

# Local risk-minimization for Barndorff-Nielsen and Shephard models

Takuji Arai

Keio University

19 August 2016

Boston University/Keio University Workshop 2016

This talk is based on a joint work with Yuto Imai (Waseda) and Ryoichi Suzuki (Keio)

# Outline

- 1 Local risk-minimization
- 2 Barndorff-Nielsen and Shephard models
- 3 Main results
- 4 Numerical experiments

- 1 Local risk-minimization
- 2 Barndorff-Nielsen and Shephard models
- 3 Main results
- 4 Numerical experiments

## Introduction to pricing theory

We consider a financial market with maturity  $T$ , which is composed of **one risky asset** whose price is represented by a semimartingale  $\mathbf{S}$ , and **one riskless asset** with  $\mathbf{0}$  interest rate.

## Introduction to pricing theory

We consider a financial market with maturity  $T$ , which is composed of **one risky asset** whose price is represented by a semimartingale  $\mathbf{S}$ , and **one riskless asset** with  $\mathbf{0}$  interest rate.

Consider a **contingent claim**  $\mathbf{F} \in L^2(\mathbb{P})$ .

For example, if  $\mathbf{F}$  is a call option with strike price  $\mathbf{K}$ , then  $\mathbf{F} = (\mathbf{S}_T - \mathbf{K})^+$ .

## Introduction to pricing theory

We consider a financial market with maturity  $T$ , which is composed of **one risky asset** whose price is represented by a semimartingale  $\mathbf{S}$ , and **one riskless asset** with  $\mathbf{0}$  interest rate.

Consider a **contingent claim**  $F \in L^2(\mathbb{P})$ .

For example, if  $F$  is a call option with strike price  $K$ , then  $F = (\mathbf{S}_T - K)^+$ .

Let  $\xi_t$  (resp.  $\eta_t$ ) be the amount of units of the risky asset (resp. the risk-free asset) an investor holds at time  $t \in [0, T]$ , respectively.

## Introduction to pricing theory

We consider a financial market with maturity  $T$ , which is composed of **one risky asset** whose price is represented by a semimartingale  $\mathbf{S}$ , and **one riskless asset** with  $\mathbf{0}$  interest rate.

Consider a **contingent claim**  $\mathbf{F} \in L^2(\mathbb{P})$ .

For example, if  $\mathbf{F}$  is a call option with strike price  $\mathbf{K}$ , then  $\mathbf{F} = (\mathbf{S}_T - \mathbf{K})^+$ .

Let  $\xi_t$  (resp.  $\eta_t$ ) be the amount of units of the risky asset (resp. the risk-free asset) an investor holds at time  $t \in [0, T]$ , respectively.

$\varphi := (\xi, \eta)$  is called a **strategy**. The **wealth** is given as  $V_t(\varphi) = \eta_t + \xi_t \mathbf{S}_t$ .

## Introduction to pricing theory

We consider a financial market with maturity  $T$ , which is composed of **one risky asset** whose price is represented by a semimartingale  $\mathbf{S}$ , and **one riskless asset** with  $\mathbf{0}$  interest rate.

Consider a **contingent claim**  $F \in L^2(\mathbb{P})$ .

For example, if  $F$  is a call option with strike price  $K$ , then  $F = (\mathbf{S}_T - K)^+$ .

Let  $\xi_t$  (resp.  $\eta_t$ ) be the amount of units of the risky asset (resp. the risk-free asset) an investor holds at time  $t \in [0, T]$ , respectively.

$\varphi := (\xi, \eta)$  is called a **strategy**. The **wealth** is given as  $V_t(\varphi) = \eta_t + \xi_t \mathbf{S}_t$ .

$\varphi$  is **self-financing** if  $V_t(\varphi) = V_0(\varphi) + \int_0^t \xi_s d\mathbf{S}_s$ .



## Introduction to pricing theory

We consider a financial market with maturity  $T$ , which is composed of **one risky asset** whose price is represented by a semimartingale  $\mathbf{S}$ , and **one riskless asset** with  $\mathbf{0}$  interest rate.

Consider a **contingent claim**  $F \in L^2(\mathbb{P})$ .

For example, if  $F$  is a call option with strike price  $K$ , then  $F = (\mathbf{S}_T - K)^+$ .

Let  $\xi_t$  (resp.  $\eta_t$ ) be the amount of units of the risky asset (resp. the risk-free asset) an investor holds at time  $t \in [0, T]$ , respectively.

$\varphi := (\xi, \eta)$  is called a **strategy**. The **wealth** is given as  $V_t(\varphi) = \eta_t + \xi_t \mathbf{S}_t$ .

$\varphi$  is **self-financing** if  $V_t(\varphi) = V_0(\varphi) + \int_0^t \xi_s d\mathbf{S}_s$ .

### Discrete time model

A self-financing strategy satisfies the following at each trading time  $n$ :

$$\eta_n + \xi_n \mathbf{S}_n = \eta_{n+1} + \xi_{n+1} \mathbf{S}_n (= V_n(\varphi))$$

In particular, we have  $V_n(\varphi) = V_0(\varphi) + \sum_{k=1}^n \xi_k \Delta \mathbf{S}_k$ .

## Introduction to pricing theory (cont'd)

A market is called **complete**, if, for any  $F$ , we can find a  $c \in \mathbb{R}$  and a predictable process  $\xi$  satisfying

$$F = c + \int_0^T \xi_t dS_t.$$

The pair  $(c, \xi)$  is called the **perfect hedge** of  $F$ .

Note that  $(c, \xi)$  has a one-to-one corresponding to  $\varphi$  for self-financing strategies.

## Introduction to pricing theory (cont'd)

A market is called **complete**, if, for any  $F$ , we can find a  $c \in \mathbb{R}$  and a predictable process  $\xi$  satisfying

$$F = c + \int_0^T \xi_t dS_t.$$

The pair  $(c, \xi)$  is called the **perfect hedge** of  $F$ .

Note that  $(c, \xi)$  has a one-to-one corresponding to  $\varphi$  for self-financing strategies.

In this talk, we focus on hedging strategies for **incomplete** markets.

Instead of the perfect hedge, we consider a replicating strategy which is not self-financing.

# Risk minimization

Consider a strategy  $\varphi := (\xi, \eta)$ , which is not necessarily self-financing. Define the **cumulative cost process**  $C_t(\varphi)$  as

$$C_t(\varphi) := V_t(\varphi) - \int_0^t \xi_s dS_s.$$

# Risk minimization

Consider a strategy  $\varphi := (\xi, \eta)$ , which is not necessarily self-financing. Define the **cumulative cost process**  $C_t(\varphi)$  as

$$C_t(\varphi) := V_t(\varphi) - \int_0^t \xi_s dS_s.$$

Define the **risk process**  $R_t(\varphi)$  as

$$R_t(\varphi) := E \left[ (C_T(\varphi) - C_t(\varphi))^2 \middle| \mathcal{F}_t \right].$$

# Risk minimization

Consider a strategy  $\varphi := (\xi, \eta)$ , which is not necessarily self-financing. Define the **cumulative cost process**  $C_t(\varphi)$  as

$$C_t(\varphi) := V_t(\varphi) - \int_0^t \xi_s dS_s.$$

Define the **risk process**  $R_t(\varphi)$  as

$$R_t(\varphi) := E \left[ (C_T(\varphi) - C_t(\varphi))^2 \middle| \mathcal{F}_t \right].$$

$\varphi$  is said **risk-minimizing** if  $\varphi$  satisfies  $F = V_T(\varphi)$  and

$$R_t(\varphi) \leq R_t(\tilde{\varphi}) \quad \mathbb{P}\text{-a.s. for every } t \in [0, T]$$

for any strategy  $\tilde{\varphi}$  satisfying  $F = V_T(\tilde{\varphi})$ .

# Local risk-minimization

## Assumption 1

- 1  $\mathbf{S}$  is a special semimartingale with the canonical decomposition  $\mathbf{S} = \mathbf{S}_0 + \mathbf{M} + \mathbf{A}$ .
- 2 We can find a predictable process  $\mathbf{\Lambda}$  such that  $d\mathbf{A} = \mathbf{\Lambda}d\langle \mathbf{M} \rangle$ .
- 3 The mean-variance trade-off process  $\mathbf{K}_t := \int_0^t \mathbf{\Lambda}_s^2 d\langle \mathbf{M} \rangle_s$  is finite, that is,  $\mathbf{K}_T$  is finite  $\mathbb{P}$ -a.s.

