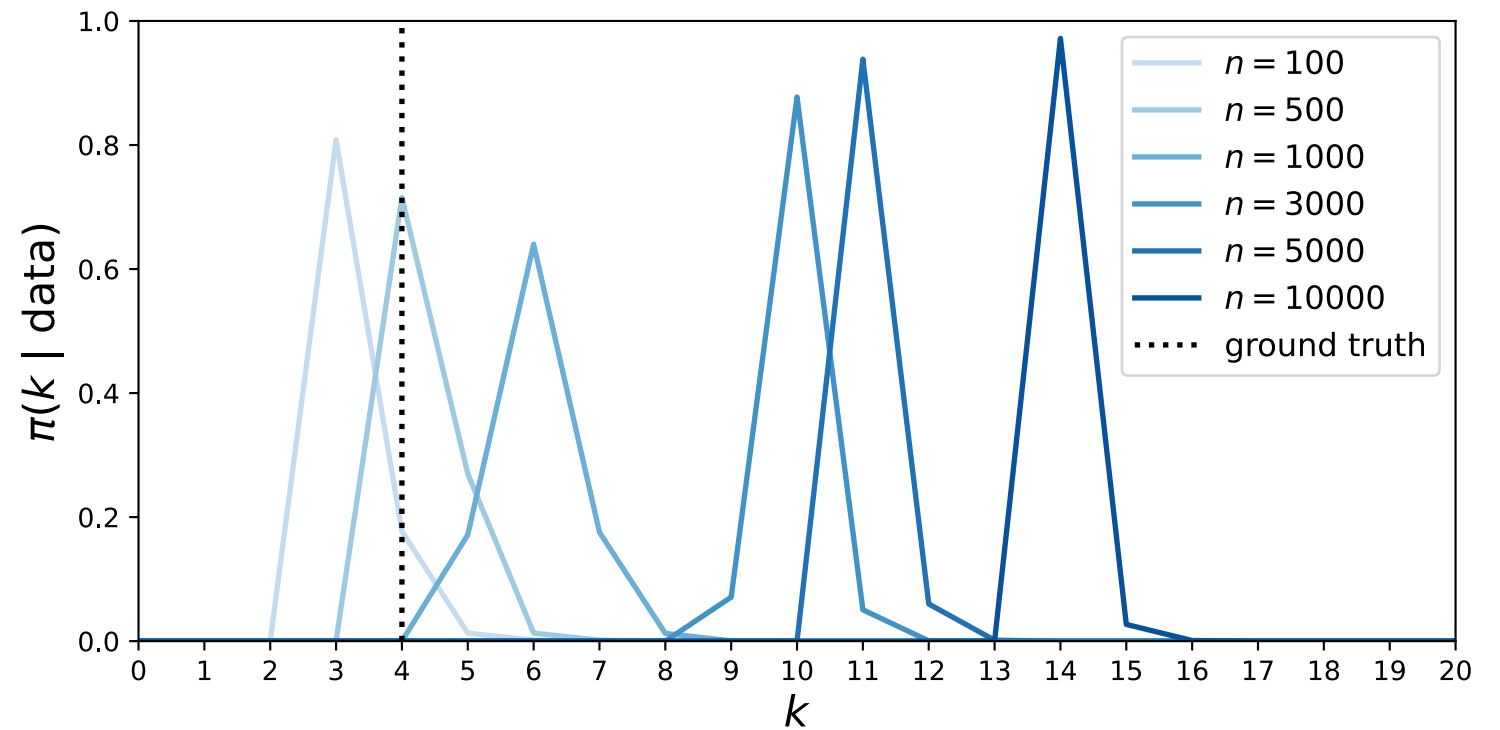
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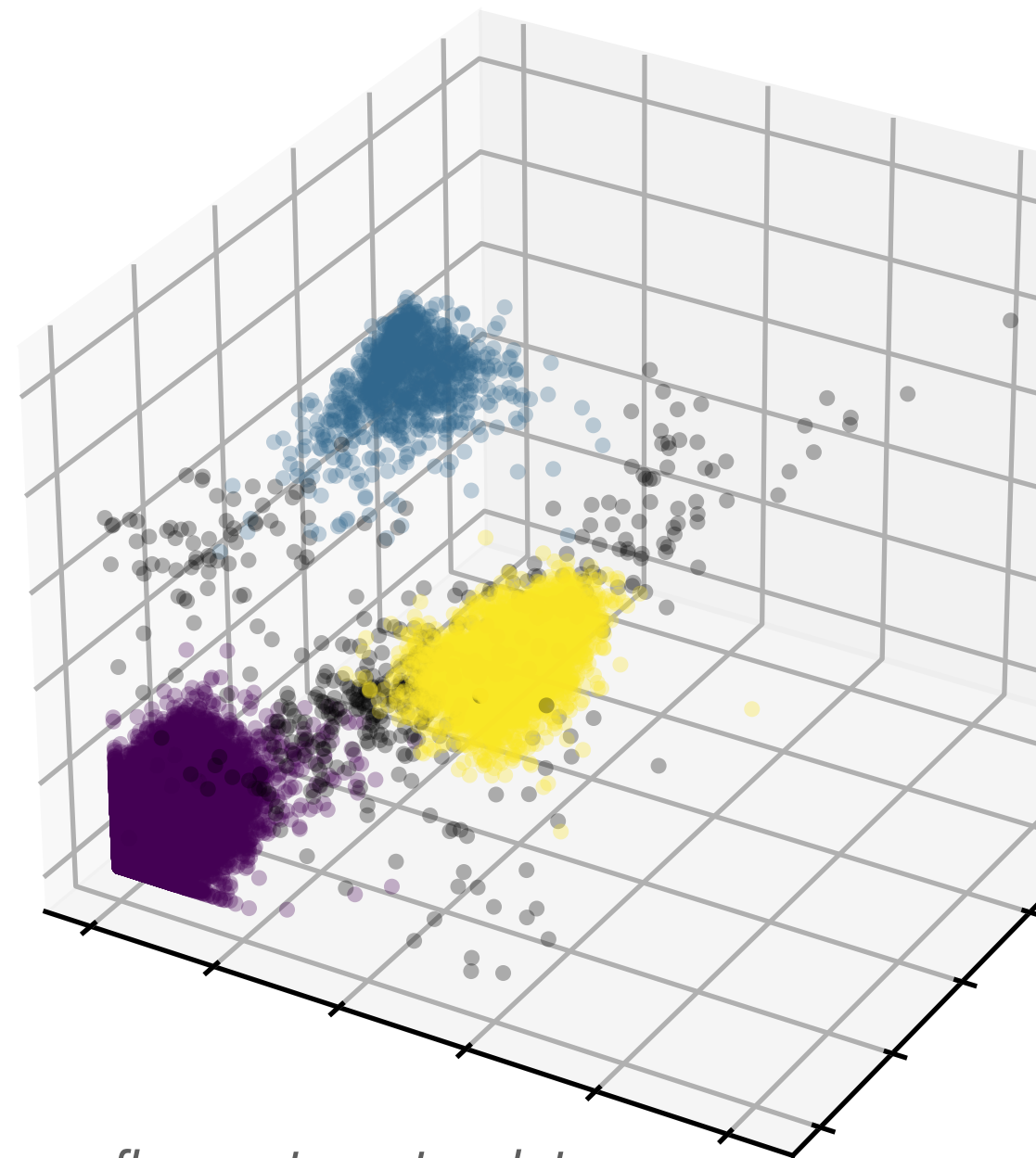


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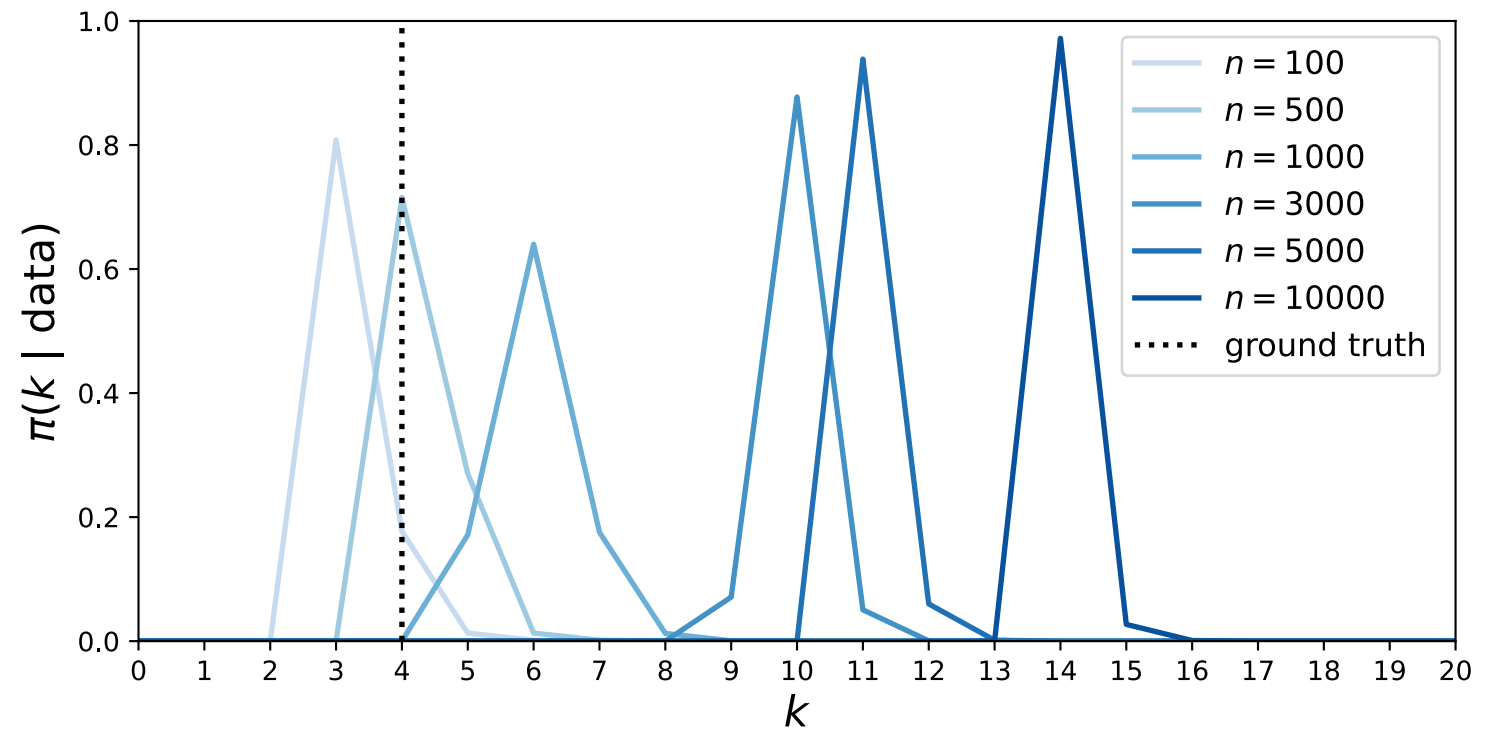


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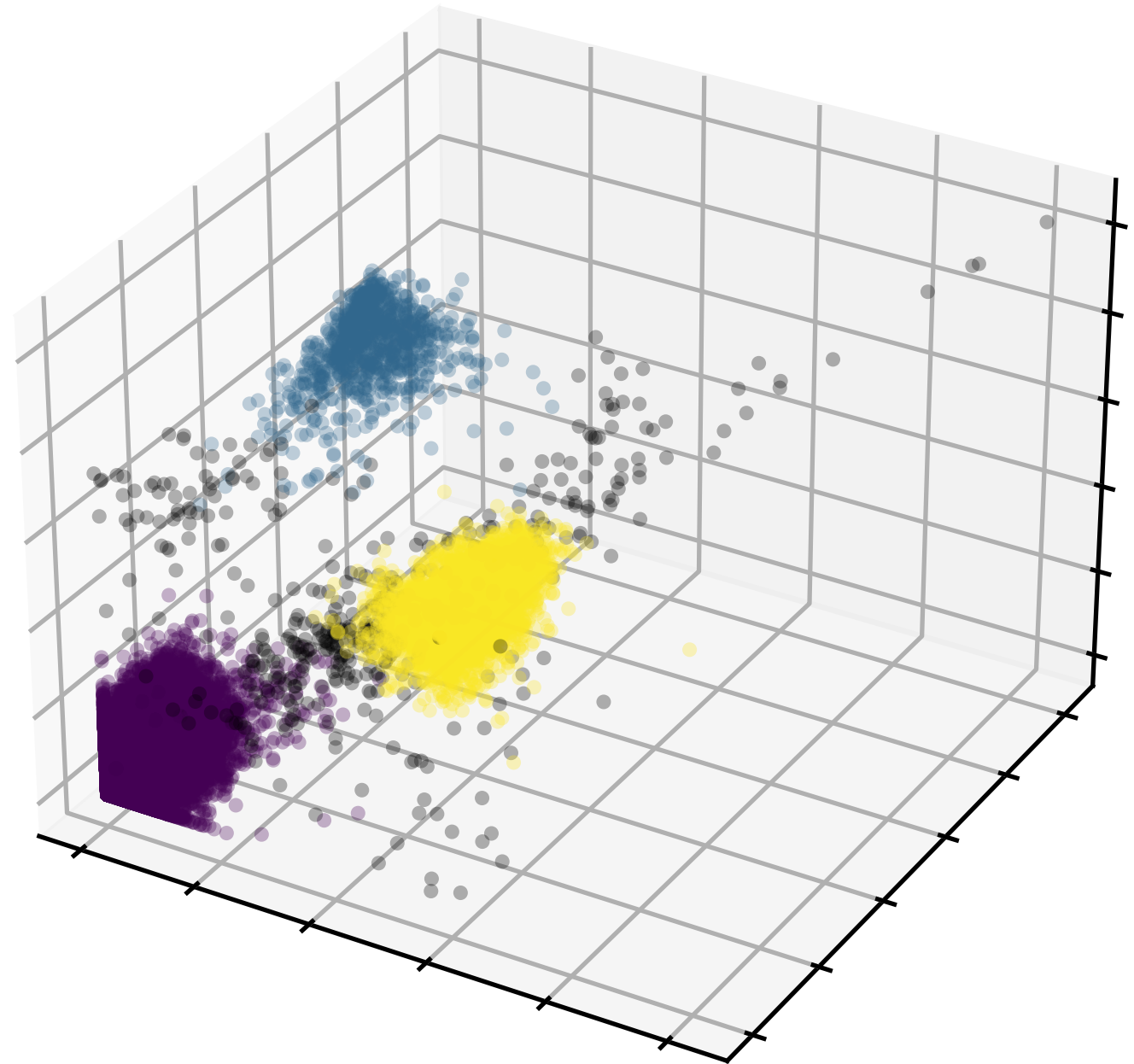
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More data, more contradictions!