## Local risk-minimization (cont'd)

### Definition

- ①  $\Theta_S$  denotes the space of all  $\mathbb{R}$ -valued predictable processes  $\xi$  satisfying
- $$\mathbb{E}\left[\int_0^T \xi_t^2 d\langle M \rangle_t + \left(\int_0^T |\xi_t dA_t|\right)^2\right] < \infty.$$



# Local risk-minimization (cont'd)

## Definition

- ①  $\Theta_S$  denotes the space of all  $\mathbb{R}$ -valued predictable processes  $\xi$  satisfying 
$$\mathbb{E}\left[\int_0^T \xi_t^2 d\langle M \rangle_t + \left(\int_0^T |\xi_t dA_t|\right)^2\right] < \infty.$$
- ② An  **$L^2$ -strategy** is given by a pair  $\varphi = (\xi, \eta)$ , where  $\xi \in \Theta_S$  and  $\eta$  is an adapted process such that  $V(\varphi) := \xi S + \eta$  is a right continuous process with  $\mathbb{E}[V_t^2(\varphi)] < \infty$  for every  $t \in [0, T]$ .  
Note that  $\xi_t$  (resp.  $\eta_t$ ) represents the amount of units of the risky asset (resp. the riskfree asset) an investor holds at time  $t$ .

# Local risk-minimization (cont'd)

## Definition

- 1  $\Theta_S$  denotes the space of all  $\mathbb{R}$ -valued predictable processes  $\xi$  satisfying  $\mathbb{E}[\int_0^T \xi_t^2 d\langle M \rangle_t + (\int_0^T |\xi_t dA_t|)^2] < \infty$ .
- 2 An  **$L^2$ -strategy** is given by a pair  $\varphi = (\xi, \eta)$ , where  $\xi \in \Theta_S$  and  $\eta$  is an adapted process such that  $V(\varphi) := \xi S + \eta$  is a right continuous process with  $\mathbb{E}[V_t^2(\varphi)] < \infty$  for every  $t \in [0, T]$ .  
Note that  $\xi_t$  (resp.  $\eta_t$ ) represents the amount of units of the risky asset (resp. the riskfree asset) an investor holds at time  $t$ .
- 3 For  $F \in L^2(\mathbb{P})$ , the process  $C^F(\varphi)$  defined by

$$C_t^F(\varphi) := F1_{\{t=T\}} + V_t(\varphi) - \int_0^t \xi_s dS_s$$

is called the **cost process** of  $\varphi = (\xi, \eta)$  for  $F$ .

# Local risk-minimization (cont'd)

## Definition

- ①  $\Theta_S$  denotes the space of all  $\mathbb{R}$ -valued predictable processes  $\xi$  satisfying  $\mathbb{E}[\int_0^T \xi_t^2 d\langle M \rangle_t + (\int_0^T |\xi_t dA_t|)^2] < \infty$ .
- ② An  **$L^2$ -strategy** is given by a pair  $\varphi = (\xi, \eta)$ , where  $\xi \in \Theta_S$  and  $\eta$  is an adapted process such that  $V(\varphi) := \xi S + \eta$  is a right continuous process with  $\mathbb{E}[V_t^2(\varphi)] < \infty$  for every  $t \in [0, T]$ .  
Note that  $\xi_t$  (resp.  $\eta_t$ ) represents the amount of units of the risky asset (resp. the riskfree asset) an investor holds at time  $t$ .
- ③ For  $F \in L^2(\mathbb{P})$ , the process  $C^F(\varphi)$  defined by

$$C_t^F(\varphi) := F1_{\{t=T\}} + V_t(\varphi) - \int_0^t \xi_s dS_s$$

is called the **cost process** of  $\varphi = (\xi, \eta)$  for  $F$ .

- ④ An  $L^2$ -strategy  $\varphi$  is said **locally risk-minimizing** strategy for  $F$  if  $V_T(\varphi) = 0$  and  $C^F(\varphi)$  is a martingale orthogonal to  $M$ , that is,  $C^F(\varphi)M$  is a martingale.

# Föllmer-Schweizer decomposition

An  $F \in L^2(\mathbb{P})$  admits an **FS decomposition** if it can be described by

$$F = F_0 + \int_0^T \xi_t^F dS_t + L_T^F, \quad (1)$$

where  $F_0 \in \mathbb{R}$ ,  $\xi^F \in \Theta_S$  and  $L^F$  is a square-integrable martingale orthogonal to  $\mathbf{M}$  with  $L_0^F = 0$ .

---

<sup>1</sup>Schweizer, M.: Local Risk-Minimization for Multidimensional Assets and Payment Streams. Banach Center Publ. 83, 213–229 (2008)

# Föllmer-Schweizer decomposition

An  $F \in L^2(\mathbb{P})$  admits an **FS decomposition** if it can be described by

$$F = F_0 + \int_0^T \xi_t^F dS_t + L_T^F, \quad (1)$$

where  $F_0 \in \mathbb{R}$ ,  $\xi^F \in \Theta_S$  and  $L^F$  is a square-integrable martingale orthogonal to  $\mathbf{M}$  with  $L_0^F = 0$ .

## Proposition 5.2 of Schweizer<sup>1</sup>

Under Assumption 1, LRM  $\varphi = (\xi, \eta)$  for  $F$  exists if and only if  $F$  admits an FS decomposition, and its relationship is given by

$$\xi_t = \xi_t^F, \quad \eta_t = F_0 + \int_0^t \xi_s^F dS_s + L_t^F - F \mathbf{1}_{\{t=T\}} - \xi_t^F S_t.$$

<sup>1</sup>Schweizer, M.: Local Risk-Minimization for Multidimensional Assets and Payment Streams. Banach Center Publ. 83, 213–229 (2008)

# Minimal martingale measure (MMM)

As a result, it suffices to obtain a representation of  $\xi^F$  in (1) in order to obtain LRM.

## Minimal martingale measure (MMM)

As a result, it suffices to obtain a representation of  $\xi^F$  in (1) in order to obtain LRM.

A martingale measure  $\mathbb{P}^* \sim \mathbb{P}$  is called **minimal** if any square-integrable  $\mathbb{P}$ -martingale orthogonal to  $\mathbf{M}$  remains a martingale under  $\mathbb{P}^*$ .

## Minimal martingale measure (MMM)

As a result, it suffices to obtain a representation of  $\xi^F$  in (1) in order to obtain LRM.

A martingale measure  $\mathbb{P}^* \sim \mathbb{P}$  is called **minimal** if any square-integrable  $\mathbb{P}$ -martingale orthogonal to  $\mathbf{M}$  remains a martingale under  $\mathbb{P}^*$ .

We define  $\mathbf{Z} := \mathcal{E}\left(-\int \boldsymbol{\Lambda} d\mathbf{M}\right)$ , where  $\mathcal{E}(\mathbf{Y})$  represents the stochastic exponential of  $\mathbf{Y}$ .

Note that  $\mathbf{Z}$  is a solution to the SDE  $d\mathbf{Z}_t = -\boldsymbol{\Lambda}_t \mathbf{Z}_{t-} d\mathbf{M}_t$ .



## Minimal martingale measure (MMM)

As a result, it suffices to obtain a representation of  $\xi^F$  in (1) in order to obtain LRM.

A martingale measure  $\mathbb{P}^* \sim \mathbb{P}$  is called **minimal** if any square-integrable  $\mathbb{P}$ -martingale orthogonal to  $\mathbf{M}$  remains a martingale under  $\mathbb{P}^*$ .

We define  $\mathbf{Z} := \mathcal{E}\left(-\int \boldsymbol{\Lambda} d\mathbf{M}\right)$ , where  $\mathcal{E}(\mathbf{Y})$  represents the stochastic exponential of  $\mathbf{Y}$ .

Note that  $\mathbf{Z}$  is a solution to the SDE  $d\mathbf{Z}_t = -\boldsymbol{\Lambda}_t \mathbf{Z}_{t-} d\mathbf{M}_t$ .

Under Assumption 1, if  $\mathbf{Z}$  is a positive square integrable martingale, then an MMM  $\mathbb{P}^*$  exists with  $d\mathbb{P}^* = \mathbf{Z}_T d\mathbb{P}$ .

## Precedence research<sup>2</sup>

We consider a **Lévy market** as follows:

---

<sup>2</sup>Arai, T., Suzuki, R.: Local risk minimization for Lévy markets. *International Journal of Financial Engineering*, 2, 1550015 (2015)

## Precedence research<sup>2</sup>

We consider a **Lévy market** as follows:

Assume that  $\mathbf{S}$  is given by a solution to the following SDE:

$$d\mathbf{S}_t = \mathbf{S}_{t-} \left[ \alpha_t dt + \beta_t d\mathbf{W}_t + \int_{\mathbb{R}_0} \gamma_{t,z} \tilde{\mathbf{N}}(dt, dz) \right], \quad \mathbf{S}_0 > \mathbf{0},$$

where  $\alpha$ ,  $\beta$  and  $\gamma$  are predictable processes.