Clustering for type discovery

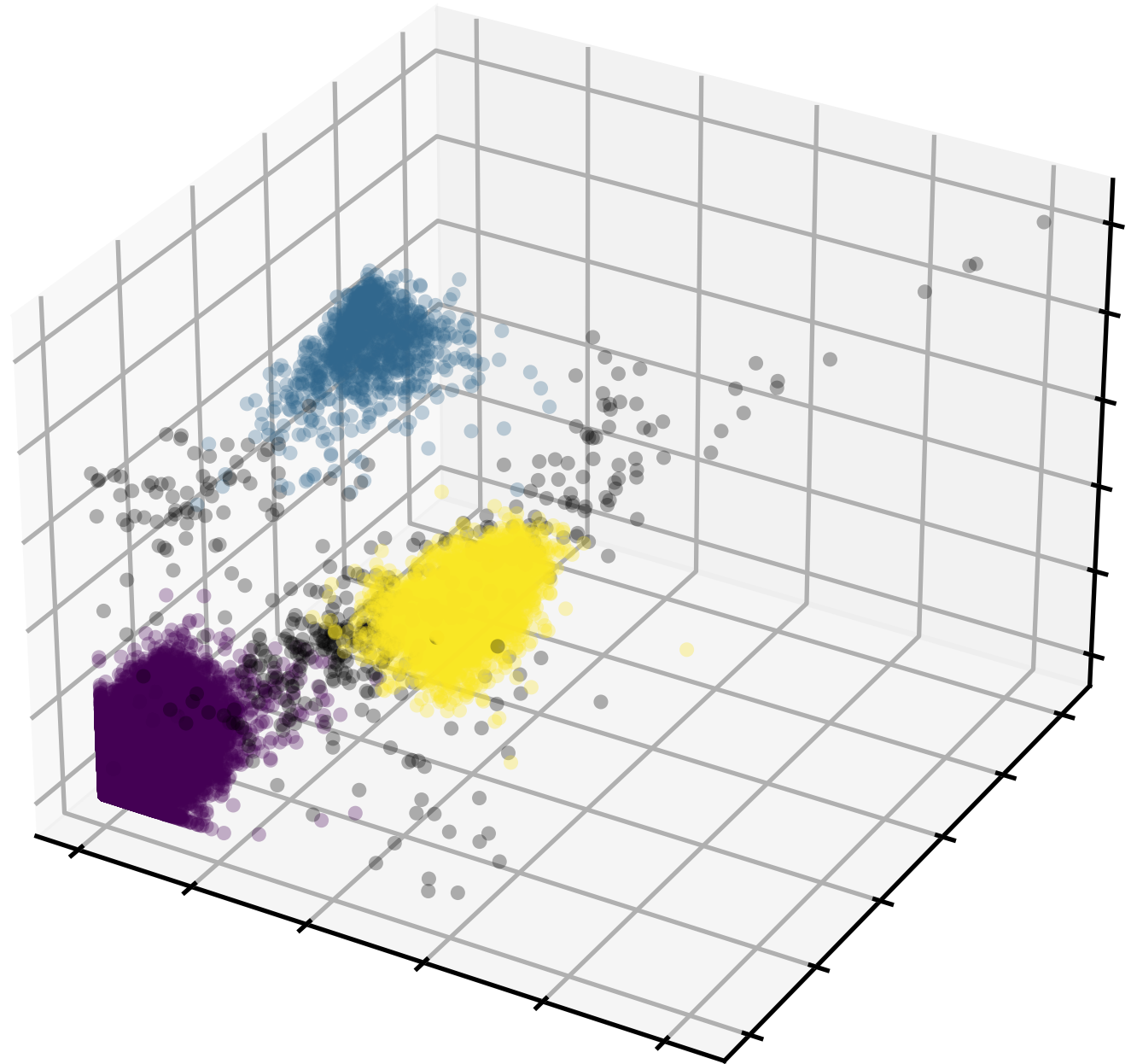


Clustering for type discovery

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- ▶ Each component is a meaningful *type*