Here,  $\mathbf{W}$  is a **1**-dimensional Brownian motion,  $\mathbf{N}$  is a Poisson random measure,  $\nu$  is its Lévy measure and  $\tilde{\mathbf{N}}(dt, d\mathbf{x}) = \mathbf{N}(dt, d\mathbf{x}) - \nu(d\mathbf{x})dt$ .

---

<sup>2</sup>Arai, T., Suzuki, R.: Local risk minimization for Lévy markets. International Journal of Financial Engineering, 2, 1550015 (2015)

## Precedence research (cont'd)

Under some assumptions, A. and Suzuki gave the following expression of **locally risk minimizing (LRM)** strategy  $\xi^F$  for claim  $F$ :

$$\xi_t^F := \frac{\Lambda_t}{\alpha_t} \left\{ h_t^0 \beta_t + \int_{\mathbb{R}_0} h_{t,z}^1 \gamma_{t,z} \nu(dz) \right\},$$

where  $\Lambda_t := \frac{\alpha_t}{S_{t-}(\beta_t^2 + \int_{\mathbb{R}_0} \gamma_{t,z}^2 \nu(dz))}$ ,  $u_t := \Lambda_t S_{t-} \beta_t$ ,  $\theta_{t,z} := \Lambda_t S_{t-} \gamma_{t,z}$ .

Moreover,

$$h_t^0 := \mathbb{E}_{\mathbb{P}^*} \left[ D_{t,0} F - F \left[ \int_0^T D_{t,0} u_s dW_s^{\mathbb{P}^*} + \int_0^T \int_{\mathbb{R}_0} \frac{D_{t,0} \theta_{s,x}}{1 - \theta_{s,x}} \tilde{N}^{\mathbb{P}^*}(ds, dx) \right] \middle| \mathcal{F}_{t-} \right],$$

$$h_{t,z}^1 := \mathbb{E}_{\mathbb{P}^*} [F(H_{t,z}^* - 1) + z H_{t,z}^* D_{t,z} F | \mathcal{F}_{t-}],$$

where  $H_{t,z}^* := \exp\{z D_{t,z} \log Z_T - \log(1 - \theta_{t,z})\}$ .

## Precedence research (cont'd)

However, we need to assume the following:

## Assumption (A)

- ①  $u, u^2 \in \mathbb{L}_0^{1,2}$ ; and  $2u_s D_{t,z} u_s + z(D_{t,z} u_s)^2 \in L^2(\mathbf{q} \times \mathbb{P})$  for a.e.  $s \in [0, T]$ .
- ②  $\theta + \log(1 - \theta) \in \widetilde{\mathbb{L}}_1^{1,2}$ , and  $\log(1 - \theta) \in \mathbb{L}_1^{1,2}$
- ③ For  $\mathbf{q}$ -a.e.  $(s, x) \in [0, T] \times \mathbb{R}_0$ ,  $\exists \varepsilon_{s,x} \in (0, 1)$  such that  $\theta_{s,x} < 1 - \varepsilon_{s,x}$ .
- ④  $Z_T \left\{ D_{t,0} \log Z_T 1_{\{0\}}(z) + \frac{e^{zD_{t,z} \log Z_T} - 1}{z} 1_{\mathbb{R}_0}(z) \right\} \in L^2(\mathbf{q} \times \mathbb{P})$ .
- ⑤  $F \in \mathbb{D}^{1,2}$ ; and  $Z_T D_{t,z} F + F D_{t,z} Z_T + z D_{t,z} F \cdot D_{t,z} Z_T \in L^2(\mathbf{q} \times \mathbb{P})$ .
- ⑥  $FH_{t,z}^*, H_{t,z}^* D_{t,z} F \in L^1(\mathbb{P}^*)$  for  $\mathbf{q}$ -a.e.  $(t, z) \in [0, T] \times \mathbb{R}$ ,  
where  $H_{t,z}^* := \exp\{zD_{t,z} \log Z_T - \log(1 - \theta_{t,z})\}$ .

## Precedence Research (cont'd)

### Deterministic coefficients case

In the case where  $\alpha$ ,  $\beta$ , and  $\gamma$  are given by **deterministic** functions satisfying the following three conditions, if condition 5 in Assumption (A) and  $\mathbf{Z}_T \mathbf{F} \in L^2(\mathbb{P})$  are satisfied, then  $\xi^F$  is given as

$$\xi_t^F = \frac{\beta_t \mathbb{E}_{\mathbb{P}^*} [D_{t,0} \mathbf{F} | \mathcal{F}_{t-}] + \int_{\mathbb{R}_0} \mathbb{E}_{\mathbb{P}^*} [z D_{t,z} \mathbf{F} | \mathcal{F}_{t-}] \gamma_{t,z} \nu(dz)}{S_{t-} \left( \beta_t^2 + \int_{\mathbb{R}_0} \gamma_{t,z}^2 \nu(dz) \right)}.$$

### Conditions on $\alpha$ , $\beta$ , and $\gamma$

- 1  $\gamma_{t,z} > -1$ ,  $dt\nu(dz)$ -a.e.
- 2  $\sup_{t \in [0, T]} (|\alpha_t| + \beta_t^2 + \int_{\mathbb{R}_0} \gamma_{t,z}^2 \nu(dz)) < C$  for some  $C > 0$ .
- 3 We can find a positive number  $\varepsilon$  such that

$$\beta_t^2 + \int_{\mathbb{R}_0} \gamma_{t,z}^2 \nu(dz) > \varepsilon \text{ and } \frac{\alpha_t \gamma_{t,z}}{\beta_t^2 + \int_{\mathbb{R}_0} \gamma_{t,z}^2 \nu(dz)} < 1 - \varepsilon.$$

- 1 Local risk-minimization
- 2 Barndorff-Nielsen and Shephard models**
- 3 Main results
- 4 Numerical experiments

# BNS models

In this talk, we calculate **locally risk-minimizing (LRM) strategies** for **Barndorff-Nielsen and Shephard (BNS) models** <sup>3</sup>:

$$S_t = S_0 \exp \left\{ \int_0^t \left( \mu - \frac{1}{2} \sigma_s^2 \right) ds + \int_0^t \sigma_s dW_s + \rho H_{\lambda t} \right\}.$$

where  $S_0 > 0$ ,  $\rho \leq 0$ ,  $\mu \in \mathbb{R}$ ,  $\lambda > 0$ ,  $H$  is a subordinator without drift, and

---

<sup>3</sup>Barndorff-Nielsen, O.E., Shephard, N.: Non-Gaussian Ornstein-Uhlenbeck based models and some of their uses in financial econometrics. J.R. Statistic. Soc. 63, 167–241 (2001)



# BNS models

In this talk, we calculate **locally risk-minimizing (LRM) strategies** for **Barndorff-Nielsen and Shephard (BNS) models**<sup>3</sup>:

$$S_t = S_0 \exp \left\{ \int_0^t \left( \mu - \frac{1}{2} \sigma_s^2 \right) ds + \int_0^t \sigma_s dW_s + \rho H_{\lambda t} \right\}.$$

where  $S_0 > 0$ ,  $\rho \leq 0$ ,  $\mu \in \mathbb{R}$ ,  $\lambda > 0$ ,  $H$  is a subordinator without drift, and

$$d\sigma_t^2 = -\lambda \sigma_t^2 dt + dH_{\lambda t}, \quad \sigma_0^2 > 0.$$

<sup>3</sup>Barndorff-Nielsen, O.E., Shephard, N.: Non-Gaussian Ornstein-Uhlenbeck based models and some of their uses in financial econometrics. J.R. Statistic. Soc. 63, 167–241 (2001)

# BNS models

In this talk, we calculate **locally risk-minimizing (LRM) strategies** for **Barndorff-Nielsen and Shephard (BNS) models**<sup>3</sup>:

$$S_t = S_0 \exp \left\{ \int_0^t \left( \mu - \frac{1}{2} \sigma_s^2 \right) ds + \int_0^t \sigma_s dW_s + \rho H_{\lambda t} \right\}.$$

where  $S_0 > 0$ ,  $\rho \leq 0$ ,  $\mu \in \mathbb{R}$ ,  $\lambda > 0$ ,  $H$  is a subordinator without drift, and

$$d\sigma_t^2 = -\lambda \sigma_t^2 dt + dH_{\lambda t}, \quad \sigma_0^2 > 0.$$

$\rho H_{\lambda t}$  represents **leverage effect**.

<sup>3</sup>Barndorff-Nielsen, O.E., Shephard, N.: Non-Gaussian Ornstein-Uhlenbeck based models and some of their uses in financial econometrics. J.R. Statistic. Soc. 63, 167–241 (2001)

## BNS models (cont'd)

$\mathbf{S}$  is a solution to the following SDE:

$$d\mathbf{S}_t = \mathbf{S}_{t-} \left\{ \alpha dt + \sigma_t dW_t + \int_0^\infty (e^{\rho x} - 1) \tilde{N}(dt, dx) \right\},$$

where  $\alpha = \mu + \int_0^\infty (e^{\rho x} - 1) \nu(dx)$ .