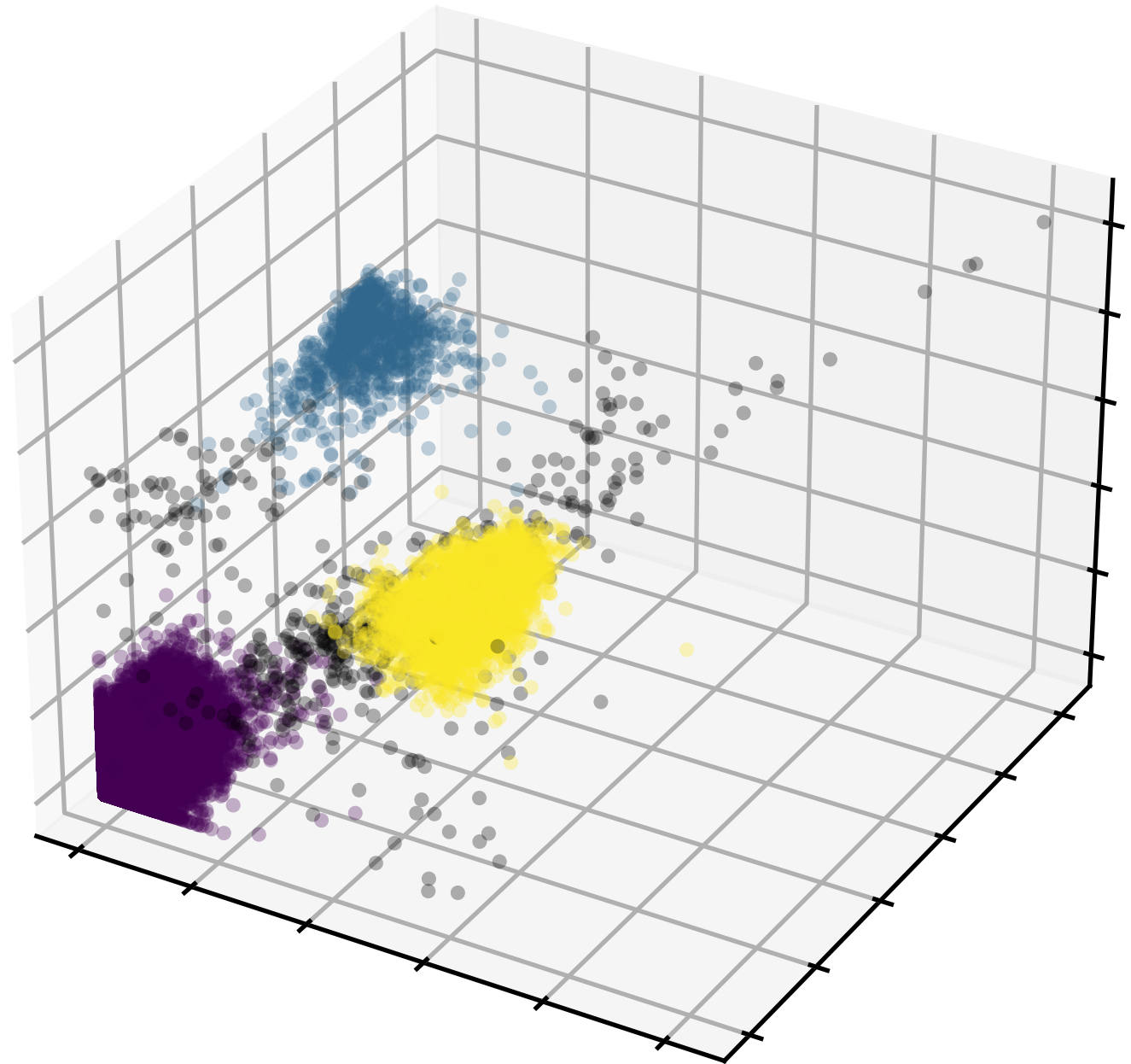


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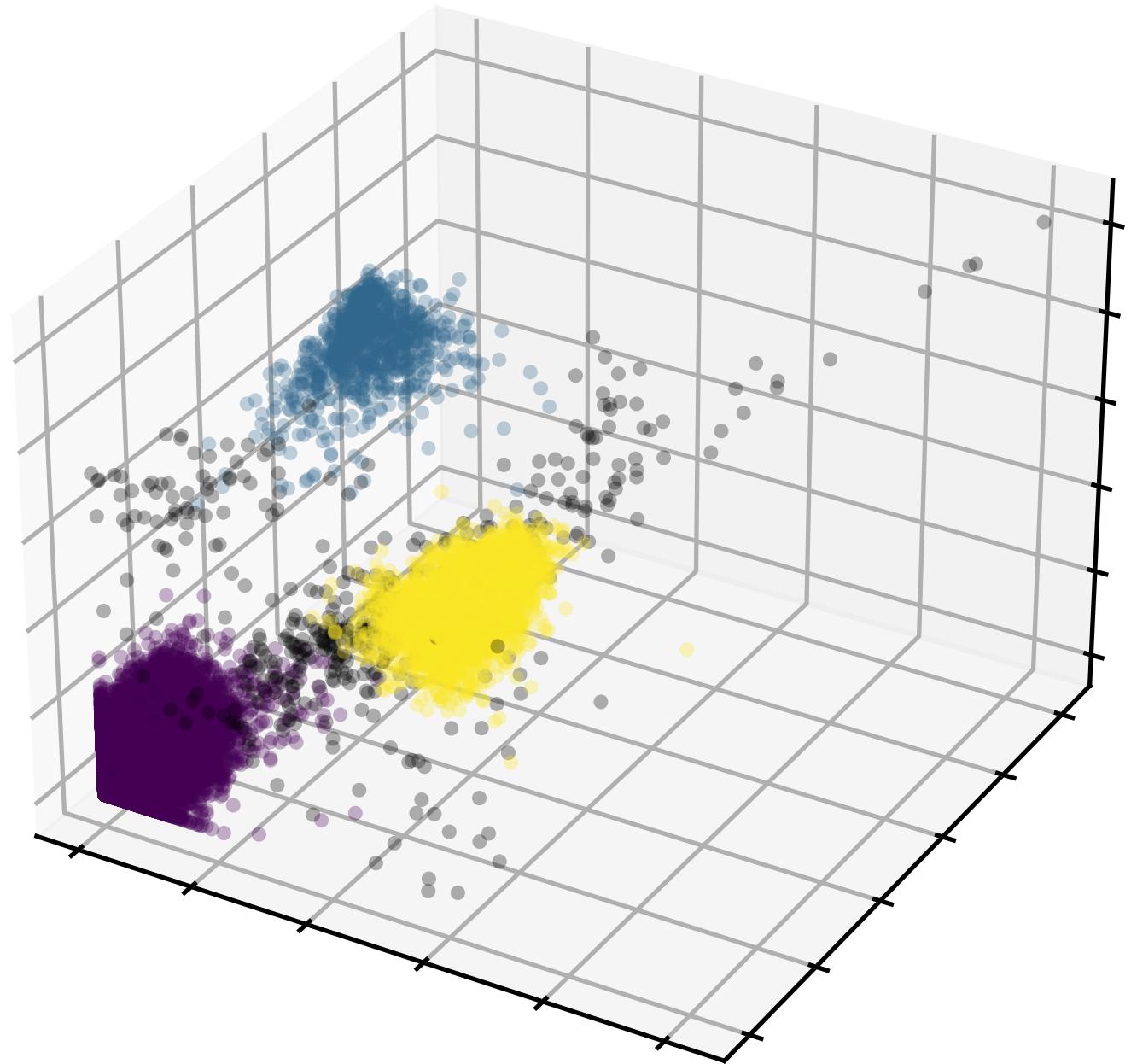
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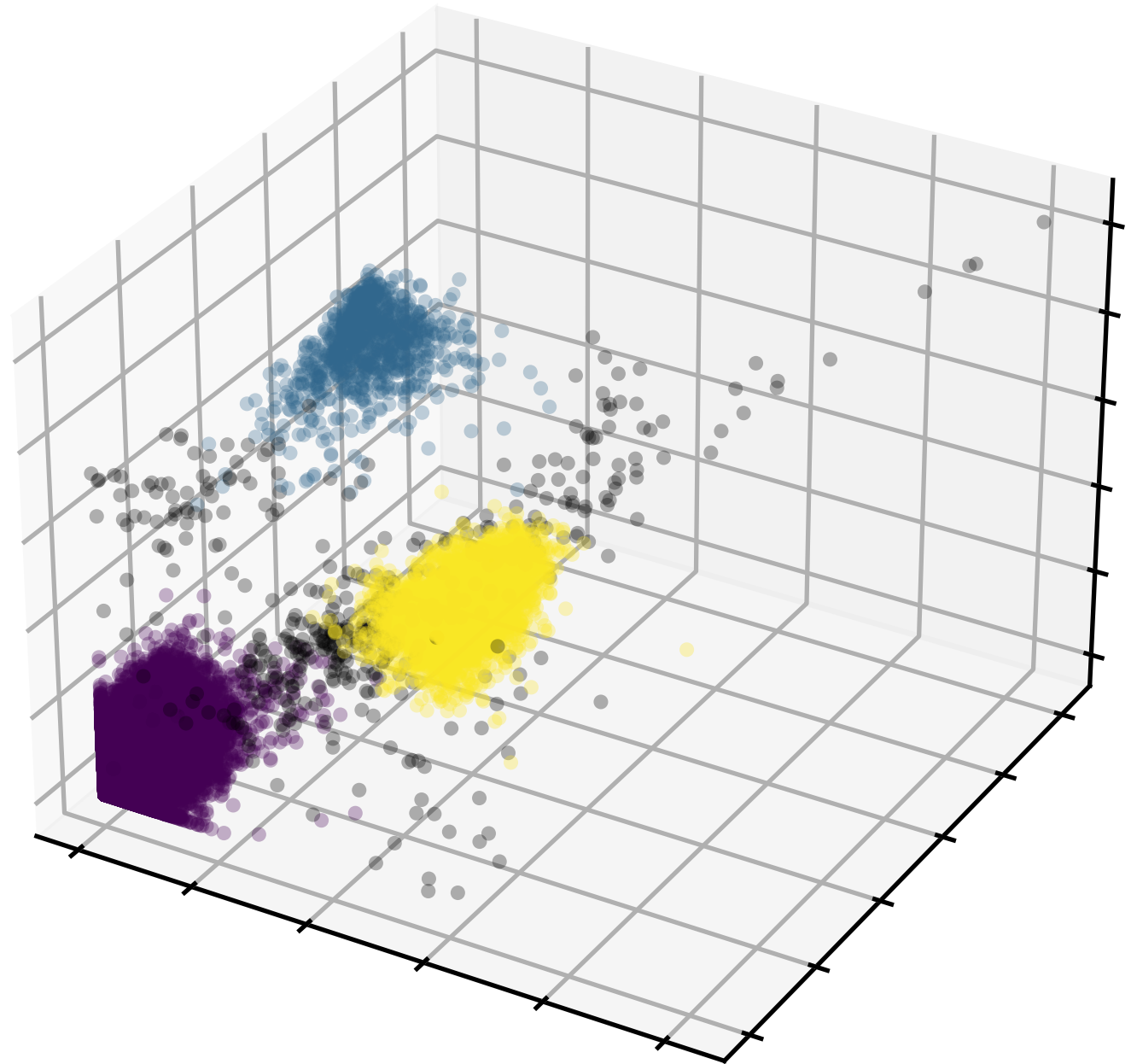
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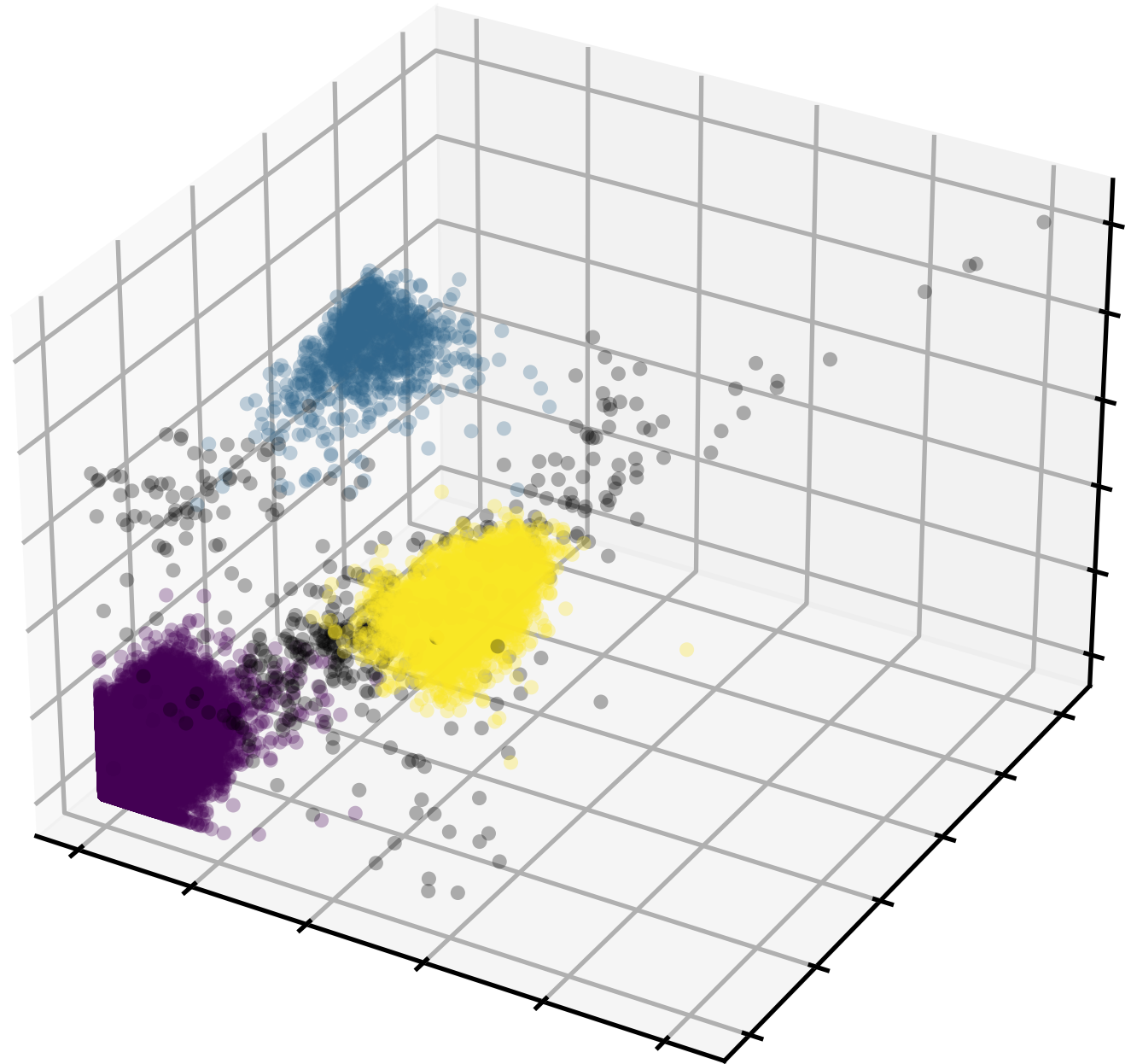
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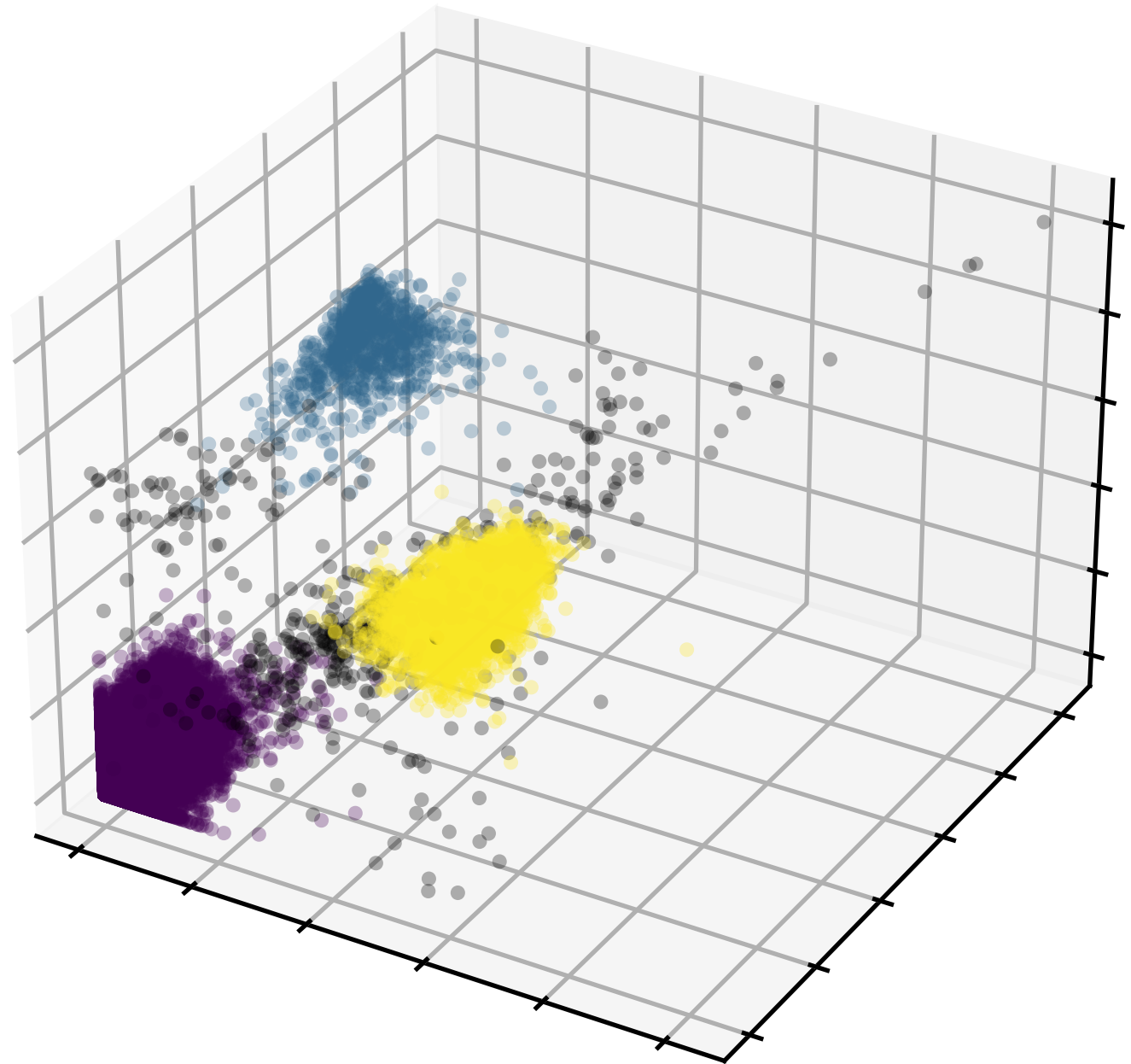
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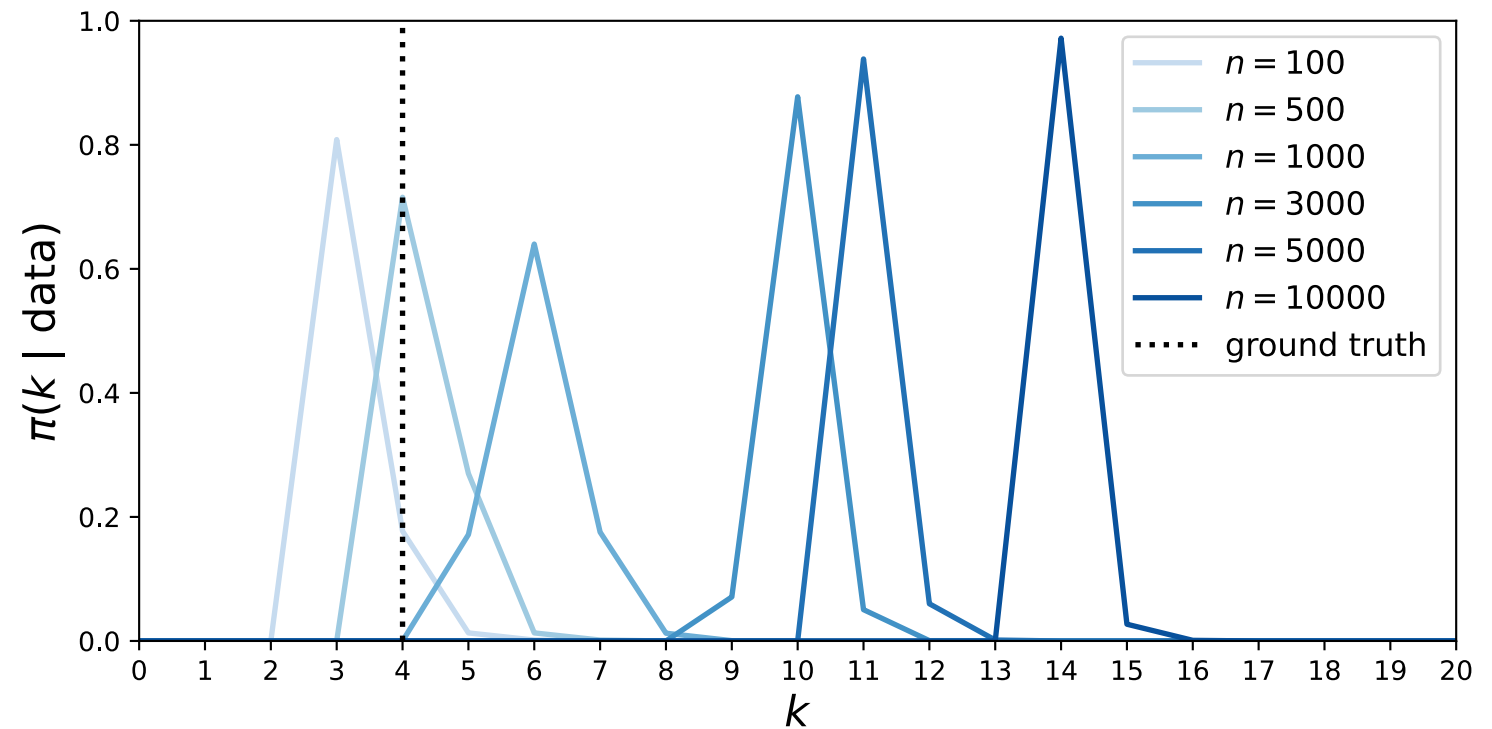
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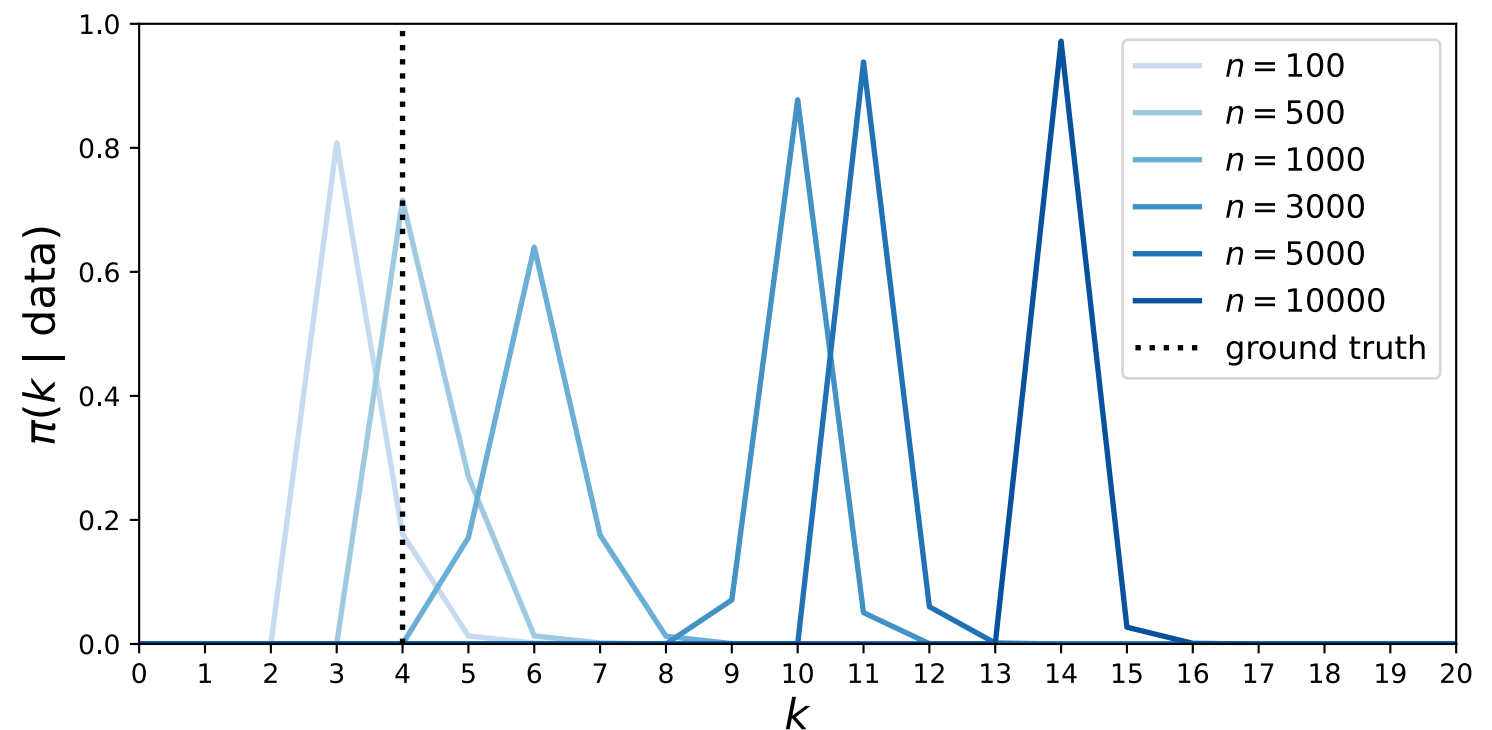


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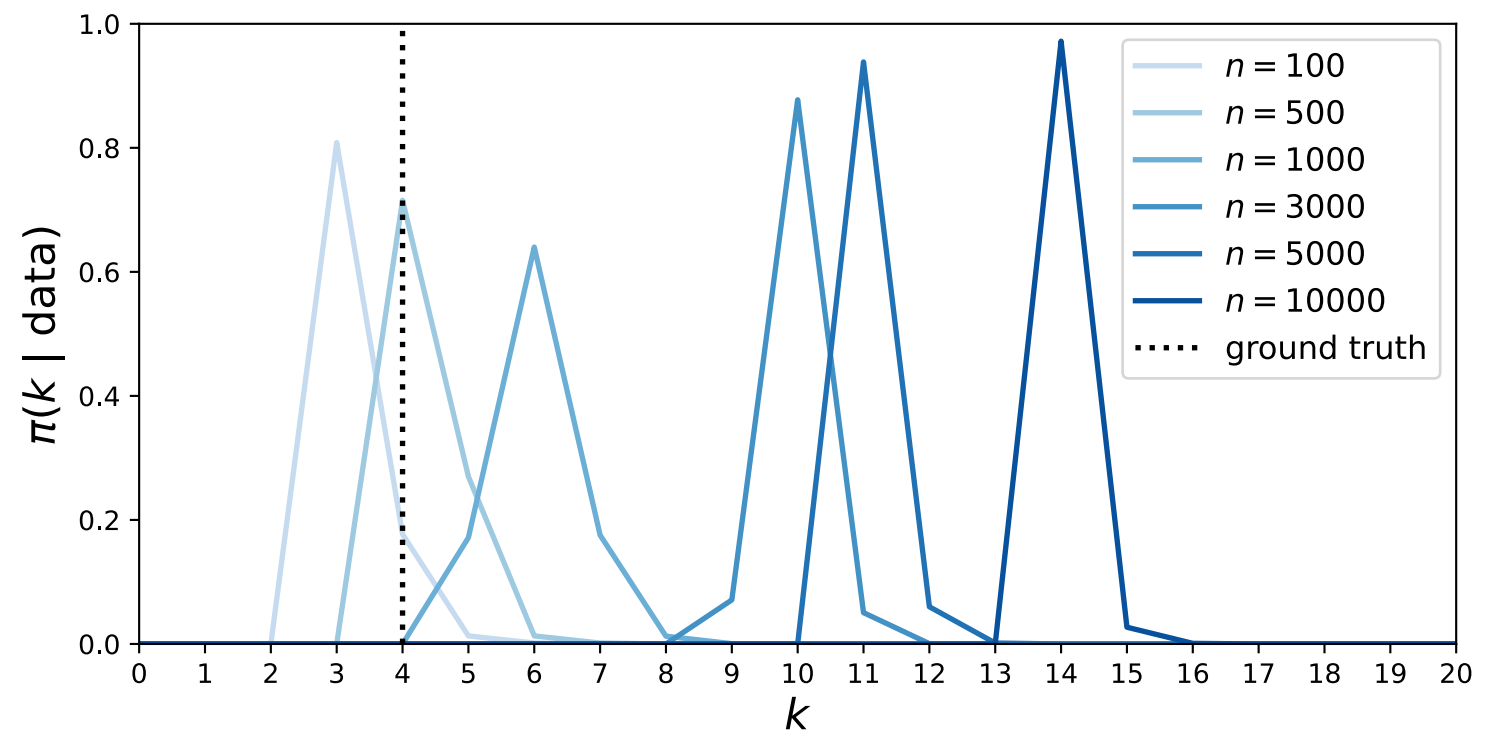


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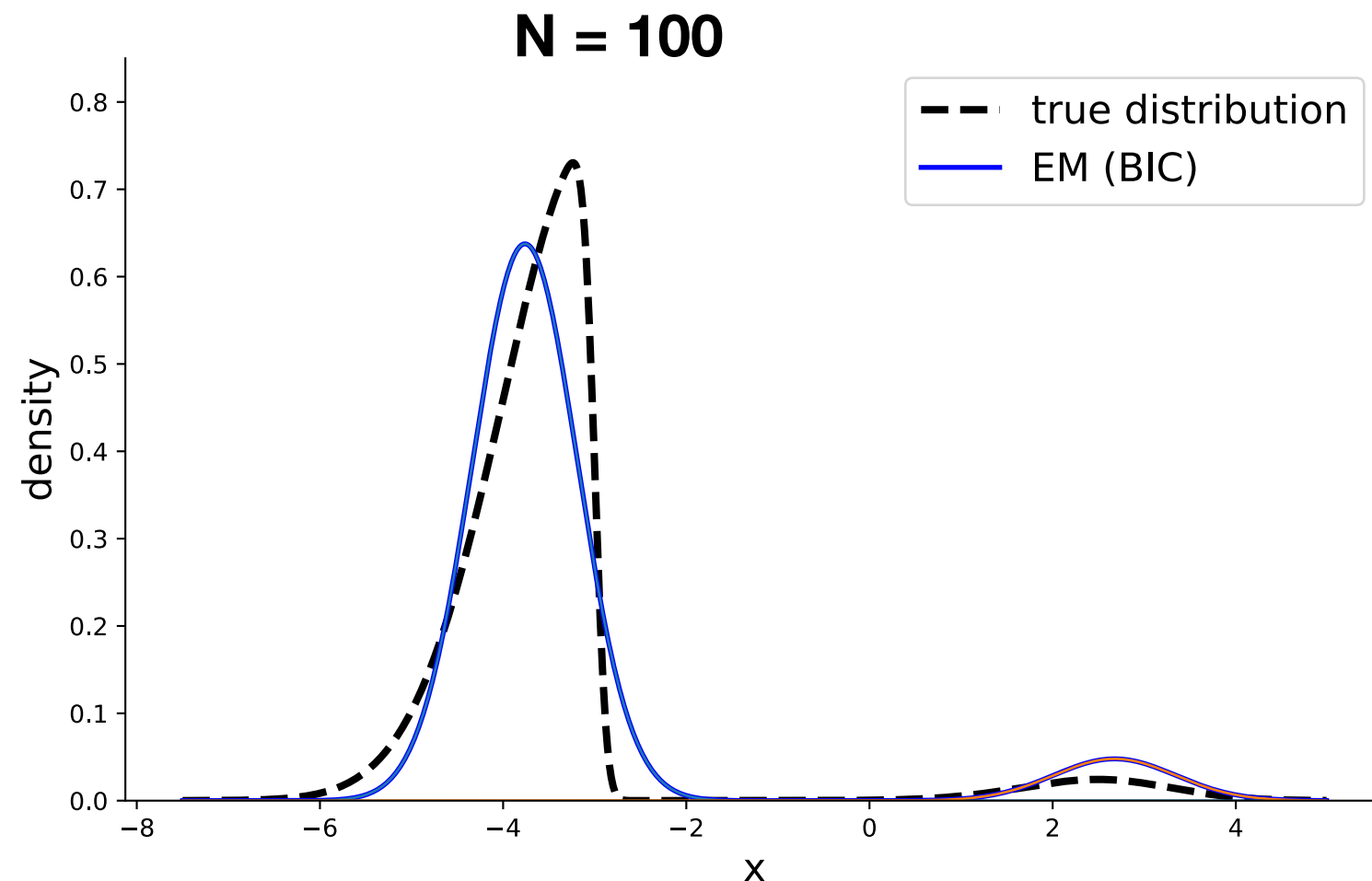


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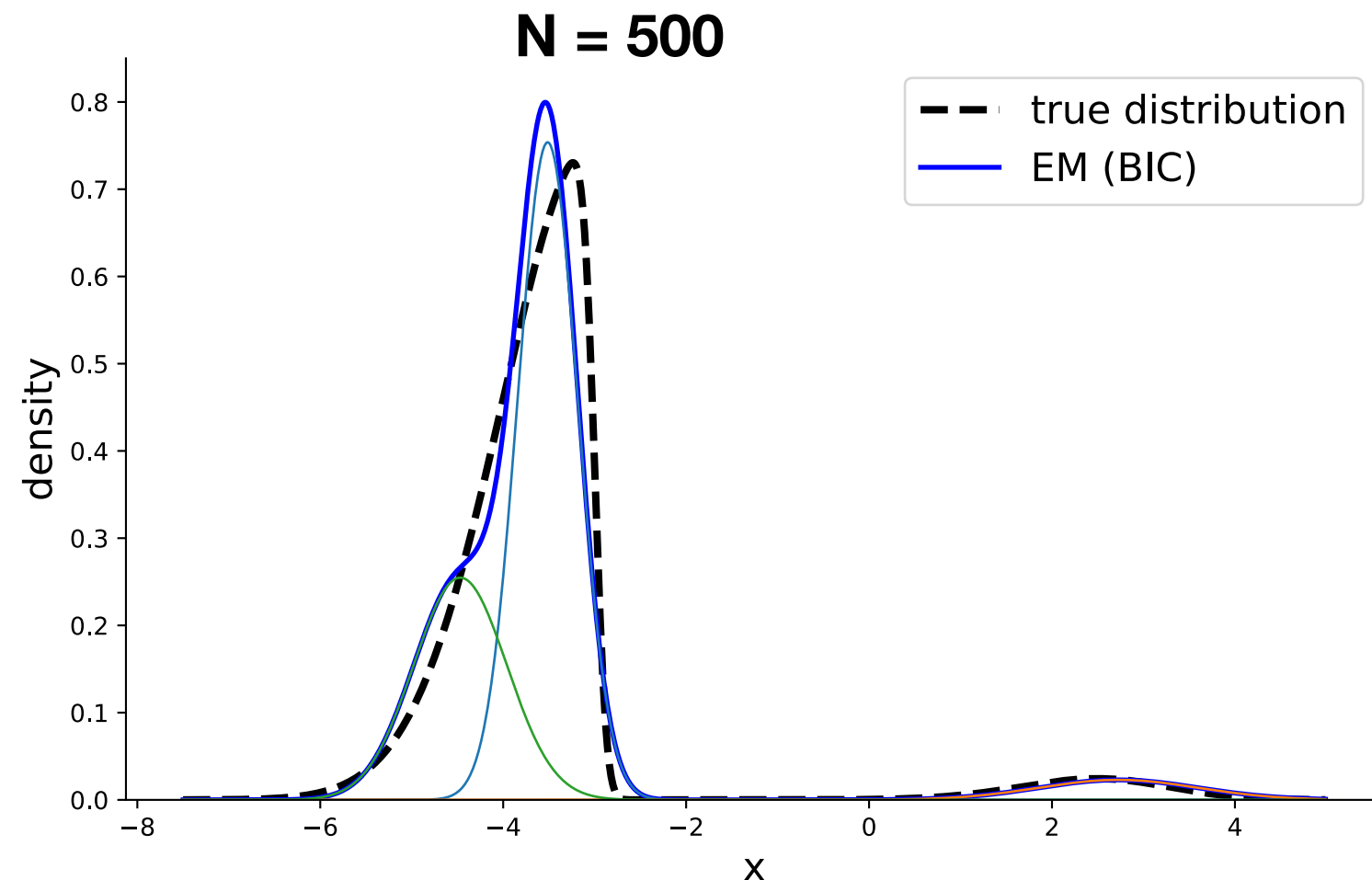


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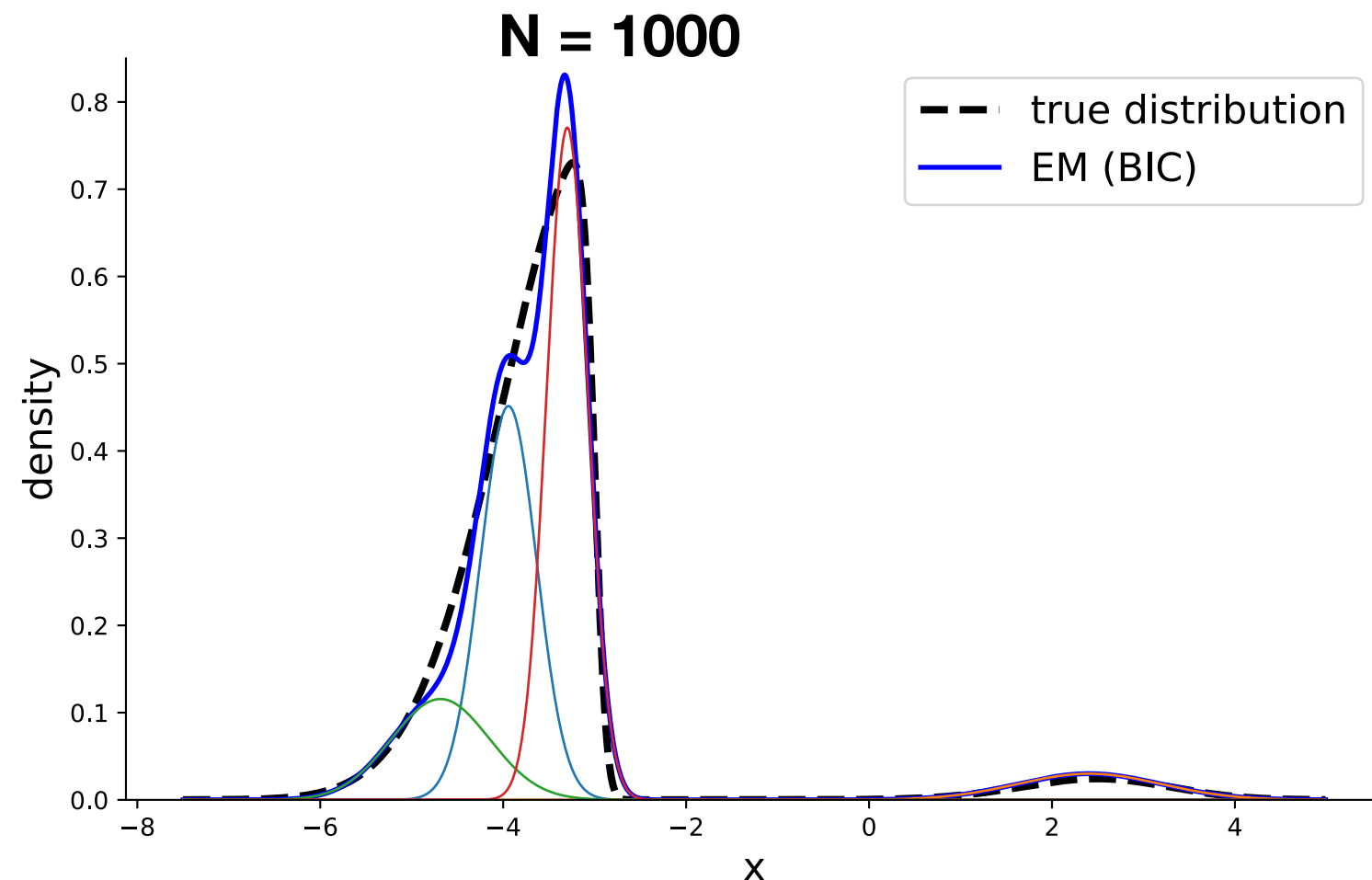


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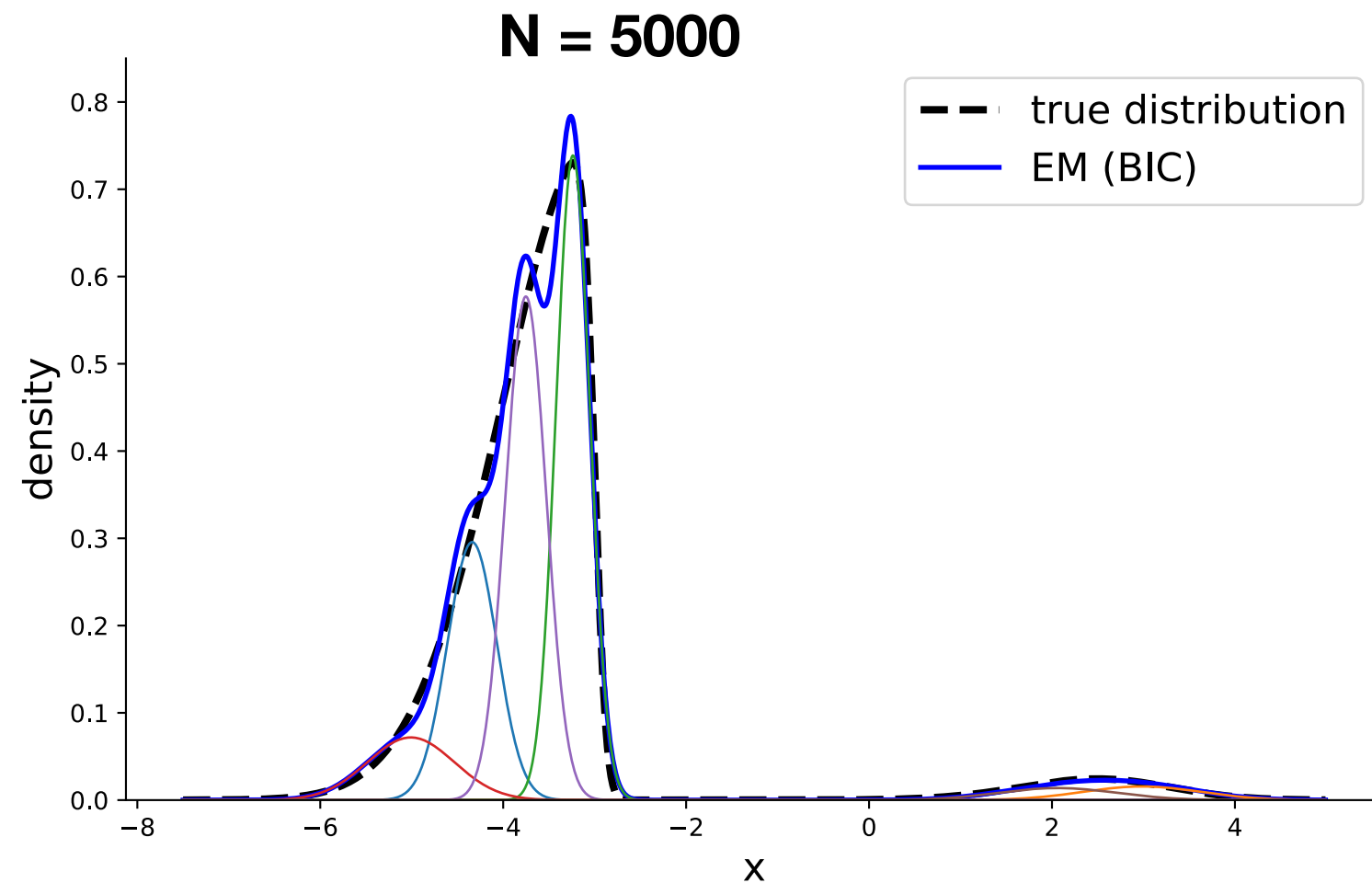


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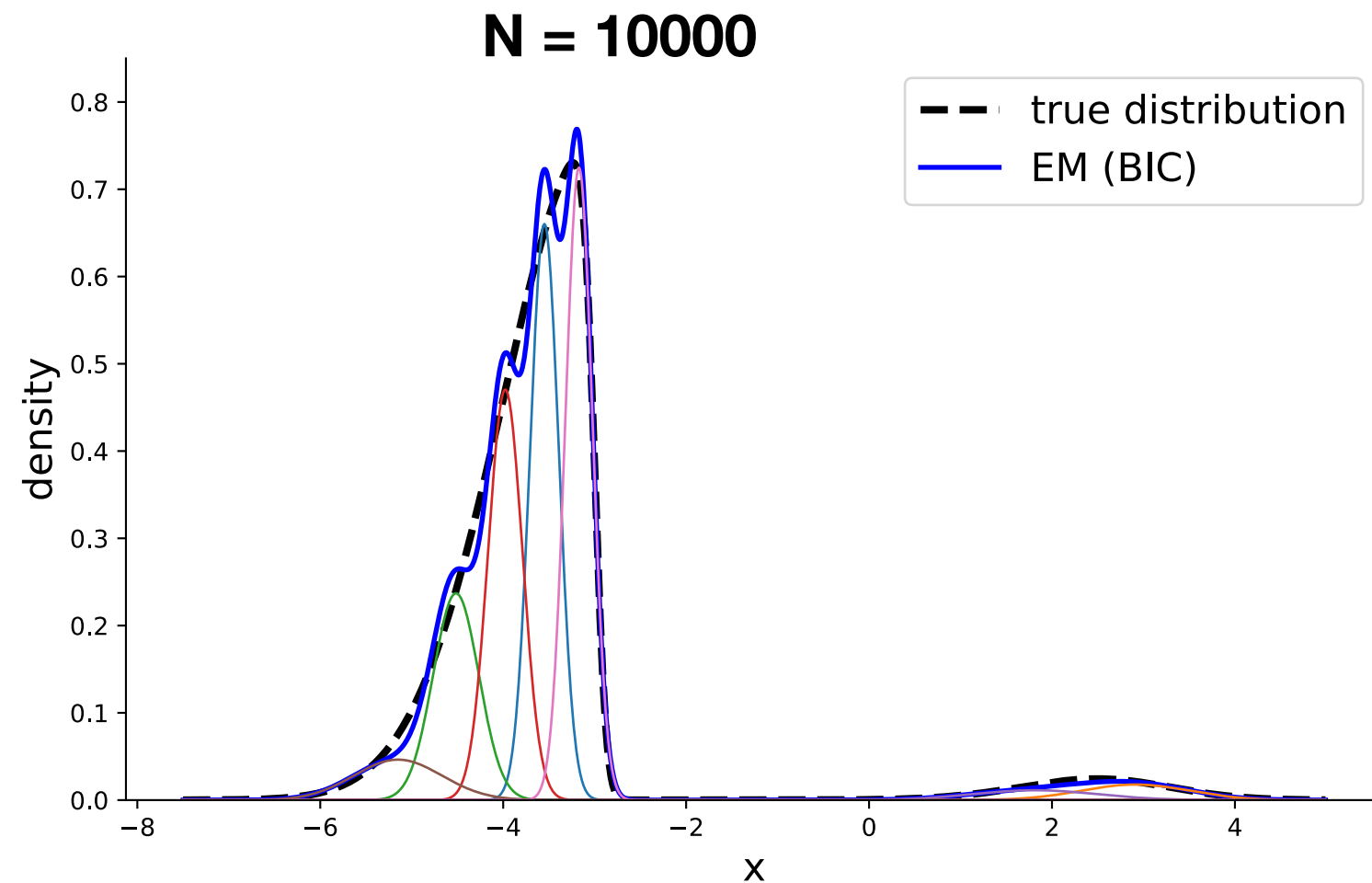