Here,

$$(\mathbf{J}_t :=) H_{\lambda t} = \int_0^\infty x N([0, t], dx) \text{ and } \tilde{N}(dt, dx) = N(dt, dx) - \nu(dx) dt,$$

where  $\nu$  is the Lévy measure of  $\mathbf{J}$ .

## BNS models (cont'd)

$S$  is a solution to the following SDE:

$$dS_t = S_{t-} \left\{ \alpha dt + \sigma_t dW_t + \int_0^\infty (e^{\rho x} - 1) \tilde{N}(dt, dx) \right\},$$

where  $\alpha = \mu + \int_0^\infty (e^{\rho x} - 1) \nu(dx)$ .

Here,

$$(J_t :=) H_{\lambda t} = \int_0^\infty x N([0, t], dx) \text{ and } \tilde{N}(dt, dx) = N(dt, dx) - \nu(dx) dt,$$

where  $\nu$  is the Lévy measure of  $J$ .

Note that  $\sigma^2$  is represented as

$$\sigma_t^2 = e^{-\lambda t} \sigma_0^2 + \int_0^t e^{-\lambda(t-s)} dJ_s.$$

- 1 Local risk-minimization
- 2 Barndorff-Nielsen and Shephard models
- 3 Main results**
- 4 Numerical experiments

# Assumption

Denote  $L_t := \log(\mathbf{S}_t/\mathbf{S}_0)$  for  $t \in [0, T]$ , that is,

$$L_t = \int_0^t \left( \mu - \frac{1}{2} \sigma_s^2 \right) ds + \int_0^t \sigma_s dW_s + \rho J_t.$$

# Assumption

Denote  $L_t := \log(\mathbf{S}_t/\mathbf{S}_0)$  for  $t \in [0, T]$ , that is,

$$L_t = \int_0^t \left( \mu - \frac{1}{2} \sigma_s^2 \right) ds + \int_0^t \sigma_s dW_s + \rho J_t.$$

Assumption (BNS)

- ①  $\int_0^1 x \nu(dx) + \int_1^\infty \exp\{2(\mathcal{B}(T) \vee |\rho|)x\} \nu(dx) < \infty$ , where  $\mathcal{B}(t) := \frac{1-e^{-\lambda t}}{\lambda}$  for  $t \in [0, T]$ .
- ②  $\frac{\alpha}{e^{-\lambda T} \sigma_0^2 + C_\rho} > -1$ , where  $C_\rho := \int_0^\infty (e^{\rho x} - 1)^2 \nu(dx)$ .

# Assumption

Denote  $L_t := \log(\mathbf{S}_t/\mathbf{S}_0)$  for  $t \in [0, T]$ , that is,

$$L_t = \int_0^t \left( \mu - \frac{1}{2} \sigma_s^2 \right) ds + \int_0^t \sigma_s dW_s + \rho J_t.$$

Assumption (BNS)

- ①  $\int_0^1 x \nu(dx) + \int_1^\infty \exp\{2(\mathcal{B}(T) \vee |\rho|)x\} \nu(dx) < \infty$ , where  $\mathcal{B}(t) := \frac{1-e^{-\lambda t}}{\lambda}$  for  $t \in [0, T]$ .
  - ②  $\frac{\alpha}{e^{-\lambda T} \sigma_0^2 + C_\rho} > -1$ , where  $C_\rho := \int_0^\infty (e^{\rho x} - 1)^2 \nu(dx)$ .
- 
- ① Item 1 ensures  $\int_0^\infty x^2 \nu(dx) < \infty$ .



# Assumption

Denote  $L_t := \log(\mathbf{S}_t/\mathbf{S}_0)$  for  $t \in [0, T]$ , that is,

$$L_t = \int_0^t \left( \mu - \frac{1}{2} \sigma_s^2 \right) ds + \int_0^t \sigma_s dW_s + \rho J_t.$$

## Assumption (BNS)

- ①  $\int_0^1 x \nu(dx) + \int_1^\infty \exp\{2(\mathcal{B}(T) \vee |\rho|)x\} \nu(dx) < \infty$ , where  $\mathcal{B}(t) := \frac{1-e^{-\lambda t}}{\lambda}$  for  $t \in [0, T]$ .
- ②  $\frac{\alpha}{e^{-\lambda T} \sigma_0^2 + C_\rho} > -1$ , where  $C_\rho := \int_0^\infty (e^{\rho x} - 1)^2 \nu(dx)$ .

- ① Item 1 ensures  $\int_0^\infty x^2 \nu(dx) < \infty$ .
- ② By  $|e^{\rho x} - 1| \leq -\rho x$ , we have  $\int_0^\infty (e^{\rho x} - 1)^2 \nu(dx) \leq \int_0^\infty \rho^2 x^2 \nu(dx) < \infty$ .

# Assumption

Denote  $L_t := \log(\mathbf{S}_t/\mathbf{S}_0)$  for  $t \in [0, T]$ , that is,

$$L_t = \int_0^t \left( \mu - \frac{1}{2} \sigma_s^2 \right) ds + \int_0^t \sigma_s dW_s + \rho J_t.$$

## Assumption (BNS)

- ①  $\int_0^1 x \nu(dx) + \int_1^\infty \exp\{2(\mathcal{B}(T) \vee |\rho|)x\} \nu(dx) < \infty$ , where  $\mathcal{B}(t) := \frac{1-e^{-\lambda t}}{\lambda}$  for  $t \in [0, T]$ .
  - ②  $\frac{\alpha}{e^{-\lambda T} \sigma_0^2 + C_\rho} > -1$ , where  $C_\rho := \int_0^\infty (e^{\rho x} - 1)^2 \nu(dx)$ .
- 
- ① Item 1 ensures  $\int_0^\infty x^2 \nu(dx) < \infty$ .
  - ② By  $|e^{\rho x} - 1| \leq -\rho x$ , we have  $\int_0^\infty (e^{\rho x} - 1)^2 \nu(dx) \leq \int_0^\infty \rho^2 x^2 \nu(dx) < \infty$ .
  - ③ Item 2 ensures  $\frac{\alpha}{\sigma_t^2 + C_\rho} > -1$  for any  $t \in [0, T]$ .

# Representative examples of $\sigma^2$

## IG-OU

The first is the case where the Lévy measure  $\nu^H$  of the subordinator  $H$  is given as

$$\nu^H(dx) = \frac{a}{2\sqrt{2\pi}} x^{-\frac{3}{2}} (1 + b^2 x) e^{-\frac{1}{2}b^2 x} \mathbf{1}_{(0,\infty)}(x) dx$$

where  $a > 0$  and  $b > 0$ .

In this case, the invariant distribution of the squared volatility process  $\sigma^2$  follows an **inverse-Gaussian** distribution with parameters  $a > 0$  and  $b > 0$ .  $\sigma^2$  is called an **IG-OU** process.

# Representative examples of $\sigma^2$

## IG-OU

The first is the case where the Lévy measure  $\nu^H$  of the subordinator  $H$  is given as

$$\nu^H(dx) = \frac{a}{2\sqrt{2\pi}} x^{-\frac{3}{2}} (1 + b^2 x) e^{-\frac{1}{2} b^2 x} \mathbf{1}_{(0, \infty)}(x) dx$$

where  $a > 0$  and  $b > 0$ .

In this case, the invariant distribution of the squared volatility process  $\sigma^2$  follows an **inverse-Gaussian** distribution with parameters  $a > 0$  and  $b > 0$ .  $\sigma^2$  is called an **IG-OU** process.

If  $\frac{b^2}{2} > 2(\mathcal{B}(T) \vee |\rho|)$ , then item 1 of Assumption (BNS) is satisfied.

## Representative examples of $\sigma^2$ (cont'd)

### Gamma-OU

The second example is what we call **Gamma-OU** case, that is, the case where the invariant distribution of  $\sigma^2$  is given by a **Gamma** distribution with parameter  $\mathbf{a} > \mathbf{0}$  and  $\mathbf{b} > \mathbf{0}$ .

In this case,  $\nu^H$  is described as

$$\nu^H(dx) = \mathbf{a}\mathbf{b}e^{-\mathbf{b}x}\mathbf{1}_{(0,\infty)}(x)dx.$$

where  $\mathbf{a} > \mathbf{0}$  and  $\mathbf{b} > \mathbf{0}$ .

## Representative examples of $\sigma^2$ (cont'd)

### Gamma-OU

The second example is what we call **Gamma-OU** case, that is, the case where the invariant distribution of  $\sigma^2$  is given by a **Gamma** distribution with parameter  $\mathbf{a} > \mathbf{0}$  and  $\mathbf{b} > \mathbf{0}$ .