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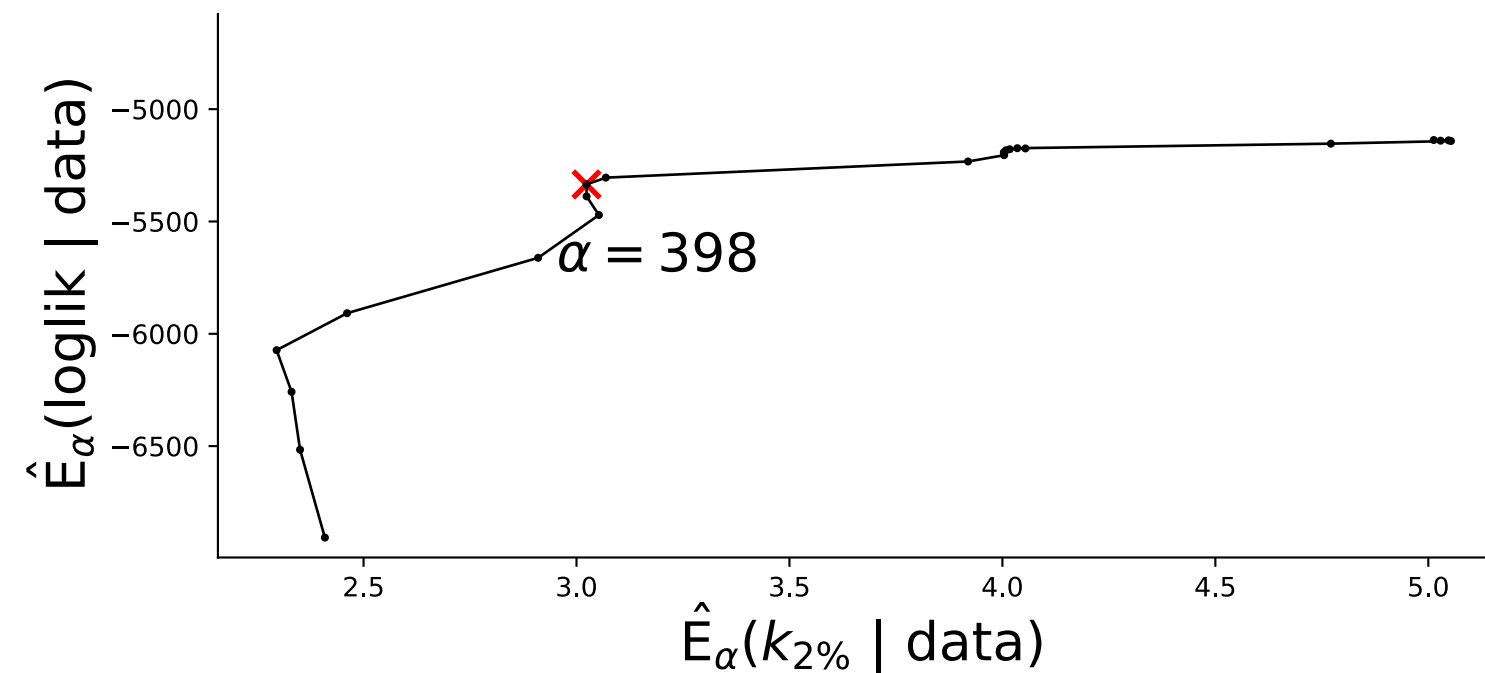
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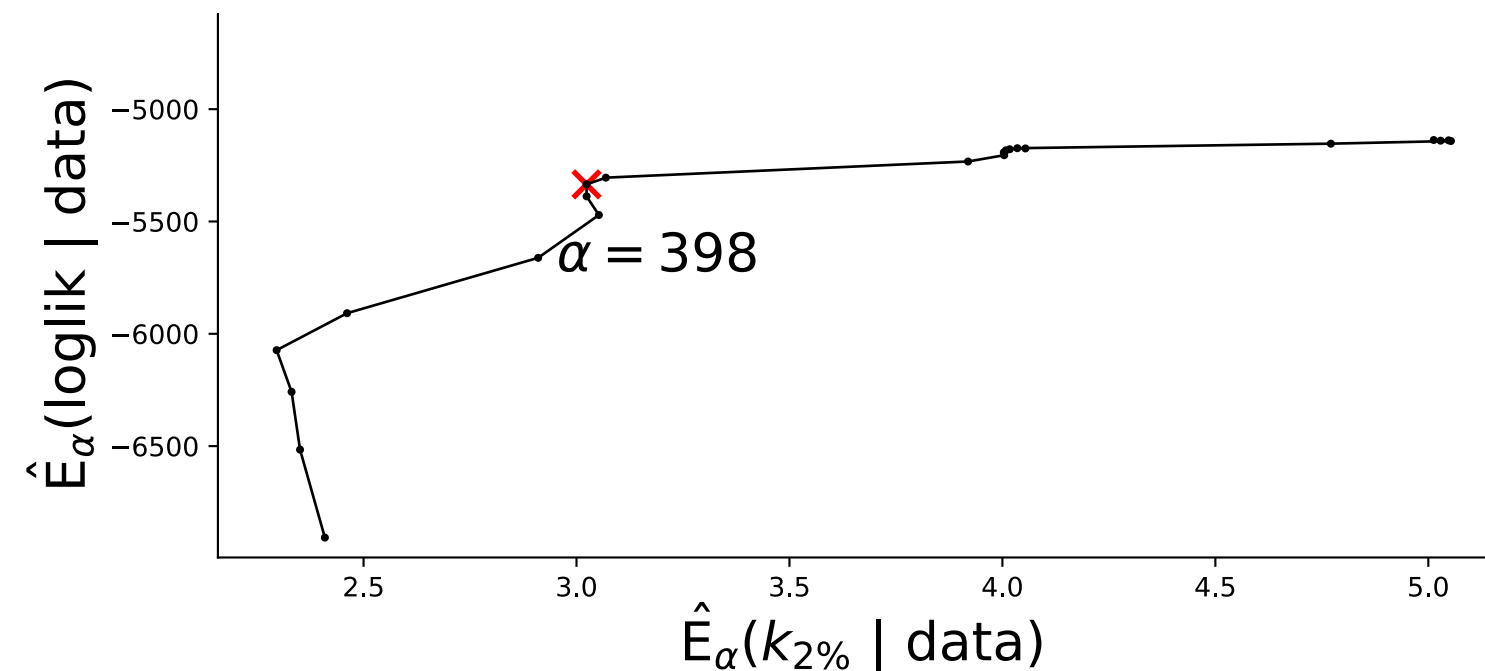
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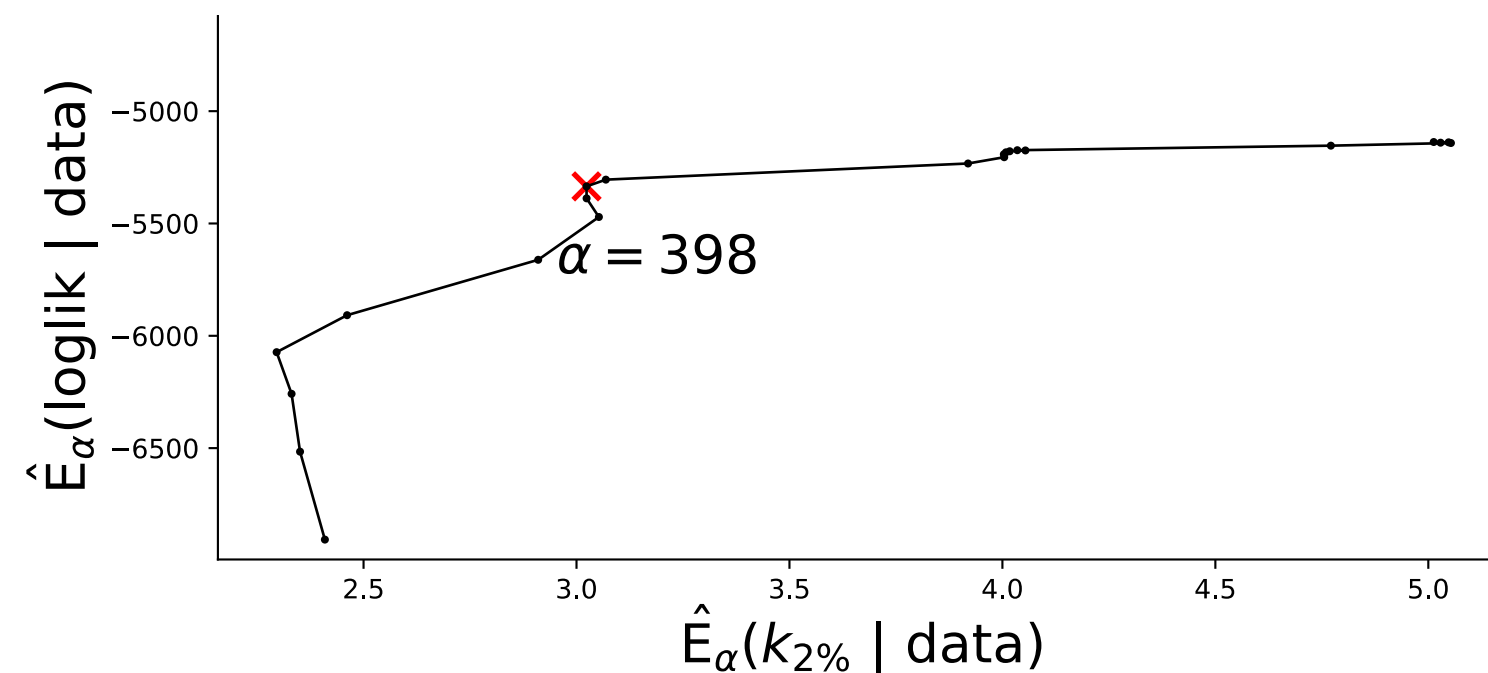
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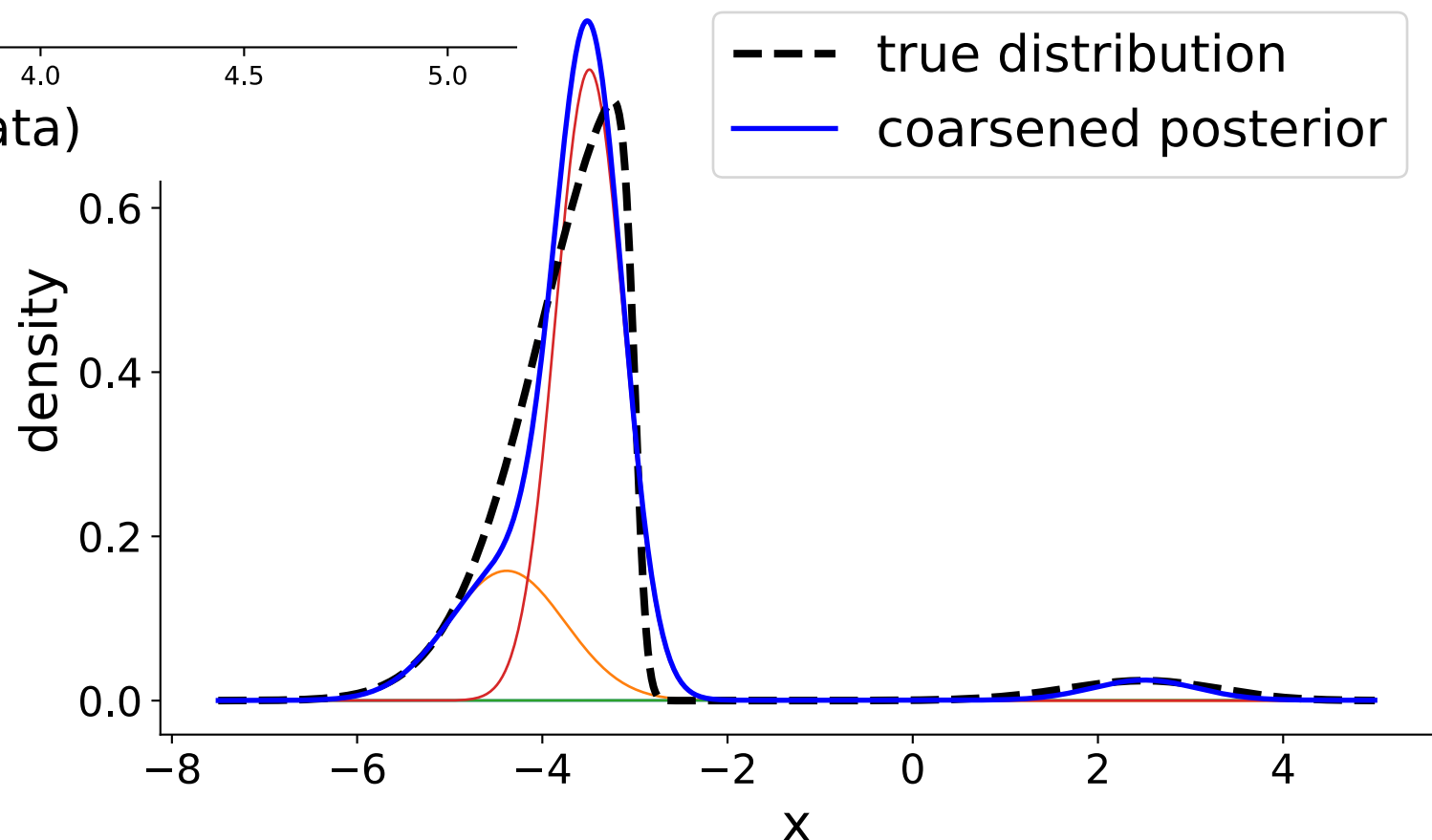
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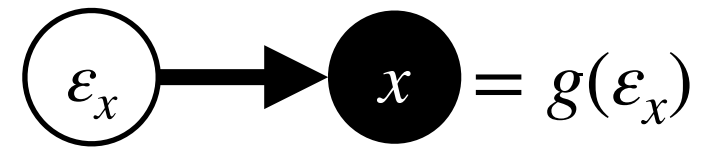
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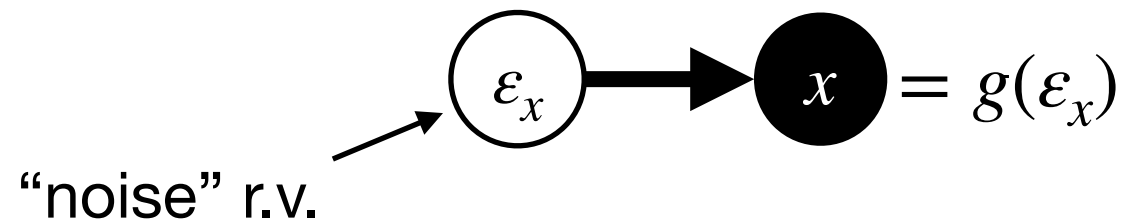


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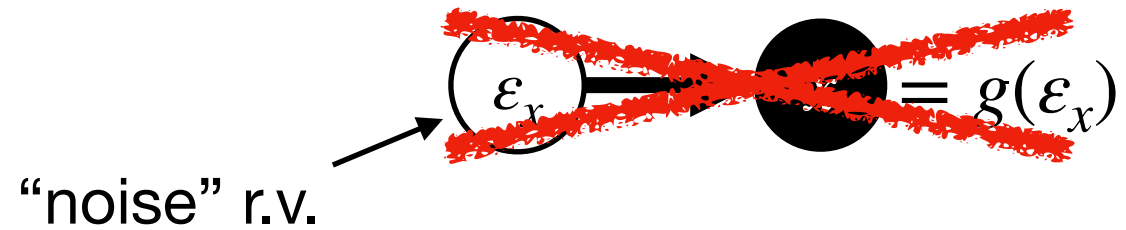


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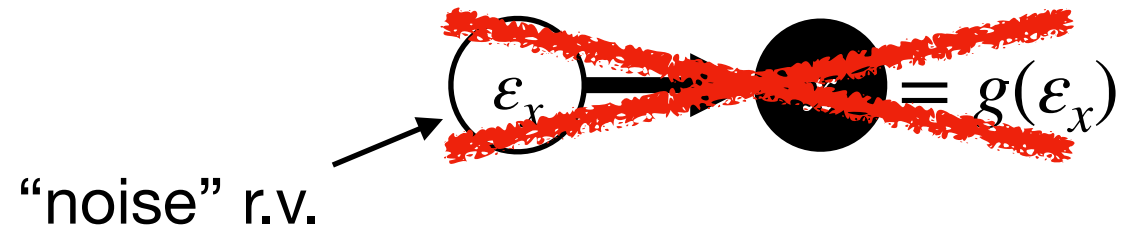
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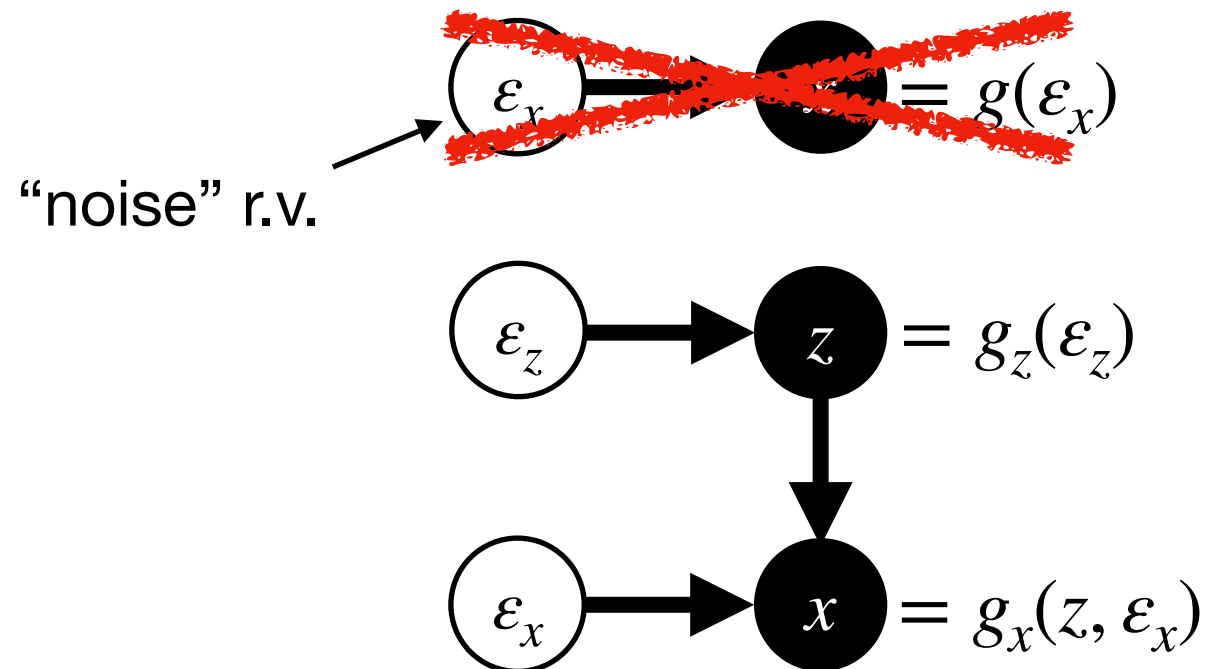
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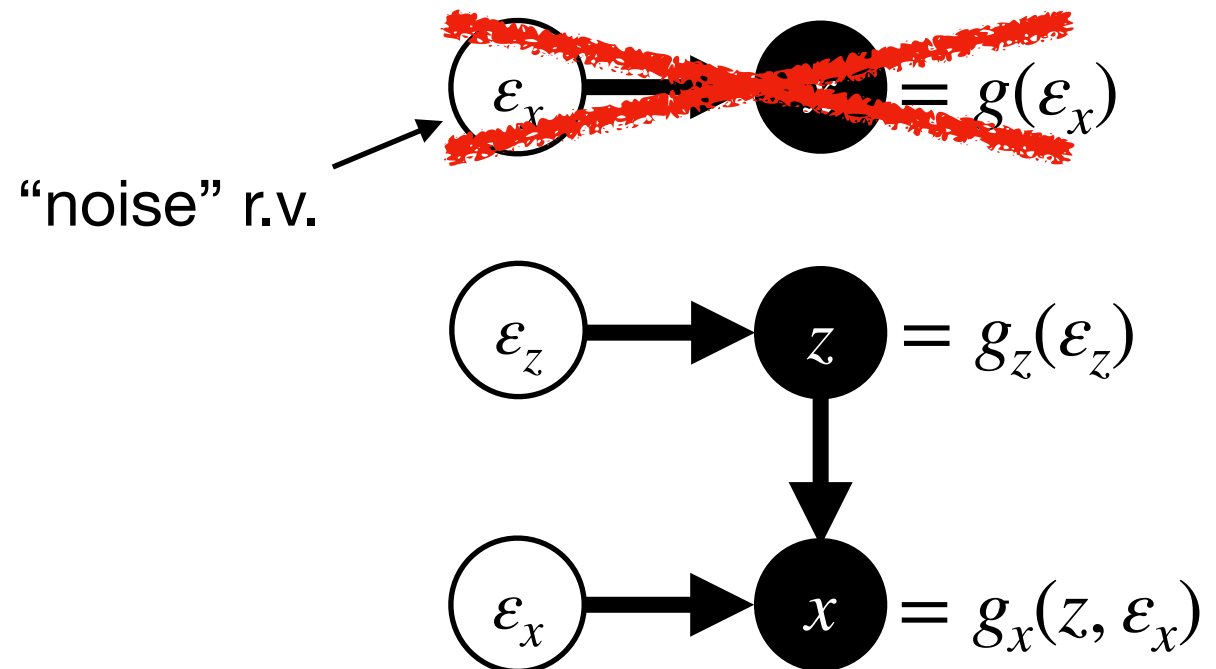
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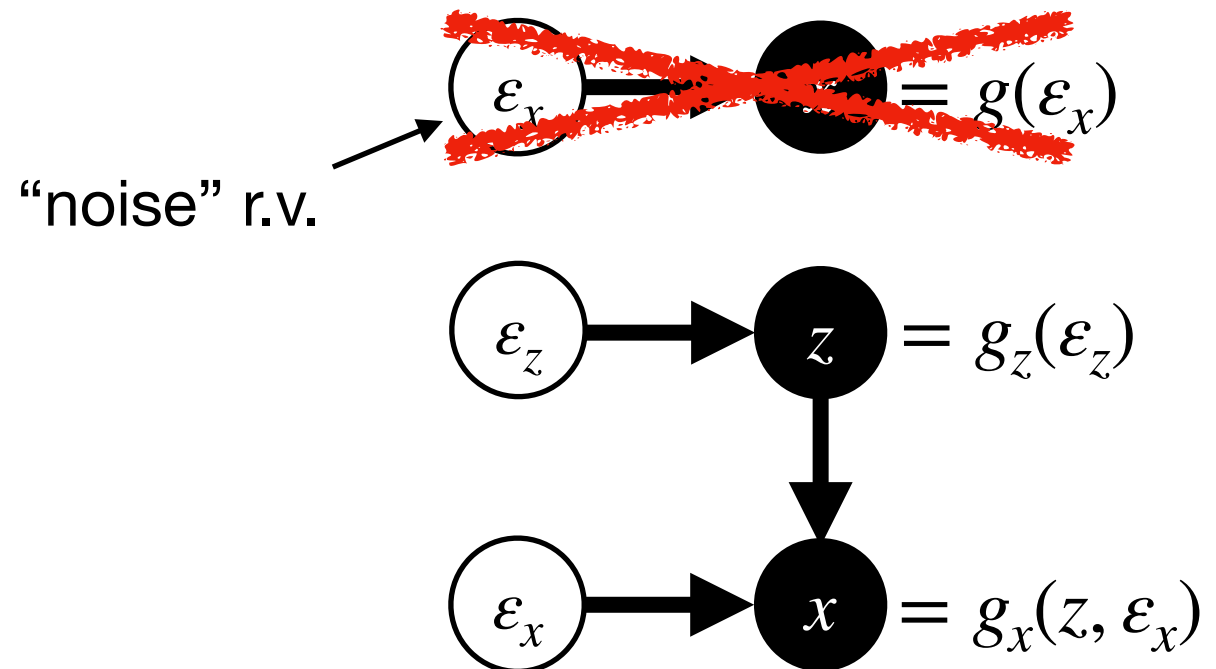
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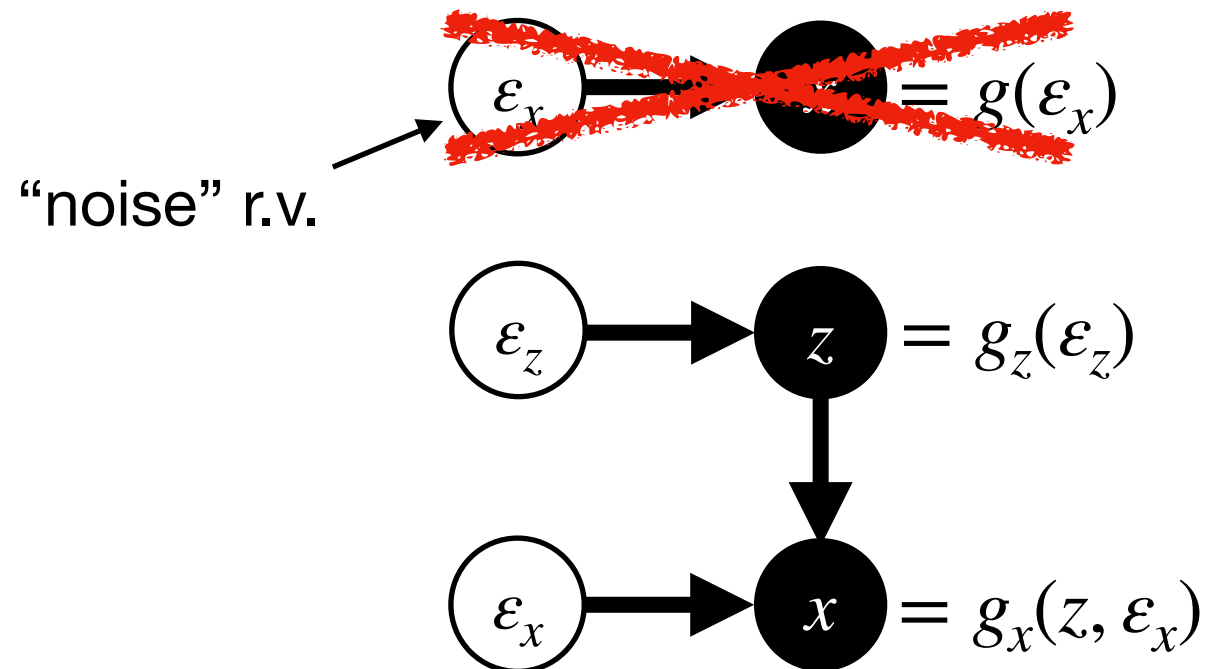
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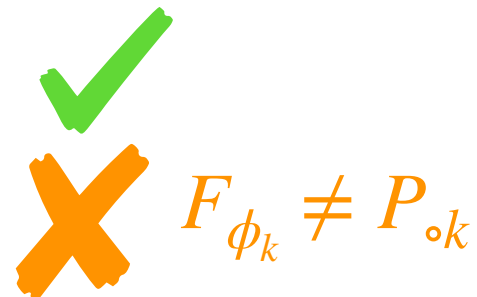
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AIC-like penalty



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► Special case of our general consistency result

Skew-Gaussian mixture: STARE

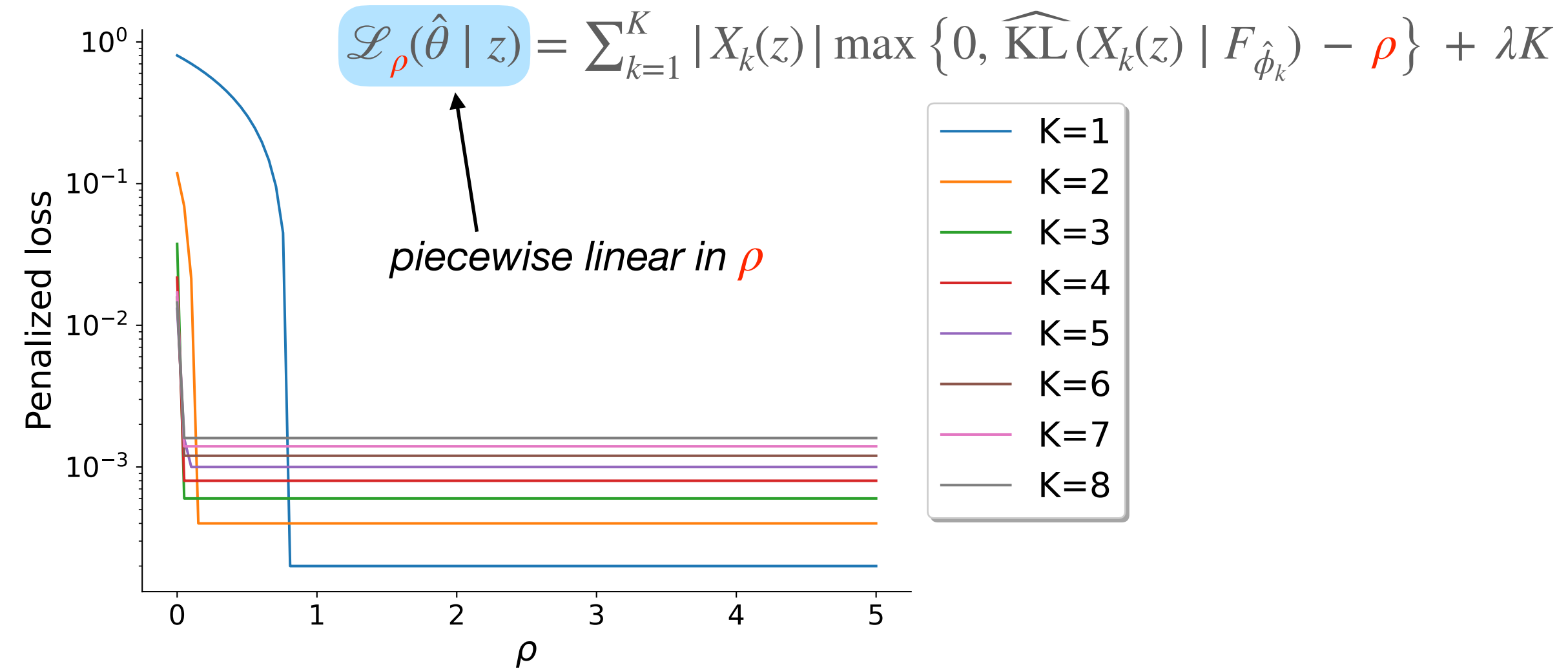
$$\mathcal{L}_{\rho}(\hat{\theta} | z) = \sum_{k=1}^K |X_k(z)| \max \{0, \widehat{\text{KL}}(X_k(z) | F_{\hat{\phi}_k}) - \rho\} + \lambda K$$