In this case,  $\nu^H$  is described as

$$\nu^H(dx) = \mathbf{a}\mathbf{b}e^{-\mathbf{b}x}\mathbf{1}_{(0,\infty)}(x)dx.$$

where  $\mathbf{a} > \mathbf{0}$  and  $\mathbf{b} > \mathbf{0}$ .

If  $\mathbf{b} > 2(\mathcal{B}(\mathbf{T}) \vee |\rho|)$ , then item 1 of Assumption (BNS) is satisfied.

## Minimal martingale measure

Now, we consider the following SDE:

$$dZ_t = -Z_{t-}\Lambda_t dM_t, \quad Z_0 = 1,$$

where  $\Lambda_s := \frac{1}{S_{s-}} \frac{\alpha}{\sigma_s^2 + C_\rho}$  and  $C_\rho := \int_0^\infty (e^{\rho x} - 1)^2 \nu(dx)$ .

## Minimal martingale measure

Now, we consider the following SDE:

$$dZ_t = -Z_{t-}\Lambda_t dM_t, \quad Z_0 = 1,$$

where  $\Lambda_s := \frac{1}{S_{s-}} \frac{\alpha}{\sigma_s^2 + C_\rho}$  and  $C_\rho := \int_0^\infty (e^{\rho x} - 1)^2 \nu(dx)$ . Denoting

$$u_s := \Lambda_s S_{s-} \sigma_s = \frac{\alpha \sigma_s}{\sigma_s^2 + C_\rho} \quad \text{and} \quad \theta_{s,x} := \Lambda_s S_{s-} (e^{\rho x} - 1) = \frac{\alpha (e^{\rho x} - 1)}{\sigma_s^2 + C_\rho}$$

for  $s \in [0, T]$  and  $x \in (0, \infty)$ , we have  $\Lambda_t dM_t = u_t dW_t + \int_0^\infty \theta_{t,z} \tilde{N}(dt, dz)$ ; and



## Minimal martingale measure

Now, we consider the following SDE:

$$dZ_t = -Z_t \Lambda_t dM_t, \quad Z_0 = 1,$$

where  $\Lambda_s := \frac{1}{S_s - \sigma_s^2 + C_\rho}$  and  $C_\rho := \int_0^\infty (e^{\rho x} - 1)^2 \nu(dx)$ . Denoting

$$u_s := \Lambda_s S_s \sigma_s = \frac{\alpha \sigma_s}{\sigma_s^2 + C_\rho} \quad \text{and} \quad \theta_{s,x} := \Lambda_s S_s (e^{\rho x} - 1) = \frac{\alpha (e^{\rho x} - 1)}{\sigma_s^2 + C_\rho}$$

for  $s \in [0, T]$  and  $x \in (0, \infty)$ , we have  $\Lambda_t dM_t = u_t dW_t + \int_0^\infty \theta_{t,z} \tilde{N}(dt, dz)$ ; and

$$Z_t = \exp \left\{ - \int_0^t u_s dW_s - \frac{1}{2} \int_0^t u_s^2 ds + \int_0^t \int_0^\infty \log(1 - \theta_{s,x}) \tilde{N}(ds, dx) + \int_0^t \int_0^\infty (\log(1 - \theta_{s,x}) + \theta_{s,x}) \nu(dx) ds \right\}.$$

## Minimal martingale measure

Now, we consider the following SDE:

$$dZ_t = -Z_t \Lambda_t dM_t, \quad Z_0 = 1,$$

where  $\Lambda_s := \frac{1}{S_s - \sigma_s^2 + C_\rho}$  and  $C_\rho := \int_0^\infty (e^{\rho x} - 1)^2 \nu(dx)$ . Denoting

$$u_s := \Lambda_s S_s \sigma_s = \frac{\alpha \sigma_s}{\sigma_s^2 + C_\rho} \quad \text{and} \quad \theta_{s,x} := \Lambda_s S_s (e^{\rho x} - 1) = \frac{\alpha (e^{\rho x} - 1)}{\sigma_s^2 + C_\rho}$$

for  $s \in [0, T]$  and  $x \in (0, \infty)$ , we have  $\Lambda_t dM_t = u_t dW_t + \int_0^\infty \theta_{t,z} \tilde{N}(dt, dz)$ ; and

$$Z_t = \exp \left\{ - \int_0^t u_s dW_s - \frac{1}{2} \int_0^t u_s^2 ds + \int_0^t \int_0^\infty \log(1 - \theta_{s,x}) \tilde{N}(ds, dx) + \int_0^t \int_0^\infty (\log(1 - \theta_{s,x}) + \theta_{s,x}) \nu(dx) ds \right\}.$$

### Proposition

The process  $Z$  is a martingale with  $Z_T \in L^2(\mathbb{P})$ .

# LRM for put options

## Theorem

For  $K > 0$ , LRM  $\xi^{(K-S_T)^+}$  of put option  $(K - S_T)^+$  is represented as

$$\xi_t^{(K-S_T)^+} = \frac{1}{S_{t-}(\sigma_t^2 + C_\rho)} \left\{ \sigma_t^2 \mathbb{E}_{\mathbb{P}^*} [-1_{\{S_T < K\}} S_T | \mathcal{F}_{t-}] \right. \\ \left. + \int_0^\infty \mathbb{E}_{\mathbb{P}^*} [(K - S_T)^+ (H_{t,z}^* - 1) + z H_{t,z}^* D_{t,z} (K - S_T)^+ | \mathcal{F}_{t-}] \right. \\ \left. \times (e^{\rho z} - 1) \nu(dz) \right\},$$

where  $H_{t,z}^* := \exp\{z D_{t,z} \log Z_T - \log(1 - \theta_{t,z})\}$  for  $(t, z) \in [0, T] \times (0, \infty)$ .

# Reminder

## Assumption (A)

- 1  $u, u^2 \in \mathbb{L}_0^{1,2}$ ; and  $2u_s D_{t,z} u_s + z(D_{t,z} u_s)^2 \in L^2(q \times \mathbb{P})$  for a.e.  $s \in [0, T]$ .
- 2  $\theta + \log(1 - \theta) \in \widetilde{\mathbb{L}}_1^{1,2}$ , and  $\log(1 - \theta) \in \mathbb{L}_1^{1,2}$
- 3 For  $q$ -a.e.  $(s, x) \in [0, T] \times \mathbb{R}_0$ ,  $\exists \varepsilon_{s,x} \in (0, 1)$  such that  $\theta_{s,x} < 1 - \varepsilon_{s,x}$ .
- 4  $Z_T \left\{ D_{t,0} \log Z_T \mathbf{1}_{\{0\}}(z) + \frac{e^{z D_{t,z} \log Z_T - 1}}{z} \mathbf{1}_{\mathbb{R}_0}(z) \right\} \in L^2(q \times \mathbb{P})$ .
- 5  $F \in \mathbb{D}^{1,2}$ ; and  $Z_T D_{t,z} F + F D_{t,z} Z_T + z D_{t,z} F \cdot D_{t,z} Z_T \in L^2(q \times \mathbb{P})$ .
- 6  $FH_{t,z}^*, H_{t,z}^* D_{t,z} F \in L^1(\mathbb{P}^*)$  for  $q$ -a.e.  $(t, z) \in [0, T] \times \mathbb{R}$ .

# LRM for call options

## Corollary

LRM for call option  $(S_T - K)^+$  is given as  $\xi^{(S_T - K)^+} = \mathbf{1} + \xi^{(K - S_T)^+}$ .

# LRM for call options

## Corollary

LRM for call option  $(S_T - K)^+$  is given as  $\xi^{(S_T - K)^+} = \mathbf{1} + \xi^{(K - S_T)^+}$ .

Proof

$$\begin{aligned}
 (S_T - K)^+ &= S_T - K + (K - S_T)^+ \\
 &= S_0 + \int_0^T dS_t - K + \mathbb{E}_{\mathbb{P}^*}[(K - S_T)^+] + \int_0^T \xi_t^{(K - S_T)^+} dS_t + L_T^{(K - S_T)^+} \\
 &= \mathbb{E}_{\mathbb{P}^*}[S_T - K + (K - S_T)^+] + \int_0^T \left(1 + \xi_t^{(K - S_T)^+}\right) dS_t + L_T^{(K - S_T)^+} \\
 &= \mathbb{E}_{\mathbb{P}^*}[(S_T - K)^+] + \int_0^T \left(1 + \xi_t^{(K - S_T)^+}\right) dS_t + L_T^{(K - S_T)^+}.
 \end{aligned}$$

This is an FS-decomposition of  $(S_T - K)^+$  since  $\mathbf{1} \in \Theta_S$ .

# Proof of Theorem

In order to see condition 4, we need to show  $Z_T \in \mathbb{D}^{1,2}$ .

Condition 4:  $Z_T \left\{ D_{t,0} \log Z_T \mathbf{1}_{\{0\}}(z) + \frac{e^{zD_{t,z} \log Z_T} - 1}{z} \mathbf{1}_{\mathbb{R}_0}(z) \right\} \in L^2(\mathbf{q} \times \mathbb{P})$ .

Reminder:

$$dZ_t = -Z_{t-} \left\{ u_t dW_t + \int_0^\infty \theta_{t,z} \tilde{N}(dt, dz) \right\}, \quad Z_0 = 1,$$

where  $u_s = \frac{\alpha \sigma_s}{\sigma_s^2 + C_\rho}$ ,  $\theta_{s,x} = \frac{\alpha(e^{\rho x} - 1)}{\sigma_s^2 + C_\rho}$  and  $C_\rho := \int_0^\infty (e^{\rho x} - 1)^2 \nu(dx)$ .

## Proof of Theorem (cont'd)

For  $t \in [0, T]$ , we define  $Z_t^{(0)} := 1$  and

$$Z_t^{(n+1)} := 1 - \int_0^t Z_{s-}^{(n)} u_s dW_s - \int_0^t \int_0^\infty Z_{s-}^{(n)} \theta_{s,x} \tilde{N}(ds, dx)$$

for  $n \geq 0$ .



## Proof of Theorem (cont'd)

For  $t \in [0, T]$ , we define  $Z_t^{(0)} := 1$  and

$$Z_t^{(n+1)} := 1 - \int_0^t Z_{s-}^{(n)} u_s dW_s - \int_0^t \int_0^\infty Z_{s-}^{(n)} \theta_{s,x} \tilde{N}(ds, dx)$$

for  $n \geq 0$ .

Besides, we denote, for  $n \geq 0$ ,

$$\phi_n(t) := \mathbb{E} \left[ \int_{[0,t] \times [0,\infty)} (D_{r,z} Z_t^{(n)})^2 q(dr, dz) \right].$$

Note that  $\phi_0(t) \equiv 0$ .

## Proof of Theorem (cont'd)

For  $t \in [0, T]$ , we define  $Z_t^{(0)} := 1$  and

$$Z_t^{(n+1)} := 1 - \int_0^t Z_{s-}^{(n)} u_s dW_s - \int_0^t \int_0^\infty Z_{s-}^{(n)} \theta_{s,x} \tilde{N}(ds, dx)$$

for  $n \geq 0$ .

Besides, we denote, for  $n \geq 0$ ,

$$\phi_n(t) := \mathbb{E} \left[ \int_{[0,t] \times [0,\infty)} (D_{r,z} Z_t^{(n)})^2 q(dr, dz) \right].$$

Note that  $\phi_0(t) \equiv 0$ .

### Lemma 1

We have  $Z_t^{(n)} \in \mathbb{D}^{1,2}$  for every  $n \geq 0$  and any  $t \in [0, T]$ . Moreover, there exist constants  $k_1 > 0$  and  $k_2 > 0$  such that

$$\phi_{n+1}(t) \leq k_1 + k_2 \int_0^t \phi_n(s) ds$$

for every  $n \geq 0$  and any  $t \in [0, T]$ .

- 1 Local risk-minimization
- 2 Barndorff-Nielsen and Shephard models
- 3 Main results
- 4 Numerical experiments**

# Setting

We treat **Gamma-OU model**:  $\nu(dx) = ab\lambda e^{-bx} \mathbf{1}_{(0,\infty)}(x) dx$ , where  $a > 0$ ,  $b > 0$ .

---

<sup>4</sup>Schoutens, W.: Lévy Processes in Finance: Pricing Financial Derivatives. John Wiley & Sons, Hoboken (2003)

# Setting

We treat **Gamma-OU model**:  $\nu(dx) = ab\lambda e^{-bx} \mathbf{1}_{(0,\infty)}(x) dx$ , where  $a > 0$ ,  $b > 0$ .

We use a parameter set estimated in Schoutens' text book<sup>4</sup>.

---

<sup>4</sup>Schoutens, W.: Lévy Processes in Finance: Pricing Financial Derivatives. John Wiley & Sons, Hoboken (2003)

# Setting

We treat **Gamma-OU model**:  $\nu(dx) = ab\lambda e^{-bx} \mathbf{1}_{(0,\infty)}(x) dx$ , where  $a > 0$ ,  $b > 0$ .

We use a parameter set estimated in Schoutens' text book<sup>4</sup>.

Fix  $T = 1$ ,  $r = 0.019$  and  $q = 0.012$ .

The asset price and the squared volatility at time  $t$  are fixed to  $S_t = 1124.47$  and  $\sigma_t^2 = 0.0145$ , respectively.

---

<sup>4</sup>Schoutens, W.: Lévy Processes in Finance: Pricing Financial Derivatives. John Wiley & Sons, Hoboken (2003)

# Setting

We treat **Gamma-OU model**:  $\nu(dx) = ab\lambda e^{-bx} \mathbf{1}_{(0,\infty)}(x) dx$ , where  $a > 0$ ,  $b > 0$ .

We use a parameter set estimated in Schoutens' text book<sup>4</sup>.

Fix  $T = 1$ ,  $r = 0.019$  and  $q = 0.012$ .

The asset price and the squared volatility at time  $t$  are fixed to  $S_t = 1124.47$  and  $\sigma_t^2 = 0.0145$ , respectively.

$\rho = -1.2606$ ,  $\lambda = 0.5783$ ,  $a = 1.4338$ ,  $b = 11.6641$ .

---

<sup>4</sup>Schoutens, W.: Lévy Processes in Finance: Pricing Financial Derivatives. John Wiley & Sons, Hoboken (2003)

## Setting

We treat **Gamma-OU model**:  $\nu(dx) = ab\lambda e^{-bx} \mathbf{1}_{(0,\infty)}(x) dx$ , where  $a > 0$ ,  $b > 0$ .

We use a parameter set estimated in Schoutens' text book<sup>4</sup>.

Fix  $T = 1$ ,  $r = 0.019$  and  $q = 0.012$ .

The asset price and the squared volatility at time  $t$  are fixed to  $S_t = 1124.47$  and  $\sigma_t^2 = 0.0145$ , respectively.

$\rho = -1.2606$ ,  $\lambda = 0.5783$ ,  $a = 1.4338$ ,  $b = 11.6641$ .

Suppose that the discounted asset price process  $e^{-(r-q)t} S_t$  is a martingale. Hence,  $\mu$  is given as

$$\mu = r - q + \int_0^\infty (1 - e^{\rho x}) \nu(dx) = r - q - \frac{a\lambda\rho}{b - \rho}.$$

<sup>4</sup>Schoutens, W.: Lévy Processes in Finance: Pricing Financial Derivatives. John Wiley & Sons, Hoboken (2003)



## Setting (cont'd)

A., Imai and Suzuki <sup>5</sup> developed a numerical scheme of LRM for exponential Lévy models using the **Carr-Madan approach** <sup>6</sup>, which is a numerical method for option prices based on the **fast Fourier transform (FFT)**.

We consider a call option with strike price  $K$ . Since  $H_{t,z}^* = \mathbf{1}$  and  $Z_T = \mathbf{1}$ , we have

$$\begin{aligned} \xi_t^{(S_T - K)^+} &= \frac{e^{-(r-q)(T-t)}}{S_{t-}(\sigma_t^2 + C_\rho)} \left( \sigma_t^2 \mathbb{E}[S_T \mathbf{1}_{\{S_T \geq K\}} | \mathcal{F}_{t-}] \right. \\ &\quad \left. + \int_0^\infty \mathbb{E} \left[ (S_T e^{zD_{t,z}L_T} - K)^+ - (S_T - K)^+ | \mathcal{F}_{t-} \right] (e^{\rho z} - 1) \nu(dz) \right) \\ &=: \frac{\sigma_t^2 I_1 + I_2}{S_t(\sigma_t^2 + C_\rho)}. \end{aligned}$$

<sup>5</sup>Arai, T., Imai, Y., Suzuki, R.: Numerical analysis on local risk-minimization for exponential Lévy models, International Journal of Theoretical and Applied Finance vol.19, 1650008 (2016)

<sup>6</sup>Carr, P., D. Madan: Option valuation using the fast Fourier transform. Journal of Computational Finance, 2, 61–73 (1999)

## Reminder: main result

$$\xi_t^{(K-S_T)^+} = \frac{1}{S_{t-}(\sigma_t^2 + C_\rho)} \left\{ \sigma_t^2 \mathbb{E}_{\mathbb{P}^*} [-1_{\{S_T < K\}} S_T | \mathcal{F}_{t-}] \right. \\ \left. + \int_0^\infty \mathbb{E}_{\mathbb{P}^*} [(K - S_T)^+ (H_{t,z}^* - 1) + z H_{t,z}^* D_{t,z} (K - S_T)^+ | \mathcal{F}_{t-}] \right. \\ \left. \times (e^{\rho z} - 1) \nu(dz) \right\},$$

where  $H_{t,z}^* := \exp\{z D_{t,z} \log Z_T - \log(1 - \theta_{t,z})\}$  for  $(t, z) \in [0, T] \times (0, \infty)$ .