Skew-Gaussian mixture: STARE

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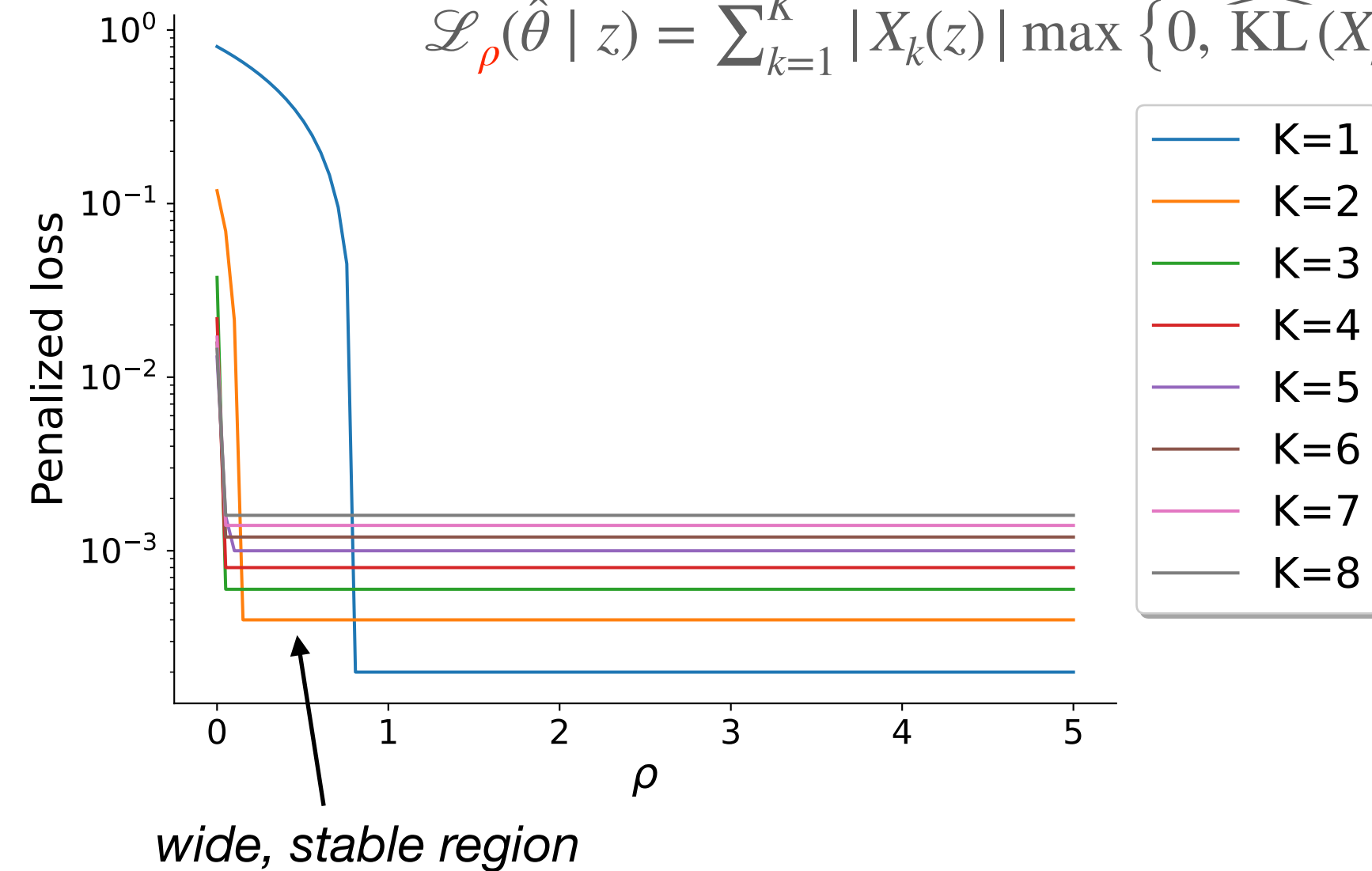
piecewise linear in ρ

Skew-Gaussian mixture: STARE



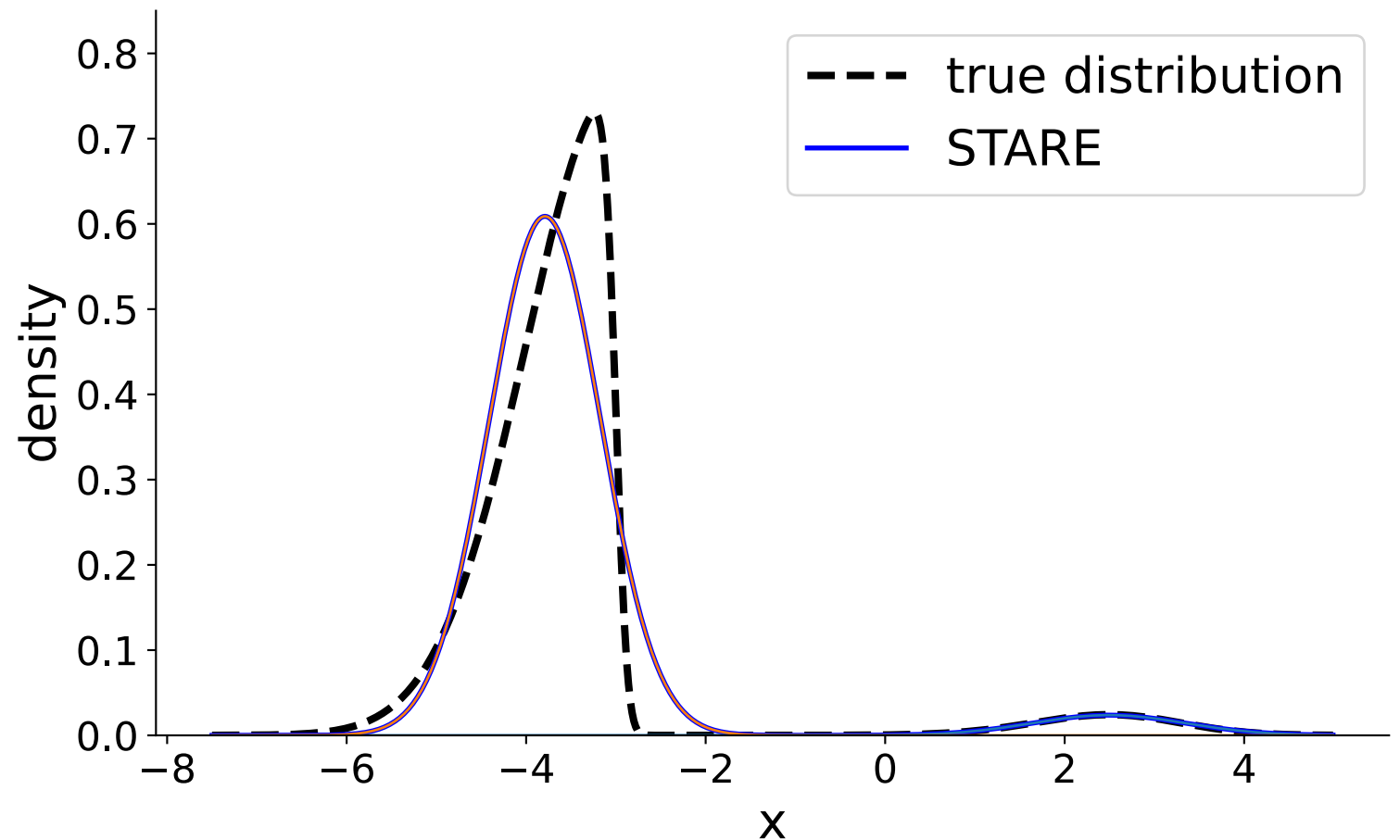
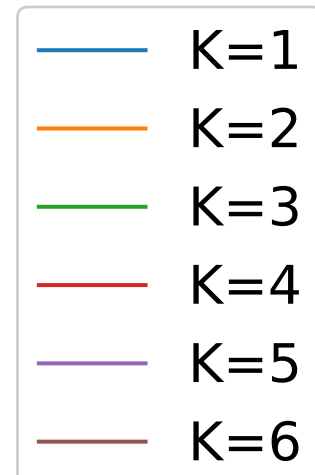
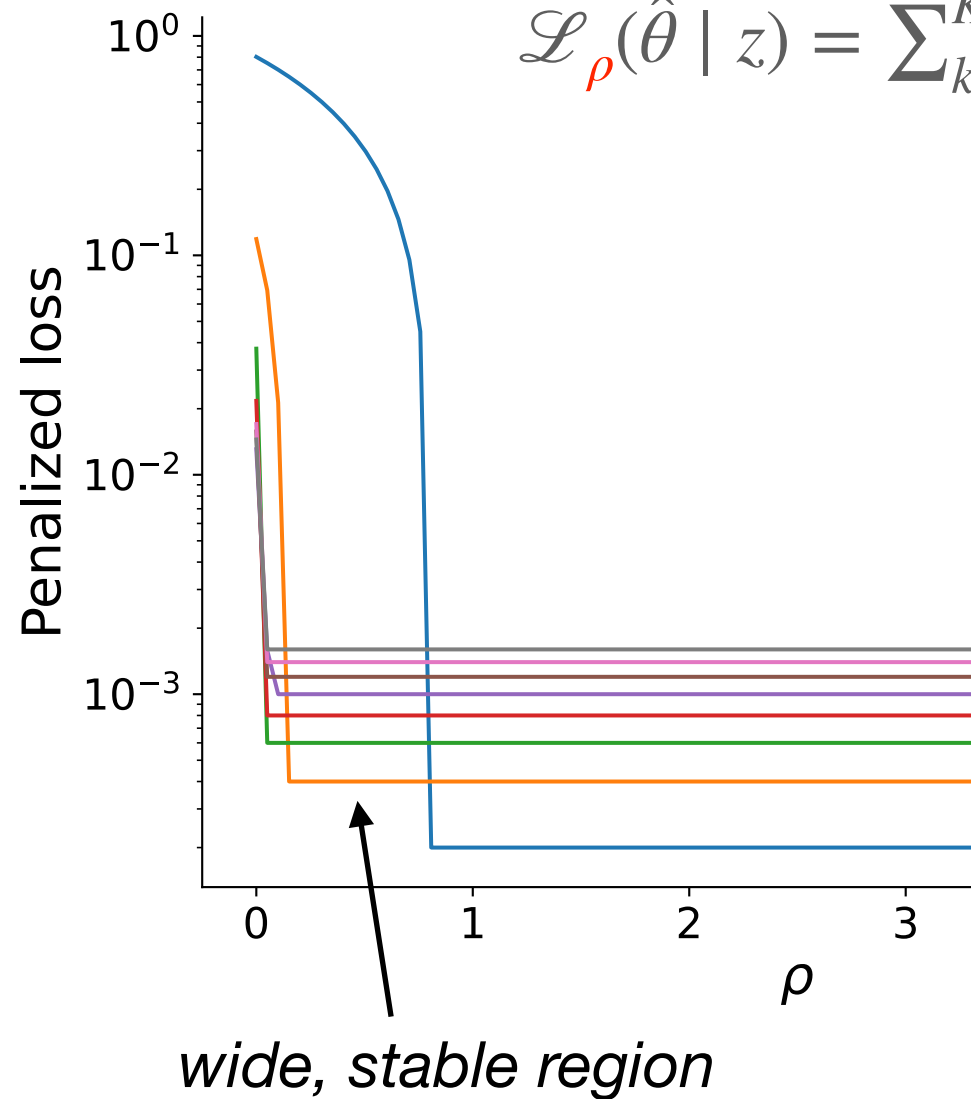
Skew-Gaussian mixture: STARE

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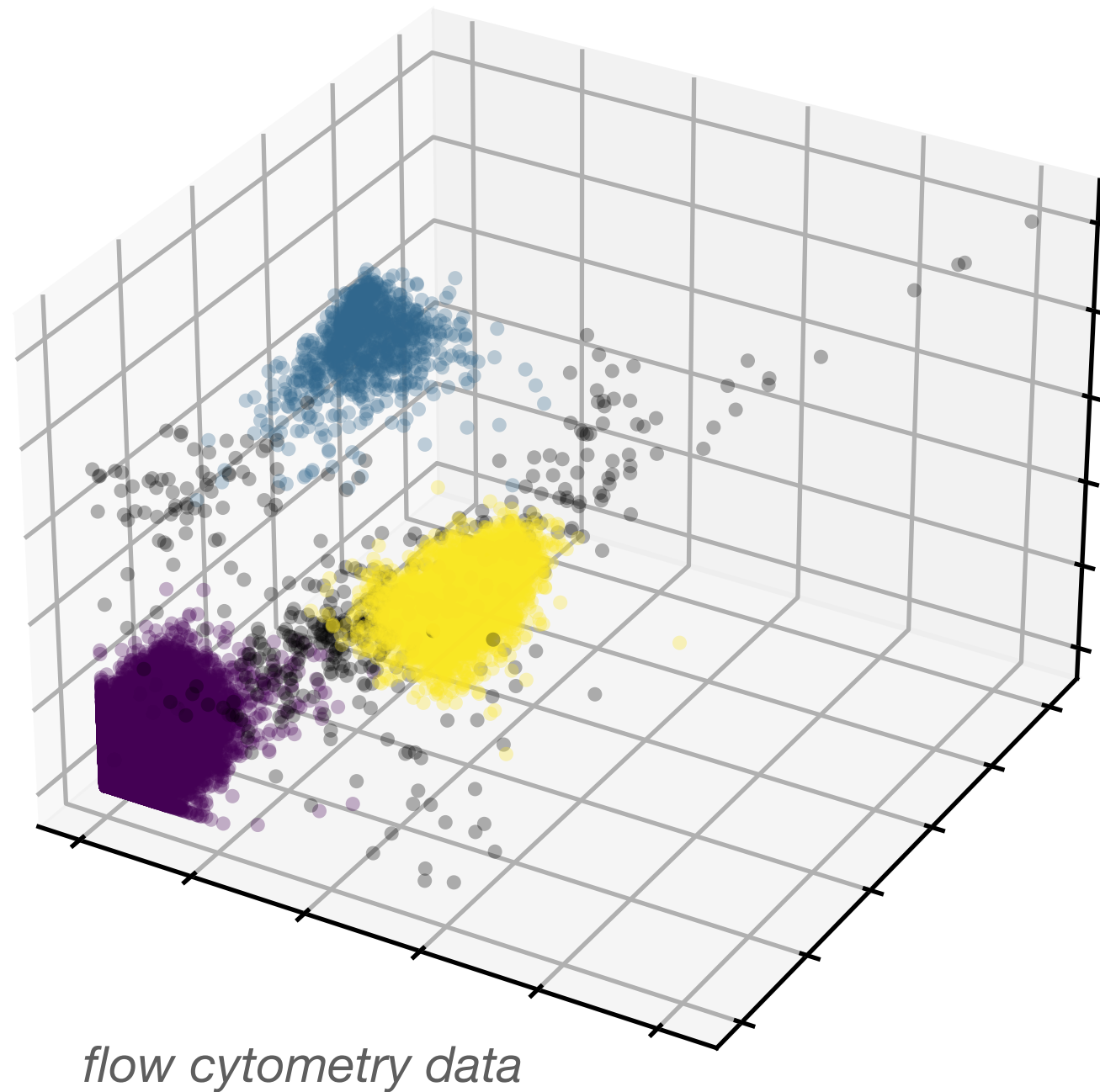


Skew-Gaussian mixture: STARE

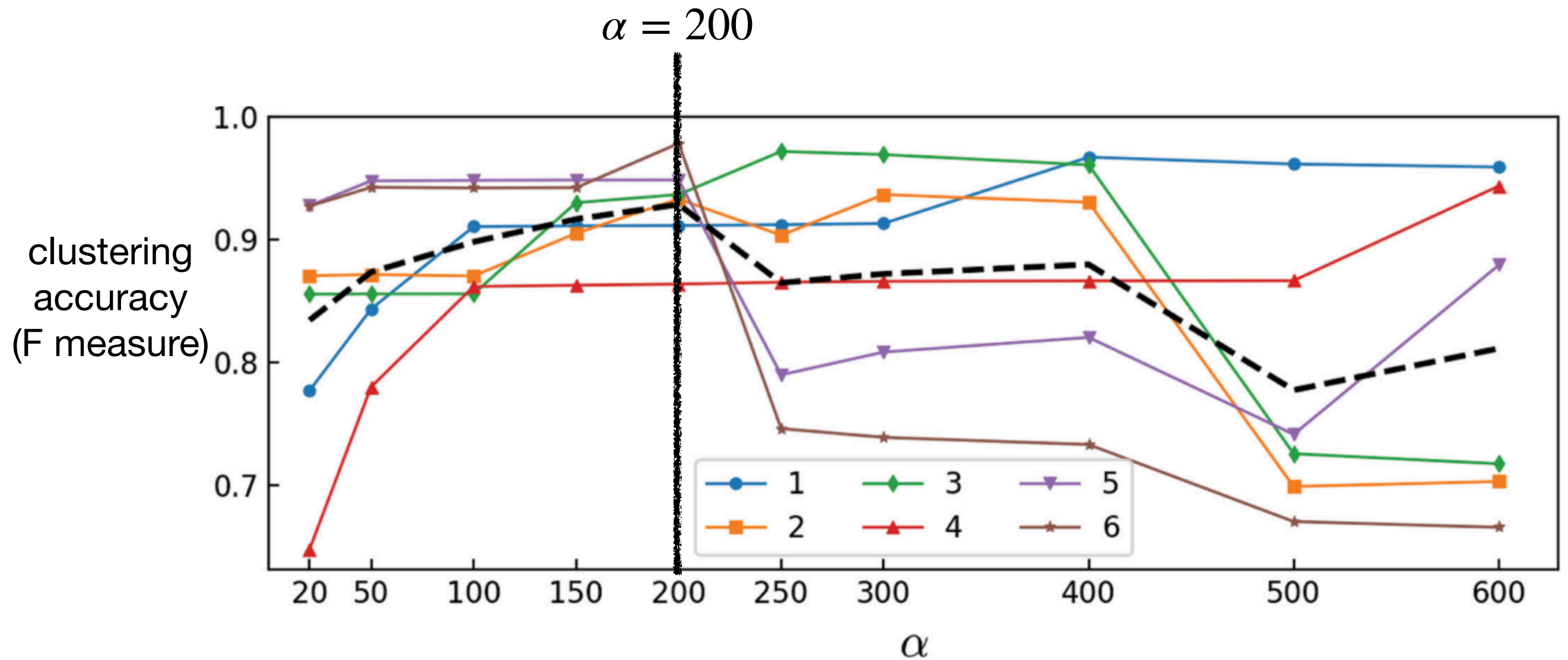
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Clustering cells by type: calibrating α

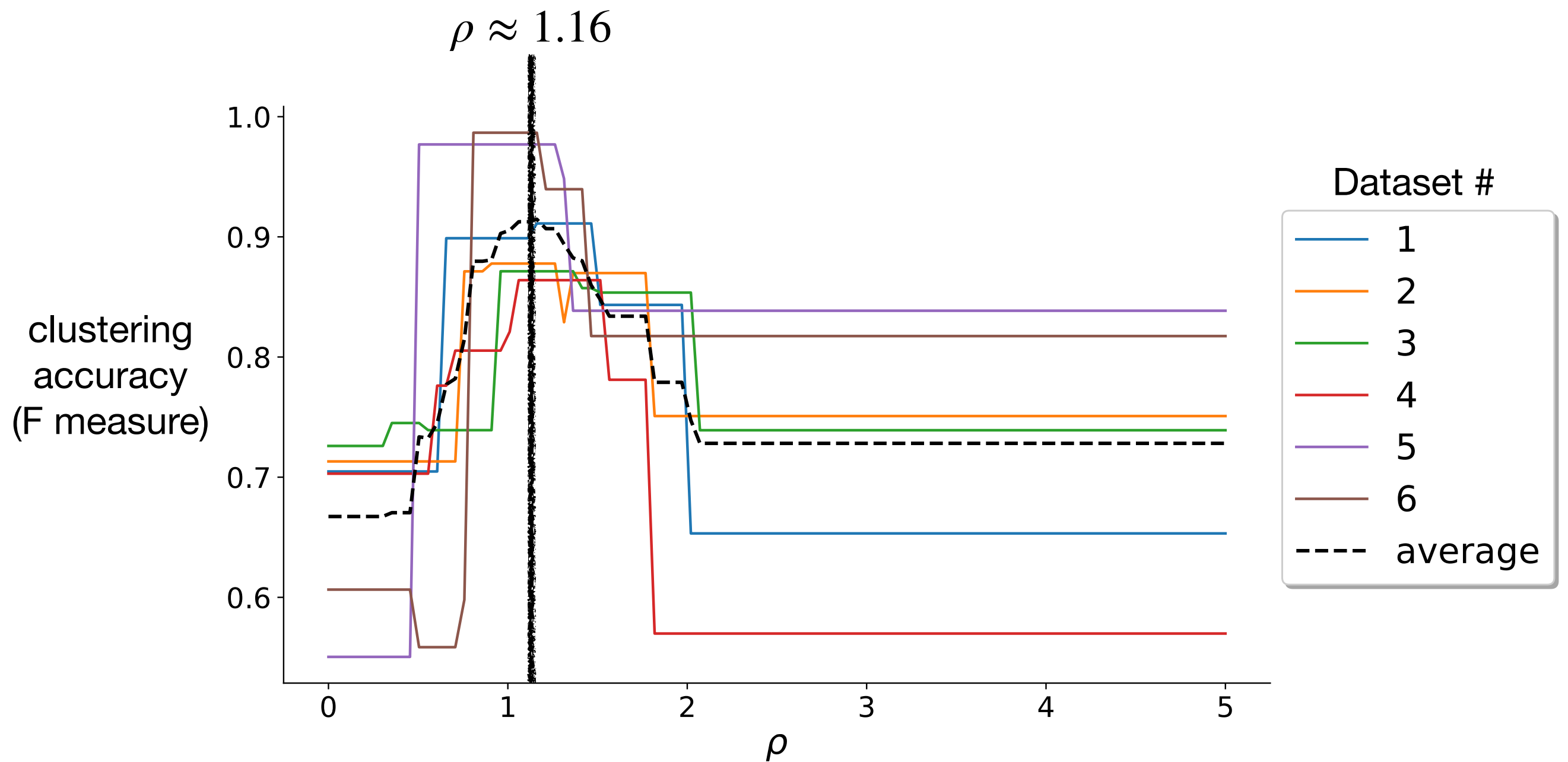


Clustering cells by type: calibrating α

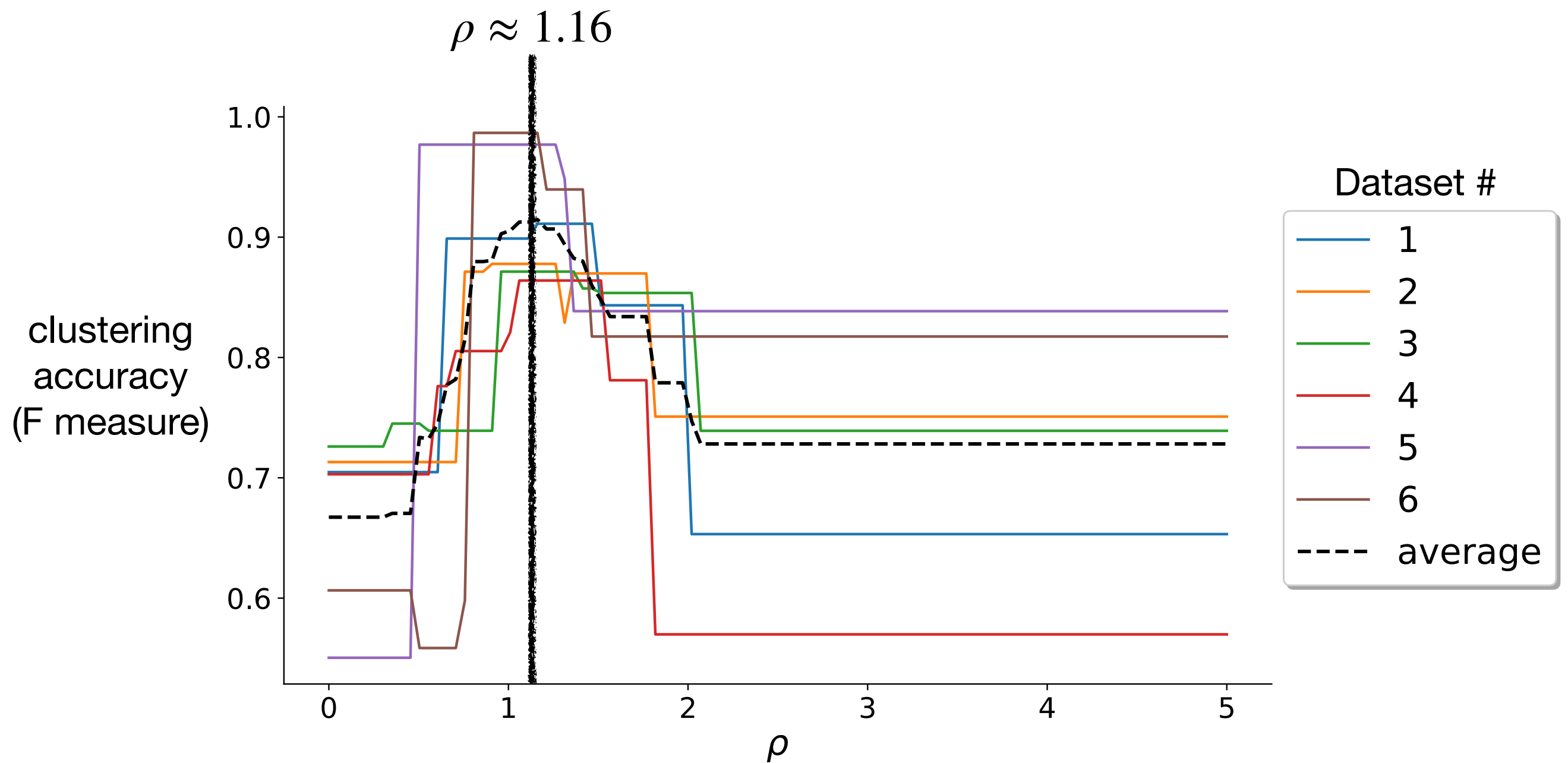


Clustering cells by type: calibrating ρ

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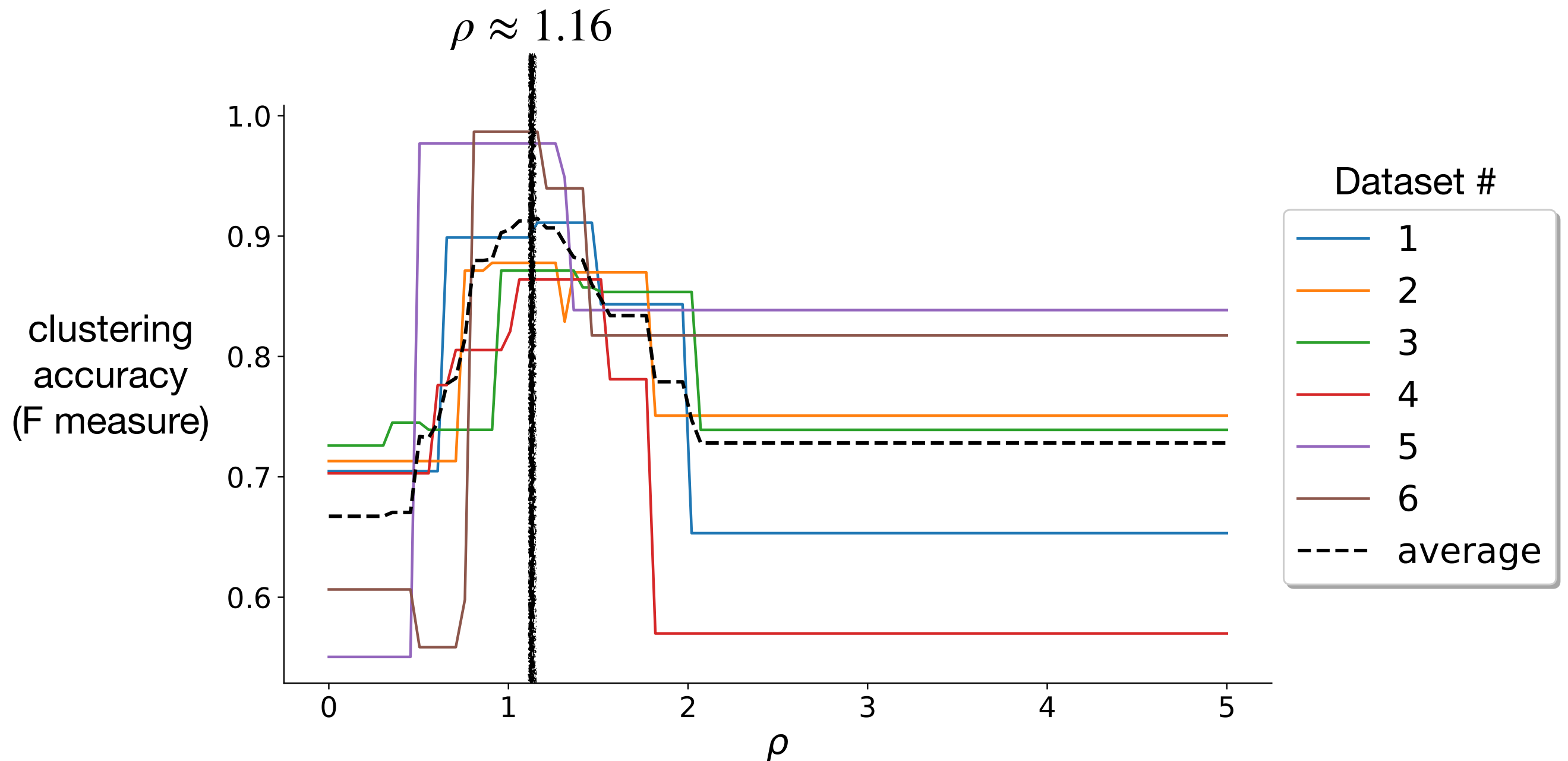


Clustering cells by type: calibrating ρ



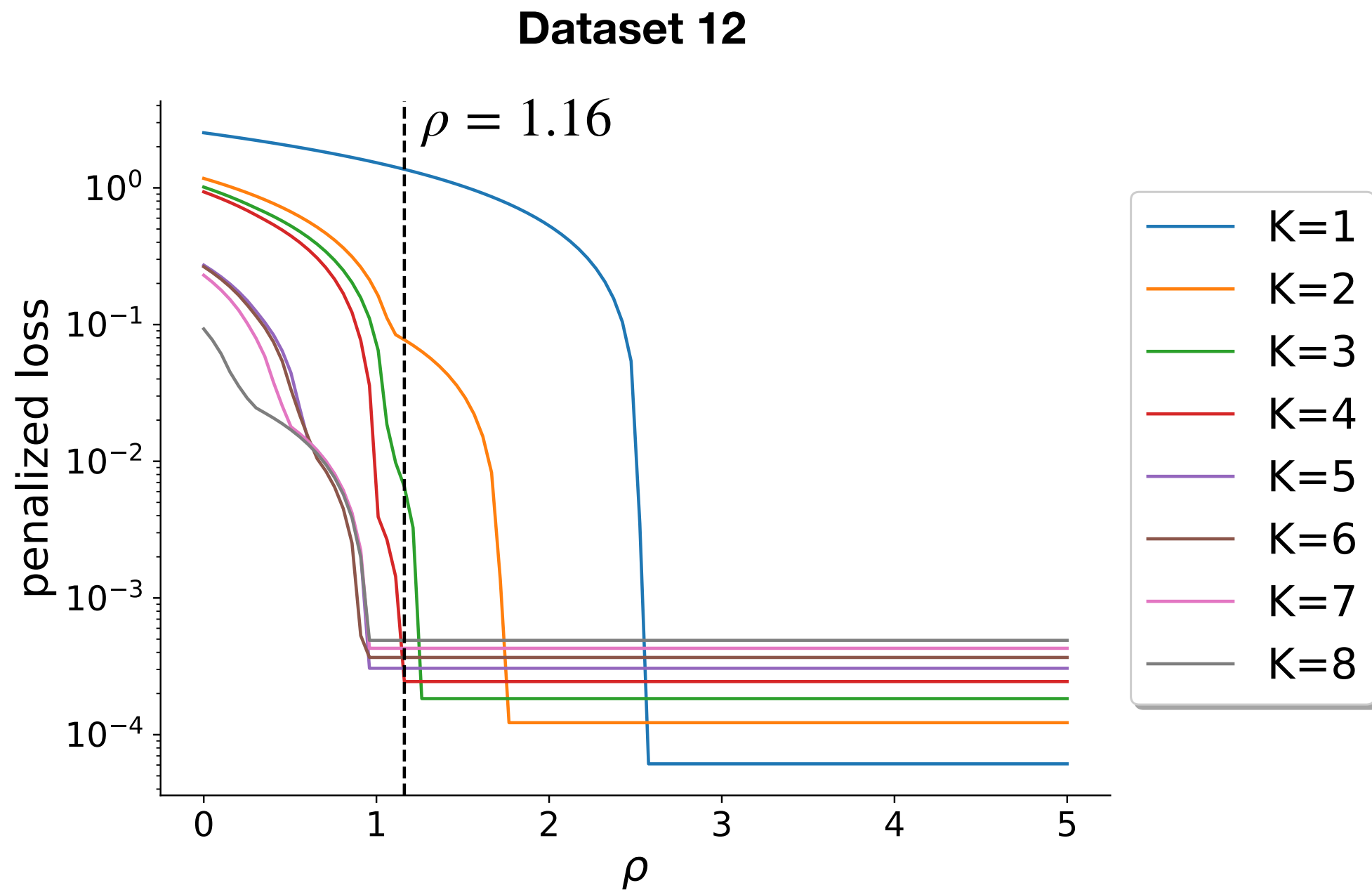
- STARE runtime: 30 min [Python]

Clustering cells by type: calibrating ρ



- STARE runtime: 30 min [Python]
- Coarsened posterior runtime: 2.5 hours [Julia]

A borderline case



Clustering cells by type: STARE is fast and accurate

Clustering accuracy (F-measure)

Dataset	7	8	9	10	11	12
STARE	0.63	0.92	0.94	0.99	0.99	0.98
Coarsened posterior	0.67	0.88	0.93	0.99	0.99	0.99

Closing thoughts

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 - ▶ Ideally, want to default to **stable methods** that don't degrade **statistical efficiency**
 - ▶ Examples: **Bagged posterior** and **STARE**

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References

J. H. Huggins & J. W. Miller (2023). Reproducible Model Selection Using Bagged Posteriors. *Bayesian Analysis* 18(1): 79–104

J. H. Huggins & J. W. Miller (2019). Robust Inference and Model Criticism Using Bagged Posteriors. arXiv:1912.07104 [stat.ME].

J. H. Huggins & J. W. Miller (2023+). Reproducible Parameter Inference Using Bagged Posteriors.

J. Li & J. H. Huggins (2023+). Robust, Structurally-Aware Inference with Mixture Models.