$$\xi^{(S_T - K)^+} = 1 + \xi^{(K - S_T)^+}.$$

# Characteristic function of $L_T$

$$\begin{aligned}
 \phi(\vartheta) &:= \mathbb{E}[\exp(i\vartheta L_T) | \mathcal{S}_t, \sigma_t^2] \\
 &= \exp\left(i\vartheta (L_t + \mu(T-t)) - (\vartheta^2 + i\vartheta) \frac{\mathcal{B}(T-t)}{2} \sigma_t^2 \right. \\
 &\quad \left. + \frac{a}{b-f_2} \left[ b \log\left(\frac{b-f_1}{b-i\vartheta\rho}\right) + f_2 \lambda(T-t) \right] \right)
 \end{aligned}$$

for  $\vartheta \in \mathbb{C}$ , where

$$f_1 := i\vartheta\rho - \frac{1}{2}(\vartheta^2 + i\vartheta)\lambda\mathcal{B}(T-t) \quad \text{and} \quad f_2 := i\vartheta\rho - \frac{1}{2}(\vartheta^2 + i\vartheta).$$

Recall that  $\mathcal{B}(t) = \frac{1-e^{-\lambda t}}{\lambda}$  for  $t \in [0, T]$ .

$I_1$ 

$$I_1 = e^{-(r-q)(T-t)} \mathbb{E}[\mathbf{S}_T \mathbf{1}_{\{\mathbf{S}_T \geq K\}} | \mathcal{F}_{t-}] = \frac{e^{-(r-q)(T-t)}}{\pi} \int_0^\infty K^{-i\zeta+1} \frac{\phi(\zeta)}{i\zeta-1} d\nu, \quad (2)$$

where  $\zeta := \nu - i\delta$ , and  $\delta$  is a real number satisfying

$$\sup_{t \leq s < T} \left\{ \frac{1}{2} - \frac{\rho}{\mathcal{B}(T-s)} - \sqrt{D_s} \right\} < \delta < \inf_{t \leq s < T} \left\{ \frac{1}{2} - \frac{\rho}{\mathcal{B}(T-s)} + \sqrt{D_s} \right\}.$$

Here,

$$D_s := \left( -\frac{1}{2} + \frac{\rho}{\mathcal{B}(T-s)} \right)^2 + \frac{2\hat{\vartheta}}{\mathcal{B}(T-s)} \text{ and } \hat{\vartheta} := \sup \left\{ \vartheta \mid \int_0^\infty (e^{\vartheta x} - 1) \nu(dx) < \infty \right\}.$$

$I_1$ 

$$I_1 = e^{-(r-q)(T-t)} \mathbb{E}[\mathbf{S}_T \mathbf{1}_{\{\mathbf{S}_T \geq K\}} | \mathcal{F}_{t-}] = \frac{e^{-(r-q)(T-t)}}{\pi} \int_0^\infty K^{-i\zeta+1} \frac{\phi(\zeta)}{i\zeta-1} d\mathbf{v}, \quad (2)$$

where  $\zeta := \mathbf{v} - i\delta$ , and  $\delta$  is a real number satisfying

$$\sup_{t \leq s < T} \left\{ \frac{1}{2} - \frac{\rho}{\mathcal{B}(T-s)} - \sqrt{D_s} \right\} < \delta < \inf_{t \leq s < T} \left\{ \frac{1}{2} - \frac{\rho}{\mathcal{B}(T-s)} + \sqrt{D_s} \right\}.$$

Here,

$$D_s := \left( -\frac{1}{2} + \frac{\rho}{\mathcal{B}(T-s)} \right)^2 + \frac{2\hat{\vartheta}}{\mathcal{B}(T-s)} \text{ and } \hat{\vartheta} := \sup \left\{ \vartheta \mid \int_0^\infty (e^{\vartheta x} - 1) \nu(dx) < \infty \right\}.$$

Note that the RHS of (2) is independent of the choice of  $\delta$ .

As a result, since the integrand of (2) is given by the product of  $K^{-i\zeta+1}$  and a function of  $\zeta$ , we can compute  $I_1$  through the FFT.

$I_2$ 

Reminder:

$$I_2 = e^{-(r-q)(T-t)} \int_0^\infty \mathbb{E} \left[ \left( S_T e^{zD_{t,z}L_T} - K \right)^+ - (S_T - K)^+ | \mathcal{F}_{t-} \right] (e^{\rho z} - 1) \nu(dz).$$

$$\text{Note that } \mathbb{E}[(S_T - K)^+ | S_t, \sigma_t^2] = \frac{1}{\pi} \int_0^\infty \frac{K^{-i\zeta+1} \phi(\zeta)}{(i\zeta-1)i\zeta} d\nu$$

$I_2$ 

Reminder:

$$I_2 = e^{-(r-q)(T-t)} \int_0^\infty \mathbb{E} \left[ \left( S_T e^{zD_{t,z}L_T} - K \right)^+ - (S_T - K)^+ | \mathcal{F}_{t-} \right] (e^{\rho z} - 1) \nu(dz).$$

Note that  $\mathbb{E}[(S_T - K)^+ | S_t, \sigma_t^2] = \frac{1}{\pi} \int_0^\infty \frac{K^{-i\zeta+1} \phi(\zeta)}{(i\zeta-1)i\zeta} d\nu$

$$\begin{aligned} & \frac{S_T}{S_t} \exp(zD_{t,z}L_T) \\ &= \exp \left( \mu(T-t) - \frac{1}{2} \int_t^T \sigma_s^2 ds + \int_t^T \sigma_s dW_s + \rho \int_t^T dJ_s \right. \\ & \quad \left. - \frac{z}{2} \mathcal{B}(T-t) + \int_t^T \left( \sqrt{\sigma_s^2 + z e^{-\lambda(s-t)}} - \sigma_s \right) dW_s + \rho z \right) \\ &= \exp \left( \mu(T-t) - \frac{1}{2} \int_t^T \sigma_{s,z}^2 ds + \int_t^T \sigma_{s,z} dW_s + \rho \int_t^T dJ_s + \rho z \right) \end{aligned}$$

where  $\sigma_{s,z}^2 := \sigma_s^2 + z e^{-\lambda(s-t)}$  for  $(s, z) \in [t, T] \times (0, \infty)$ .

$I_2$  (cont'd)

Denoting

$$L_s^{(z)} := \int_t^s \left( \mu - \frac{1}{2} \sigma_{u,z}^2 \right) du + \int_t^s \sigma_{u,z} dW_u + \rho \int_t^s dJ_u$$

for  $(\mathbf{s}, \mathbf{z}) \in [t, T] \times (0, \infty)$ , we have

$$\mathbf{S}_T \exp(\mathbf{z} D_{t,z} L_T) = \mathbf{S}_t \exp(L_T^{(z)} + \rho \mathbf{z}).$$



$l_2$  (cont'd)

Denoting

$$L_s^{(z)} := \int_t^s \left( \mu - \frac{1}{2} \sigma_{u,z}^2 \right) du + \int_t^s \sigma_{u,z} dW_u + \rho \int_t^s dJ_u$$

for  $(s, z) \in [t, T] \times (0, \infty)$ , we have

$$S_T \exp(z D_{t,z} L_T) = S_t \exp(L_T^{(z)} + \rho z).$$

$(\sigma_{s,z}^2)_{t \leq s \leq T}$  is a solution to the same SDE as  $\sigma^2$ , that is,  $d\sigma_t^2 = -\lambda \sigma_t^2 dt + dJ_t$  with initial condition  $\sigma_{t,z}^2 = \sigma_t^2 + z$ .

We denote

$$\begin{aligned} \phi^{(z)}(\vartheta) &:= \mathbb{E} \left[ \exp(i\vartheta L_T^{(z)}) | S_t, \sigma_t^2 \right] S_t^{i\vartheta} = \mathbb{E} \left[ \exp(i\vartheta L_T) | S_t, \sigma_t^2 + z \right] \\ &= \phi(\vartheta) \exp \left( -(\vartheta^2 + i\vartheta) \frac{\mathcal{B}(T-t)}{2} z \right). \end{aligned}$$

$I_2$  (cont'd)

$$\begin{aligned}
& e^{(r-q)(T-t)} I_2 \\
&= \int_0^\infty \mathbb{E} \left[ \left( S_T e^{zD_{t,z}L_T} - K \right)^+ - (S_T - K)^+ | \mathcal{F}_{t-} \right] (e^{\rho z} - 1) \nu(dz) \\
&= \int_0^\infty \mathbb{E} \left[ \left( S_t \exp(L_T^{(z)} + \rho z) - K \right)^+ - (S_T - K)^+ \middle| S_t, \sigma_t^2 \right] (e^{\rho z} - 1) \nu(dz) \\
&= \int_0^\infty \left( \frac{e^{\rho z}}{\pi} \int_0^\infty (K e^{-\rho z})^{-i\zeta+1} \frac{\phi^{(z)}(\zeta)}{(i\zeta - 1)i\zeta} dv - \frac{1}{\pi} \int_0^\infty \frac{K^{-i\zeta+1} \phi(\zeta)}{(i\zeta - 1)i\zeta} dv \right) (e^{\rho z} - 1) \nu(dz) \\
&= \int_0^\infty \frac{1}{\pi} \int_0^\infty \frac{K^{-i\zeta+1} \phi(\zeta)}{(i\zeta - 1)i\zeta} \left( e^{i\rho z \zeta} \exp\left(-(\zeta^2 + i\zeta) \frac{\mathcal{B}(T-t)}{2} z\right) - 1 \right) dv (e^{\rho z} - 1) \nu(dz) \\
&= \int_0^\infty \frac{1}{\pi} \frac{K^{-i\zeta+1} \phi(\zeta)}{(i\zeta - 1)i\zeta} \int_0^\infty (e^{\eta z} - 1) (e^{\rho z} - 1) \nu(dz) dv,
\end{aligned}$$

where  $\eta := i\rho\zeta - (\zeta^2 + i\zeta) \frac{\mathcal{B}(T-t)}{2}$ , which is a function of  $\zeta$ .

$I_2$  (cont'd)

Note that  $\Re(\eta) \leq 0$  when  $0 < \delta < 1 - \frac{2\rho}{\mathcal{B}(T)}$ .

Therefore, taking such an  $\delta$ , we have

$$\int_0^\infty (e^{\eta z} - 1)(e^{\rho z} - 1)\nu(dz) = ab\lambda \left( \frac{1}{b - \eta - \rho} - \frac{1}{b - \eta} - \frac{1}{b - \rho} + \frac{1}{b} \right),$$

from which we can compute  $I_2$  using the FFT.

$I_2$  (cont'd)

Note that  $\Re(\eta) \leq 0$  when  $0 < \delta < 1 - \frac{2\rho}{\mathcal{B}(T)}$ .

Therefore, taking such an  $\delta$ , we have

$$\int_0^\infty (e^{\eta z} - 1)(e^{\rho z} - 1)\nu(dz) = ab\lambda \left( \frac{1}{b - \eta - \rho} - \frac{1}{b - \eta} - \frac{1}{b - \rho} + \frac{1}{b} \right),$$

from which we can compute  $I_2$  using the FFT.

Reminder:  $\sup_{t \leq s < T} \left\{ \frac{1}{2} - \frac{\rho}{\mathcal{B}(T-s)} - \sqrt{D_s} \right\} < \delta < \inf_{t \leq s < T} \left\{ \frac{1}{2} - \frac{\rho}{\mathcal{B}(T-s)} + \sqrt{D_s} \right\}$ .

## Delta-hedging strategy

Next, we discuss **delta-hedging strategy**  $\Delta_t^{(S_T - K)^+}$  for a call option with strike price  $K$ , which is given as the partial derivative of the option price with respect to  $S_t$ ,

## Delta-hedging strategy

Next, we discuss **delta-hedging strategy**  $\Delta_t^{(S_T - K)^+}$  for a call option with strike price  $K$ , which is given as the partial derivative of the option price with respect to  $S_t$ , that is,

$$\Delta_t^{(S_T - K)^+} := e^{-(r-q)(T-t)} \frac{\partial}{\partial S_t} \mathbb{E}[(S_T - K)^+ | S_t, \sigma_t^2].$$

## Delta-hedging strategy

Next, we discuss **delta-hedging strategy**  $\Delta_t^{(S_T-K)^+}$  for a call option with strike price  $K$ , which is given as the partial derivative of the option price with respect to  $S_t$ , that is,

$$\Delta_t^{(S_T-K)^+} := e^{-(r-q)(T-t)} \frac{\partial}{\partial S_t} \mathbb{E}[(S_T - K)^+ | S_t, \sigma_t^2].$$

Noting that

$$\mathbb{E}[(S_T - K)^+ | S_t, \sigma_t^2] = \frac{1}{\pi} \int_0^\infty K^{-i\zeta+1} \frac{\phi(\zeta)}{(i\zeta - 1)i\zeta} dv,$$

we have

$$\begin{aligned} \Delta_t^{(S_T-K)^+} &= \frac{e^{-(r-q)(T-t)}}{\pi} \int_0^\infty \frac{K^{-i\zeta+1}}{(i\zeta - 1)i\zeta} \frac{\partial \phi(\zeta)}{\partial S_t} dv \\ &= \frac{e^{-(r-q)(T-t)}}{\pi} \int_0^\infty K^{-i\zeta+1} \frac{\phi(\zeta) S_t^{-1}}{i\zeta - 1} dv = \frac{I_1}{S_t}. \end{aligned}$$

# Numerical experiments

We show numerical results on LRM strategies  $\xi_t^{(S_T - K)^+}$  and delta-hedging strategies  $\Delta_t^{(S_T - K)^+}$ .



# Numerical experiments

We show numerical results on LRM strategies  $\xi_t^{(S_T - K)^+}$  and delta-hedging strategies  $\Delta_t^{(S_T - K)^+}$ .

Reminder:

we take  $T = 1$ ,  $r = 0.019$ ,  $q = 0.012$ ,  $S_t = 1124.47$ ,  $\sigma_t^2 = 0.0145$ ,  
 $\rho = -1.2606$ ,  $\lambda = 0.5783$ ,  $a = 1.4338$ ,  $b = 11.6641$ .

Moreover, we take  $\delta = 1.75$ .

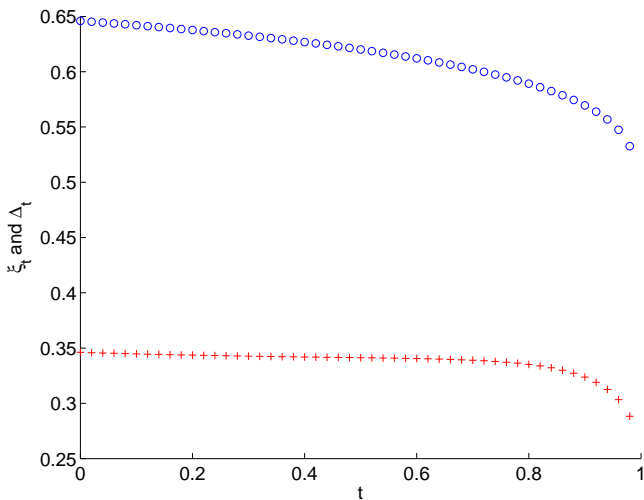


Figure: Values of  $\xi_t^{(S_T-K)^+}$  and  $\Delta_t^{(S_T-K)^+}$  when  $K$  is fixed to 1124.47(ATM) vs. times  $t = 0, 0.02, \dots, 0.98$ . In this case, the option is in the money at time  $t$ . Red crosses and blue circles represent the values of  $\xi_t^{(S_T-K)^+}$  and  $\Delta_t^{(S_T-K)^+}$ , respectively.

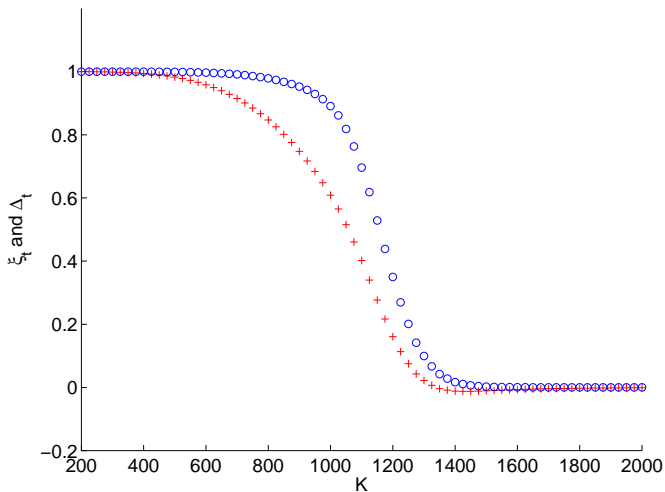


Figure: Values of  $\xi_t^{(S_T-K)^+}$  and  $\Delta_t^{(S_T-K)^+}$  at  $t = 0.5$  vs. strike price  $K$  from 200 to 2000 at steps of 25. Red crosses and blue circles represent the values of  $\xi_t^{(S_T-K)^+}$  and  $\Delta_t^{(S_T-K)^+}$ , respectively.

Thank you for your attention!